Web Appendix

# An Analysis of the Memphis Nurse-Family Partnership Program ${ }^{1}$ 

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## A Memphis NFP Randomization Protocol

The NFP Memphis trial recruited pregnant women from June 1, 1990 to August 31, 1991 through the Memphis-Shelby County Tennessee Health Department. Eligible mothers satisfied the following biological criteria: (1) less than 29 weeks of pregnancy; (2) no previous live birth; and (3) no chronic illnesses that could contribute to fetal-growth retardation or preterm delivery. They also satisfied two or more of the following socio-economic criteria: (1) unmarried; (2) less than 12 years of education; (3) unemployed. The randomization protocol was sequential, that is to say, that each participant was randomized according to the order of enrollment. Pregnant women who agreed to enroll were classified in strata defined by 5 characteristics:

1. Maternal race (African American vs non-African American);
2. Maternal age ( $<17,17-18,>18$ years $)$;
3. Gestational age at enrollment ( $<20, \geq 20$ weeks);
4. Employment status of the head of household
5. 4 geographic regions of residence. ${ }^{1}$

Within strata, randomization was performed following the Soares and Wu (1983) method as follows:

1. If the participant had a sibling already enrolled in the program, the participant was assigned to the same treatment status as the elder sibling.
2. Else, the participant was randomized into the control (C) or treatment group (T). However, if the sample size difference between treatment and control group was larger than a threshold, the participant was deterministically assigned to the treatment status that had fewer participants.

[^0]3. Next the participant was randomized again into her final treatment status. If in step 2 the participant was assigned to the control, she was next randomized to:

- Group 1: control women who received free transportation to and from their prenatal appointments (sample size: 166).
- Group 2: control women who received developmental screening and referral services at ages 6,12 and 24 months in addition to the benefits of Group 1 (sample size: 514).

If she was assigned to the treatment group previously, she is next randomized into:

- Group 3: treated women who received home visits by nurses during pregnancy, one visit in the hospital and one visit at home after childbirth in addition to the benefits of Group 2 (sample size: 230).
- Group 4: treated women that received home visits by nurses during pregnancy until the child's $2^{\text {nd }}$ birthday, in addition to the benefits of Group 2 (sample size: $228)^{2}$

As in the previous step, if the absolute difference in group size exceeded some threshold then the participant was deterministically assigned to the group with the lowest number of participants. Otherwise, the pregnant woman was randomly assigned.

Importantly, the randomization method incorporated a trigger mechanism that deterministically assigned a treatment status to participants if the sequence of assignments became too imbalanced due to sampling variation. In this context, imbalance was measured by the difference of persons assigned to T and the persons assigned to C . In practice less than $1 \%$ of the women were assigned according to the trigger mechanism. Thus, the NFP Memphis trial can be treated as a non-sequential protocol.

[^1]
## B Brief survey of the NFP Literature

The NFP program has been evaluated in three Randomized Control Trials (RCTs): Elmira, NY (1978), Memphis, TN (1990) and Denver, CO (1994), each of them targeting low-income first-time mothers from different racial backgrounds. In Elmira, the sample was mainly low-income white; in Memphis, the majority of the participants were low-income African American; and in Denver, the sample included a significant fraction of low-income Hispanics. Another important difference is that in the Elmira and Memphis trials, the visits were conducted by nurses. In contrast, in the Denver trial, one treatment group received visits by Nurses and another group by Paraprofessionals. Table B. 1 describes the main features of the randomization arms of each NFP trial.

- Denver Trial Treatment Groups:

1. Women in the control group $(\mathrm{n}=255)$ were provided developmental screening and referral services for their children at $6,12,15,21$, and 24 months old (same as Memphis Group 2).
2. Paraprofessional Group: Women assigned to the paraprofessional group ( $\mathrm{n}=$ 245) were provided the screening and referral services plus paraprofessional home visitation during pregnancy and infancy, that is, the first 2 years of the child's life (No such group in Memphis Trial).
3. Nurse Group: Women in the nurse group ( $\mathrm{n}=235$ ) were provided screening and referral plus nurse home visitation during pregnancy and infancy (Same as Memphis Group 4).

- Elmira Trial Treatment Groups

1. Group 1. When the children were 1 and 2 years of age, an infant specialist hired by the research project screened them for sensory and developmental problems and
referred those with suspected problems to other specialists for further evaluation and treatment (similar to Memphis Group 2).
2. Group 2. Families were provided free transportation for regular prenatal care at local clinics and physicians' offices through a contract with a local taxicab company, as well as the sensory and developmental screening outlined in treatment 1 (same as Memphis Group 2).
3. Group 3. Families were provided a nurse home visitor during pregnancy, in addition to the screening and transportation services. The nurses visited families approximately once every 2 weeks and made an average of nine visits during pregnancy. The average visit lasted 1 hour and 15 minutes (similar but weaker treatment than Memphis Group 3).
4. Group 4. Families received the same services as those in treatment 3, but in addition the nurse continued to visit until the children were 2 years of age. For 6 weeks after delivery the nurses visited families every week; from 6 weeks to 4 months, they visited every 2 weeks; from 4 to 14 months, every 3 weeks; from 14 to 20 months, every 4 weeks; and from 20 to 24 months, every 6 weeks. Under predetermined crisis conditions the nurses visited weekly. As with the visits during pregnancy, the average visit lasted approximately 1 hour and 15 minutes, but the mean number of visits completed from birth through the end of the program was 23 (same as Memphis Group 4).

## B. 1 Summary of key Findings of the NFP Literature

We summarize selected papers on NFP below. We also describe their main results. Tables K.1-K. 15 provide further information of the published papers of the NFP literature.

1. Memphis Trial

Table B.1: Description of Randomization Groups for Each NFP Trial

Elmira Trial
Services provided (+) in Each of the Four Treatment Groups

|  | 1 <br> $(\mathrm{~N}=90)$ | 2 <br> $(\mathrm{~N}=94)$ | 3 <br> $(\mathrm{~N}=100)$ | 4 <br> $(\mathrm{~N}=116)$ |
| :--- | :---: | :---: | :---: | :---: |
| Health and developmental <br> screening at the child's $12^{\text {th }}$ and <br> $24^{\text {th }}$ month of life | X | X | X | X |
| Free transportation to regular <br> prenatal and well-child visits |  | X | X | X |
| Nurse home visitation during <br> pregnancy |  |  | X | X |
| Nurse home visitation during <br> the child's first 2 years of life |  |  | X |  |

Memphis Trial
Services provided (+) in Each of the Four Treatment Groups

|  | 1 <br> $(\mathrm{~N}=166)$ | 2 <br> $\mathrm{~N}=(514)$ | 3 <br> $(\mathrm{~N}=230)$ | 4 <br> $(\mathrm{~N}=228)$ |
| :--- | :---: | :---: | :---: | :---: |
| Free transportation to regular <br> prenatal and well-child visits | X | X | X | X |
| Health and developmental <br> screening at the child's 6 th, $12^{\text {th }}$ |  |  |  |  |
| and $24^{\text {th }}$ month of life |  |  |  |  | X

## Denver Trial

Services provided (+) in Each of the Three Treatment Groups

|  | 1 <br> $(\mathrm{~N}=255)$ | 2 <br> $(\mathrm{~N}=245)$ | 3 <br> $(\mathrm{~N}=235)$ |
| :--- | :---: | :---: | :---: |
| Developmental screening and <br> referral services at the child's <br> $6^{\text {th }}, 12^{\text {th }}, 15^{\text {th }}, 21^{\text {st }}$ and $24^{\text {th }}$ <br> month of life | X | X | X |
| Paraprofessional home <br> visitation during pregnancy and <br> the child's first 2 years of life |  | X |  |
| Nurse home visitation during <br> pregnancy and the child's first 2 <br> years of life |  | X |  |

- Kitzman et al. (2000):

Treatment group had fewer subsequent pregnancies; longer interval between the birth of the first child and the second; fewer months of using welfare. All results were statistically significant.

- Olds et al. (2004):

Women visited by nurses had fewer subsequent pregnancies and births, longer intervals between births of the first and second children, longer relationships with current partners, fewer months of using welfare. Nurse-visited children were more likely to be enrolled in out-of-home care between 2 and 4.5 years, had higher higher intellectual functioning, higher vocabulary scores, fewer behavior problems and higher arithmetic achievement

- Olds et al. (2007):

Nurse-visited women had longer intervals between the births of their first and second children, fewer subsequent births, and longer relationships with current partners. From birth through child age 9, nurse-visited women used welfare and food stamps for fewer months. Nurse-visited children born to mothers with low psychological resources had higher achievement test scores in math and reading in grades 1 through 3 .
2. Denver Trial

- Olds et al. (2002):

Nurse-visited mothers reduced smoking during pregnancy, had fewer subsequent pregnancies and births; had delayed subsequent pregnancies; worked more; interacted more with the child. At 6 months of age, nurse-visited infants were less likely to exhibit emotional vulnerability in response to fear stimuli and nursevisited infants born to women with low psychological resources were less likely to exhibit low emotional vitality in response to joy and anger stimuli. At 21 months,
nurse-visited children born to women with low psychological resources were less likely to exhibit language delays. At 24 months, they exhibited superior mental development (Development Index scores). No statistically significant program effects for use of ancillary prenatal services, educational achievement, use of welfare, child temperament or behavior problems. There is a single statistically significant result for paraprofessional-visited mothers when compared to their control group. Mothers with low psychological resources visited by paraprofessional interacted more with their children.

- Olds et al. (2004):

Women who were visited by paraprofessionals, were less likely to be married, more likely to live with the biological father, were more likely to work, reported a greater sense of mastery and had better mental health. Paraprofessional-visited women, had fewer subsequent miscarriages, gave birth to fewer low birth weight newborns, displayed greater sensitivity and responsiveness toward one another and, for mothers with low psychological resources, home environments were more supportive. Nurse-visited mothers reported greater intervals between the births of their first and second children, experienced less domestic violence and enrolled their children less frequently in preschool, Head Start, or licensed day care than did control subjects. Nurse-visited children whose mothers had low levels of psychological resources had home environments that were more supportive of childrens early learning, more advanced language, and superior executive functioning. There were no statistically significant effects of either nurse or paraprofessional visits on the number of subsequent pregnancies, mother's educational achievement, use of substances, use of welfare, or children's externalizing behavior problems.

## 3. Elmira Trial

- Olds et al. (1997):

Women who were visited by nurses during pregnancy and infancy were less likely to be perpetrators of child abuse and neglect compared to the control group. Nurse-visited women from low socioeconomic status households who were unmarried had 1.3 vs 1.6 subsequent births and delayed the second birth. Nurse-visited women were less likely to use welfare, had fewer behavioral impairments due to use of alcohol and other drugs and had fewer arrests.

- Eckenrode et al. (2010):

Nurse-visited mothers had fewer lifetime arrests and convictions. Nurse-visited mothers of low-income had fewer children and were less likely to use Medicaid.

We present a more detailed analysis of these trials in Appendix K.

## C Assessment Instruments

## C. 1 HOME

The Home Observation for Measurement of the Environment (HOME) was first developed in the 1960 s by Caldwell. It measures the quality and quantity of stimulation and support available to a child at home (Caldwell and Bradley, 1984). The assessment documents the child's home environment and is observed during a visit of 45 to 90 minutes. A more in depth explanation of the HOME inventory is in Caldwell and Bradley (1984).

There are several versions of the inventory. The initial version, Infant/Toddler HOME, is for children aged 0 to 3 years old. It consists of 45 binary-choice items grouped into 6 subscales. The Early Childhood HOME is for children aged 3 to 6 years old. It consists of 55 binary-choice items clustered into 8 subscales. Finally, the Middle Childhood HOME is used for children aged 6 to 10 years old. It consists of 59 items in 8 subscales. The NFP uses the first version of the Inventory, the Infant/Toddler HOME.

The 45 items of the HOME inventory contain the six following subscales:

1. Emotional and Verbal Responsiveness of the Mother (11 items): measures the mother's ability to communicate with the child.
2. Avoidance of Restriction and Punishment (8 items): measures the mother's ability to discipline the child.
3. Organization of the Environment (6 items): measures the daily changes in the child's environment.
4. Provision of Appropriate Play Material (9 items): measures the types of toys and their contributions to the child's motor skills.
5. Maternal Involvement with Child (6 items): measures the aspects in which the mother is involved in the child's daily life.
6. Opportunities for Variety in Daily Stimulation (5 items): measures the levels of interaction the mother and other family members have with the child.

The NFP measured the HOME Inventory when the child was 12 and 24 months old.

## C. 2 KABC

The Kaufman Assessment Battery for Children (K-ABC) was developed by Alan S. Kaufman and Nadeen L. Kaufman in 1983 with a later revision in 2004. The KABC focuses on processes required to solve problems compared to psychological instruments that focus on measuring raw cognitive skills. In broad terms, KABC focuses on the process of acquiring and manipulating information according to a determined protocol. The KABC contains 16 subtests ( 10 mental processing and 6 achievement), which can be grouped into 3 scales. Due to the nature of the subtests, 13 subtests can be taken at once, with the mandatory age range to be between 7 to 12.5 years old. The NFP used the following 11 subtests:

1. Sequential Processing Scale (Hand Movements, Number Recall, Word Order): measures short-term memory and problem-solving skills. It emphasizes how children are able to follow ordered sequences.
2. Simultaneous Processing Scale (Gestalt Closure, Triangles, Matrix Analogies, Spatial Memory, Photo Series): measures problem-solving skills. It involves several processes at once such as scenes in a partially completed picture.
3. Achievement Scale (Arithmetic, Riddles, Reading/Decoding): measures achievement and focuses on applied skills and facts learned through the home/school environment.

The NFP Program used these three scales when the child was 6 years old.

## C. 3 PPVT

The Peabody Picture Vocabulary Test (PPVT) is an individual verbal intelligence test that measures receptive vocabulary, developed by Llyod M. Dunn and Leota M. Dunn in 1959. It is a verbal test that lasts between 20 and 30 minutes. The child is presented a series of pictures. There are four pictures in a page. The examiner states a word and asks the child to associate it with a picture. The diffusion of the figures increases over time. The exam stops when the child answers six out of eight questions incorrectly. After completion, the raw score is given, normalized to a mean of 100 and standard deviation of 15 . The NFP Program used PPVT when the child was 6 years old.

## C. 4 WISC-III

The Wechsler Intelligence Scale for Children - Third Edition (WISC-III) was created in 1949. The third edition was published in 1991 (Wechsler, 1991). WISC is an intelligence test for children between the ages 6 and 16 years old. It can be completed without reading or writing. The exam takes between 65 and 80 minutes. There are two subscales: verbal and performance, which provide a Verbal IQ (VIQ), a Performance IQ (PIQ), and a Full Scale IQ (FSIQ). The NFP only used the coding part of the Processing Speed Index:

1. Coding: the child marks rows of shapes with different lines to transcribe a digit-symbol code. It measures visual or motor integration and visual scanning.

The NFP Program used WISC-III when the child was 6 years old.

## C. 5 CBCL

The Child Behavior Checklist (CBCL) is a parent-report questionnaire developed by Thomas M. Achenbach. In it, the child is rated on several behavioral and emotional problems. The goal of the inventory is to assess internalizing and externalizing behaviors. The responses are recorded using a Likert scale: $0=$ Not True, $1=$ Sometimes True, $2=$ Very True. The preschool checklist ( 18 months to 5 years) contains 100 questions and the school-age checklist ( 6 to 18 years) contains 120 questions. The preschool checklist questions can be broken down into the following subscales: anxious/depressed, withdrawn, sleeping problems, somatic problems, aggressive behavior, and destructive behavior. The school-age checklist questions can be broken down into the following subscales: withdrawn, somatic complaints, anxious/depressed, social problems, thought problems, attention problems, delinquent behavior, aggressive behavior, and other problems. The NFP Program used the CBCL when the child was 2 and 6 years old.

## C. 6 MacArthur

The MacArthur Story Stem Battery (MSSB) was created by the MacArthur Narrative Working Group that included Bretherton, Buchsbaum, and several other collaborators. The story stem method is a procedure in which the examiner presents a story to the child that culminates at a high point, at which the child is then asked to complete the story; this type of method allows insight into the inner workings of the child's mind. The MSSB uses 15 stories and measures: dysregulated aggression, empathy/warmth, emotional integration, and performance anxiety.

1. Dysregulated Aggression Dimension: aggression, injury, danger, destruction, dishonesty, escalation of conflict, negative story endings, inappropriate child power, controlling toward examiner.
2. Empathy/Warmth Dimension: empathy-helping, affiliation, affection, reparation or guilt, parental warmth.
3. Emotional Integration Construct: ability to maintain story coherence with the inclusion of emotional expression. The affects included are joy, anger, distress, concern, sadness.
4. Avoidance or Withdrawal Dimension: characters leaving the scene repetition of previous story fragments, denial of central conflict or challenge, family characters leave, avoiding separation from parents, dissociative behaviors.
5. Performance Anxiety Dimension: unwillingness to verbalize, unresponsiveness to examiner, anxious behaviors.

The NFP Program used the MSSB when the child was 6 years old.

## D Permutation-based Inference and Multiple Hypothesis Testing

The standard model of program evaluation describes the observed outcome $Y_{i}$ of participant $i \in J$ by

$$
\begin{equation*}
Y_{i}=D_{i} Y_{i, 1}+\left(1-D_{i}\right) Y_{i, 0} \tag{1}
\end{equation*}
$$

where $J=\{1, \ldots, N\}$ denotes the sample space indexing set, $D_{i}$ denotes the treatment assignment for participant $i \in J,\left(D_{i}=1\right.$ if treatment occurs, $D_{i}=0$ otherwise $)$ and $\left(Y_{i, 0}, Y_{i, 1}\right)$ are potential outcomes for participant $i$ when treatment is fixed at control and treatment status respectively.

Randomized experiments solve potential problems of selection bias by inducing independence between counterfactual outcomes $\left(Y_{i, 0}, Y_{i, 1}\right)$ and treatment status $D_{i}$ when conditioned on the pre-program variables $\boldsymbol{X}$ used in the randomization protocol. All variables are defined in the common probability space $(\Omega, \mathscr{F}, \mathrm{P})$. In our notation, a randomized experiment must satisfy the following assumption:

Assumption A-1. $Y(d) \Perp D \mid \boldsymbol{X} ; d \in \operatorname{supp}(D)$,
where variables $\boldsymbol{X}=\left(\boldsymbol{X}_{i} ; i \in J\right), D=\left(D_{i} ; i \in J\right)$ are $N$-dimensional vectors of treatment assignments and pre-program variables, and $Y(d)=\left(Y_{i, d_{i}} ; i \in J, d_{i} \in\{0,1\}\right)$ and $d \in$ $\operatorname{supp}(D)=\{0,1\}^{|J|}$ denotes the vector of counterfactual outcomes. In the same fashion, we represent the vector of observed outcomes of Equation (1) by $Y=\left(Y_{i} ; i \in \mathcal{I}\right)$. The no-treatment hypothesis is equivalent to the statement that the conditional counterfactual outcome vectors share the same distribution:

Hypothesis H-1. $Y(d) \stackrel{d}{=} Y\left(d^{\prime}\right) \mid \boldsymbol{X} ; d, d^{\prime} \in \operatorname{supp}(D)$,

Hypothesis $\mathbf{H}-\mathbf{1}$ can be restated in more tractable form:

Hypothesis H-1 ${ }^{\prime}$. Under Assumption A-1 and Hypothesis H-1, we have that $Y \Perp D \mid \boldsymbol{X}$.

Testing Hypothesis $\mathbf{H}-\mathbf{1}^{\prime}$ poses some statistical challenges. First, small sample sizes cast doubt on inference that relies on the asymptotic behavior of test statistics. We address the problem of small sample size by generating the exact test statistic conditioned on data. Second, the presence of multiple outcomes allows for the arbitrary selection of statistically significant outcomes. Selectively reporting statistically significant outcomes is often termed cherry picking and generates downward-biased $p$-values. We solve the problem of cherry picking by implementing multiple-hypothesis testing based on the stepdown procedure of (Romano and Wolf, 2005). They explain that the stepdown procedure strongly controls for family-wise error rate (FWER), while classical tests do not. Also, Romano and Wolf (2005) show that the strong FWER control can be obtained by imposing a certain monotonicity condition on the test statistics. This requirement is weaker than the assumption of subset pivotality, used in various methods of resampling outcomes presented in Westfall and Young (1993).

To summarize, our method is based on three steps. First, we seek to characterize the exact conditional distribution of $D \mid \boldsymbol{X}$. Specifically we characterize the set $D_{x}(d)$, defined by:

$$
D_{x}(d)=\left\{d^{\prime} \in\{0,1\}^{|J|} ; \mathrm{P}(D=d \mid \boldsymbol{X}=x)=\mathrm{P}\left(D=d^{\prime} \mid \boldsymbol{X}=x\right)\right\}
$$

such that the distribution of $D$ conditioned on realized data is uniform among elements of $D_{x}(d)$. Next we use the assumption of the null hypothesis of no-treatment effects, i.e. $H_{0}: Y \Perp D \mid \boldsymbol{X}$, to generate the exact conditional distribution of a test statistic $T(Y, D) \mid \boldsymbol{X}$. Under $H_{0}$, we can construct an inference that controls for the probability of falsely rejecting the null hypothesis. We control for this probability in two ways: (1) in the case of single (joint) null hypothesis, we control for the standard Type-I error; (2) in the case of multiple hypothesis inference, we control for the family-wise error rate.

More notation is helpful for describing the method. Let $K$ represent the indexing set for all available outcomes $Y_{k} ; k \in K$. We represent the single (joint) null hypothesis that a set $L \subset K$ of outcomes $Y_{k} ; k \in L$ are jointly independent of treatment status $D$ conditional on pre-program variables $\boldsymbol{X}$ by

Hypothesis H-1". $H_{L}: Y_{L} \Perp D \mid \boldsymbol{X}$, where $Y_{L}=\left(Y_{k}: k \in L\right)$.

When $L$ is a singleton, say $L=\{k\}$, then the null hypothesis is given by $H_{\{k\}}: Y_{k} \Perp$ $D \mid \boldsymbol{X}$. In this notation, we can write the joint Hypothesis $\mathbf{H - 1}{ }^{\prime \prime}$ as $H_{L}=\cap_{k \in L} H_{\{k\}}$.

Our goal is to test single (or joint) null hypotheses controlling for the probability of Type I error at level $\alpha$, that is, P ( reject $H_{L} \mid H_{L}$ is true $) \leq \alpha$. To do so, we rely on the fact that, under $H_{L}$,

$$
\begin{equation*}
\left(Y_{L}, D\right)\left|\boldsymbol{X} \stackrel{d}{=}\left(Y_{L}, g D\right)\right| \boldsymbol{X} \forall g \in \mathscr{G}_{X}, \tag{2}
\end{equation*}
$$

where $\mathscr{G}_{X}$ comprises all the permutations within strata of $\boldsymbol{X}$, that is,

$$
\mathscr{G}_{X}=\left\{g ; g: J \rightarrow J \text { is a bijection and } g(j)=j^{\prime} \Rightarrow\left(\boldsymbol{X}_{j}\right)=\left(\boldsymbol{X}_{j^{\prime}}\right)\right\},
$$

and $g D$ is a vector defined by:

$$
g D=\left(\tilde{D}_{i} \in \operatorname{supp}(D) ; i \in J \text { and } \tilde{D}_{i}=D_{g(i)}\right)
$$

We use Relation (2) to generate a statistical test where the exact distribution of the test
statistic $T_{L}\left(Y_{L}, g D\right)$ is obtained by re-evaluating $T_{L}\left(Y_{L}, g D\right)$ as $g$ varies in $\mathscr{G}_{X}$. Note that the inference for Hypothesis $\mathbf{H}-\mathbf{1}^{\prime \prime}$ depends on the choice of statistics. That is to say that even though any statistic $T_{L}\left(Y_{L}, D\right)$ whose value provides evidence against the null hypothesis can be used, the inference is dependent on this choice of statistic. An example of such statistic is the maximum of the $t$-statistic associated with the difference in means between treated and control groups over outcomes $Y_{k}$ such that $k \in L$. Formally,

$$
\begin{equation*}
T_{L}\left(Y_{L}, D\right)=\max _{k \in L} T_{k}\left(Y_{k}, D\right) \tag{3}
\end{equation*}
$$

where $T_{k}\left(Y_{k}, D\right)$ is the $t$-statistic for outcome $Y_{k}$. Relationship (2) implies that $T_{L}\left(Y_{L}, D\right) \mid \boldsymbol{X} \stackrel{d}{=}$ $T_{L}\left(Y_{L}, g D\right) \mid \boldsymbol{X}$ for any $g \in \mathscr{G}_{X}$. Moreover, let $d \in\{0,1\}^{|J|}$ such that $\mathrm{P}(D=d \mid \boldsymbol{X}=x)>0$, then the distribution of $D$ conditioned on $\boldsymbol{X}=x$ is uniform across elements of $D_{x}(d)$ (see Lehmann and Romano (2005), Chapter 15). Thus, a critical value $c_{L, x}\left(Y_{L}, d, \alpha\right)$ such that $\mathrm{P}\left(T_{L}\left(Y_{L}, D\right)>c_{L, x}\left(Y_{L}, d, \alpha\right) \mid \boldsymbol{X}=x, H_{L}\right.$ is true $) \leq \alpha$ can be computed as:

$$
c_{L, x}\left(Y_{L}, d, \alpha\right)=\inf _{t \in \mathbf{R}}\left\{\sum_{d^{\prime} \in D_{x}(d)} I\left\{T_{L}\left(Y_{L}, d^{\prime}\right) \leq t\right\} \geq(1-\alpha)\left|D_{x}\right|\right\}
$$

where $I\{\cdot\}$ is the indicator function. The following notation is useful to further characterize $c_{L, x}\left(Y_{L}, d, \alpha\right)$. Let $T_{L, x}^{(1)}, \ldots, T_{L, x}^{\left(\left|D_{x}(d)\right|\right)}$ be the sequence of increasing ordered statistics $T_{L}\left(Y_{L}, d^{\prime}\right)$ as $d^{\prime}$ varies in $D_{x}(d)$. In this notation we can write the critical value as

$$
\begin{equation*}
c_{L, x}\left(Y_{L}, d, \alpha\right)=T_{L, x}^{\left(\left\lceil(1-\alpha)\left|D_{x}\right|\right\rceil\right)} \tag{4}
\end{equation*}
$$

where $\lceil a\rceil$ stands for the smallest integer bigger or equal than $a$.
Under the null hypothesis $H_{L}$, the probability of a test statistic be bigger or equal than the statistic $T_{L}\left(Y_{L}, d\right)$ actually observed, i.e. the p -value, is given by:

$$
\begin{equation*}
p_{L, x}(d)=\inf _{\alpha \in[0,1]}\left\{c_{L, x}\left(Y_{L}, d, \alpha\right) \leq T_{L}\left(Y_{L}, d\right)\right\} . \tag{5}
\end{equation*}
$$

Now let $r_{L, x} \in\left\{1, \ldots,\left|D_{x}(d)\right|\right\}$ be the lowest rank that the value of the observed test statistic $T_{L}\left(Y_{L}, d\right)$ takes in the sequence $T_{L, x}^{(1)}, \ldots, T_{L, x}^{\left(\left|D_{x}(d)\right|\right)}$, that is to say:

$$
r_{L, x}=1+\sum_{d^{\prime} \in D_{x}(\mathbf{d})} I\left\{T_{L}\left(Y_{L}, d^{\prime}\right)<T_{L}\left(Y_{L}, d\right)\right\}
$$

Thus:

$$
\begin{equation*}
T_{L, x}^{\left(r_{L, x}\right)}=T_{L}\left(Y_{L}, d\right) \tag{6}
\end{equation*}
$$

Then, by the ordered property of $T_{L, x}^{(r)} ; r \in\left\{1, \ldots,\left|D_{x}(d)\right|\right\}$ and the definition of $r_{L, x}$, we have that:

$$
\begin{equation*}
p_{L, x}(d)=1-\frac{r_{L, x}}{\left|D_{x}(d)\right|} . \tag{7}
\end{equation*}
$$

Moreover, p -value $p_{L, x}(d)$ complies with the following property:

$$
\mathrm{P}\left(p_{L, x}(d) \leq \phi \mid \boldsymbol{X}=\boldsymbol{x}\right) \leq \phi \forall \phi \in[0,1] .
$$

We implement a method of inference that tests the multiple null hypotheses that each outcome $Y_{k} ; k \in L$ is independent of treatment status $D$ conditional on pre-program variables $\boldsymbol{X}$. The representation of these multiple hypothesis is in the same fashion as the single (joint) null hypothesis, namely, $H_{L}=\cap_{k \in L} H_{\{k\}} ; H_{\{k\}}: Y_{k} \Perp D \mid \boldsymbol{X}$. The multiple hypothesis testing differs from the single (joint) hypothesis testing in the way it controls for the probability of false rejection. Specifically, let the subset $L_{0}$ be the set of true Hypothesis $H_{\{k\}}$ such that $k \in L_{0} \subset L$. Our multiple hypothesis testing controls for the family-wise error rate (FWER), that is, the probability of even one false rejection among the set of true hypothesis $L_{0}$. Formally, we control for:

$$
\mathrm{P}\left(\text { reject at least one } H_{\{k\}} ; k \in L_{0} \mid H_{L_{0}} \text { is true }\right) \leq \alpha
$$

while single (joint) hypothesis testing controls for $\mathrm{P}\left(\right.$ reject $H_{L} \mid H_{L}$ is true $) \leq \alpha$.

Bonferroni or Holm are examples of inference methods that test multiple hypothesis controlling for FWER. These methods rely upon a "least favorable" dependence structure among the p-values. The stepdown procedure of Romano and Wolf (2005) is less conservative as it accounts for the dependence structure of $p$-values. The method is based on a monotonicity assumption, which, in our case, can be stated as:

$$
\begin{equation*}
c_{K, x}\left(Y_{K}, d, \alpha\right) \geq c_{L_{0}, x}\left(Y_{L_{0}}, d, \alpha\right) \text { for any subset } K \text { of } L \text { containing } L_{0} \text { i.e. } L_{0} \subset K \subset L \tag{8}
\end{equation*}
$$

Equation (8) is satisfied by our choice of test statistic (3) and the fact that $L_{0} \subset K$.
The stepdown procedure given in Romano and Wolf (2005) is a stepwise method summarized in the following algorithm:

## Algorithm 1.

Step 1: Set $L_{1}=L$. If

$$
\begin{equation*}
\max _{k \in L_{1}} T_{k}\left(Y_{k}, d\right) \leq c_{L, x}\left(Y_{L_{1}}, d, \alpha\right) \tag{9}
\end{equation*}
$$

then stop and reject no null hypotheses; otherwise, reject any $H_{\{k\}}$ with

$$
T_{k}\left(Y_{k}, d\right)>c_{L, x}\left(Y_{L_{1}}, d, \alpha\right)
$$

and go to Step 2.
$\vdots$

Step $j$ : Let $L_{j}$ denote the indices of remaining null hypotheses. If

$$
\begin{equation*}
\max _{k \in L_{j}} T_{k}\left(Y_{k}, d\right) \leq c_{L, x}\left(Y_{L_{j}}, d, \alpha\right) \tag{10}
\end{equation*}
$$

then stop and reject no further null hypotheses; otherwise, reject any $H_{\{k\}}$ with

$$
T_{k}\left(Y_{k}, d\right)>c_{L, x}\left(Y_{L_{j}}, \mathbf{d}, \alpha\right)
$$

and go to Step $j+1$.
$\vdots$

We can compute the multiplicity-adjusted p-values of Equations (9)-(10) in the same fashion described by Equations (5)-(7).

## D. 1 Conditioning and Linearity

A typical problem in small sample randomized trials is sampling variation, where preprogram variables differ across treatment groups by chance. One can increase the power of any statistical inference by conditioning on those pre-program variables. Let $\boldsymbol{Z}$ be the pre-program variables that were not used in the randomization protocol that we seek to control for.

Variables $\boldsymbol{Z}$ precede the treatment intervention and therefore $\boldsymbol{Z} \Perp D \mid \boldsymbol{X}$ holds due to randomization. Under the hypothesis of no-treatment, $\boldsymbol{Y} \Perp D \mid \boldsymbol{X}$ also holds. These two relations imply that $\boldsymbol{Y} \Perp D \mid(\boldsymbol{X}, \boldsymbol{Z})$. We can use this relationship to generate a permutation test that considers the strata formed by values of covariates $\boldsymbol{X}$ and $\boldsymbol{Z}$. This way we can generate an inference method that non-parametrically conditions on variables $\boldsymbol{X}$ and $\boldsymbol{Z}$.

Non-parametric conditioning through block permutation comes at a cost. A fine conditioning set decreases the share of available data that can be permuted and a sufficiently large conditioning set prohibits the implementation of a permutation-based test. We solve this problem by evoking linearity. That is to say, we condition variables through a linear regression instead of a non-parametric block permutation. Anderson and Legendre (1999) test a range of permutation methods for linear models. They find that the Freedman and

Lane (1983) procedure generates the most consistent and reliable results among the available models in this literature.

We non-parametrically condition on variables used in the randomization protocol to achieve valid exchangeable properties (i.e. we use permutations in $\mathscr{G}_{X}$ ); We linearly condition on additional pre-program variables $\boldsymbol{Z}$ not used in the randomization protocol. Following the Freedman and Lane (1983) method, our approach can be summarized by the following steps: (1) compute the residuals $\boldsymbol{Y}-\boldsymbol{Z} \hat{\boldsymbol{\beta}}$ such that $\hat{\boldsymbol{\beta}}=\left(\boldsymbol{Z}^{\prime} \boldsymbol{Z}\right)^{-1} \boldsymbol{Z}^{\prime} Y$; (2) permute these residuals according to permutations $g \in \mathscr{G}_{X}$. (3) add these permuted residuals to $\boldsymbol{Z} \hat{\boldsymbol{\beta}}$, call it $\tilde{\boldsymbol{Y}}$; (4) regress $\tilde{\boldsymbol{Y}}$ on $\boldsymbol{Z}$ and the vector treatment statuses $\boldsymbol{D}$. (5) we then use the $t$-statistic associated with covariate $\boldsymbol{D}$ of the last regression as test statistic.

Beaton (1978) and Freedman and Lane (1983) suggest permutation inference based on Shuffle Residuals. By this, we mean regressing $Y$ on $\boldsymbol{X}$, shuffling the residuals from this regression, and adding them to the predicted $\boldsymbol{Y}$, say $\hat{\boldsymbol{Y}}$, to form a new variable, say $\tilde{\boldsymbol{Y}}$, which is then regressed on $\boldsymbol{Z}$ and $\boldsymbol{D}$. Formally, let the regression:

$$
\boldsymbol{Y}=\boldsymbol{Z} \boldsymbol{\beta}+\boldsymbol{D} \boldsymbol{\delta}+\epsilon
$$

where $\boldsymbol{Z}$ stands for the pre-program variables we wish to control for and includes a vector of elements ones that play the role of a contant term for the regression. Error term $\epsilon$ is a mean-zero exogenous random variable independent of $\boldsymbol{Z}$ and $\boldsymbol{D}$.

Now let $\boldsymbol{B}_{g} ; g \in \mathscr{G}_{X}$ be a permutation matrix associated with a permutation $g$ in $\mathscr{G}_{X}$. Let the operator that projects a vector in the orthogonal space generated by columns of $\boldsymbol{Z}$ be $\boldsymbol{M}_{\boldsymbol{Z}}=\boldsymbol{I}-\boldsymbol{Z}\left(\boldsymbol{Z}^{\prime} \boldsymbol{Z}\right)^{-1} \boldsymbol{Z}^{\prime}$, where $I$ denotes the identity matrix. As properties of Matrix $\boldsymbol{M}_{\boldsymbol{Z}}$, we can say that $\boldsymbol{M}_{\boldsymbol{Z}}$ is symmetric and idempotent, that is:

$$
\begin{equation*}
M_{Z}=M_{Z}^{\prime}=M_{Z} M_{Z}=M_{Z}^{\prime} M_{Z} \tag{11}
\end{equation*}
$$

The estimated residuals of $Y$ generated by the the regression

$$
\boldsymbol{Y}=\boldsymbol{Z} \boldsymbol{\beta}+\boldsymbol{\epsilon}
$$

is given by $\hat{\boldsymbol{e}}=\boldsymbol{M}_{\boldsymbol{Z}} \boldsymbol{Y}$. The predicted outcome based on this regression is given by: $\hat{\boldsymbol{Y}}=$ $\boldsymbol{Z}\left(\boldsymbol{Z}^{\prime} \boldsymbol{Z}\right)^{-1} \boldsymbol{X}^{\prime} \boldsymbol{Y}$.

We define the new outcome based on the sum of the predicted outcome $\hat{\boldsymbol{Y}}$ with permuted errors $\hat{e}$ according to permutation $g \in \mathscr{G}_{X}$ as

$$
\begin{equation*}
\tilde{\boldsymbol{Y}}=\hat{\boldsymbol{Y}}+B_{g} \hat{e} . \tag{12}
\end{equation*}
$$

We then use the newly computed outcome in the following regression:

$$
\begin{equation*}
\tilde{\boldsymbol{Y}}=\boldsymbol{Z} \boldsymbol{\beta}+\boldsymbol{D} \boldsymbol{\delta}+\tilde{\boldsymbol{\epsilon}} \tag{13}
\end{equation*}
$$

We now examine the $\delta$ estimate on Equation (13). This estimate ia actually the same as the one computed in the following regression:

$$
\begin{equation*}
M_{Z} \tilde{\boldsymbol{Y}}=\boldsymbol{M}_{Z} \boldsymbol{D} \boldsymbol{\delta}+\tilde{\boldsymbol{\epsilon}} \tag{14}
\end{equation*}
$$

Thus, by applying the Ordinary Least Square formula, we obtain:

$$
\begin{equation*}
\hat{\delta}_{g}=\left(D^{\prime} M_{Z}^{\prime} M_{Z} D\right)^{-1} D^{\prime} M_{Z}^{\prime} M_{Z} \tilde{Y} \tag{15}
\end{equation*}
$$

We now use previous equations to transform Equation (15) into a more general formula:

$$
\begin{align*}
\hat{\boldsymbol{\delta}}_{\boldsymbol{g}} & =\left(\boldsymbol{D}^{\prime} \boldsymbol{M}_{\boldsymbol{Z}}^{\prime} \boldsymbol{M}_{\boldsymbol{Z}} \boldsymbol{D}\right)^{-1} \boldsymbol{D}^{\prime} \boldsymbol{M}_{\boldsymbol{Z}}^{\prime} \boldsymbol{M}_{\boldsymbol{Z}} \tilde{\boldsymbol{Y}}, \text { by }(15), \\
& =\left(\boldsymbol{D}^{\prime} \boldsymbol{M}_{\boldsymbol{Z}} \boldsymbol{D}\right)^{-1} \boldsymbol{D}^{\prime} \boldsymbol{M}_{\boldsymbol{Z}} \tilde{\boldsymbol{Y}}, \text { by }(11), \\
& =\left(\boldsymbol{D}^{\prime} \boldsymbol{M}_{\boldsymbol{Z}} \boldsymbol{D}\right)^{-1} \boldsymbol{D}^{\prime} \boldsymbol{M}_{\boldsymbol{Z}}\left(\boldsymbol{Y}+\boldsymbol{B}_{g} \hat{\boldsymbol{e}}\right), \text { by }(12), \\
& =\left(\boldsymbol{D}^{\prime} \boldsymbol{M}_{\boldsymbol{Z}} \boldsymbol{D}\right)^{-1} \boldsymbol{D}^{\prime} \boldsymbol{M}_{\boldsymbol{Z}}\left(\left(I-\boldsymbol{M}_{\boldsymbol{Z}}\right) \boldsymbol{Y}+\boldsymbol{B}_{g} \hat{\boldsymbol{e}}\right), \text { because } M_{Z}=I-Z\left(Z^{\prime} Z\right)^{-1} Z^{\prime}, \\
& =\left(\boldsymbol{D}^{\prime} \boldsymbol{M}_{\boldsymbol{Z}} \boldsymbol{D}\right)^{-1} \boldsymbol{D}^{\prime}\left(\left(\boldsymbol{M}_{\boldsymbol{Z}}-\boldsymbol{M}_{\boldsymbol{Z}}\right) \boldsymbol{Y}+\boldsymbol{M}_{\boldsymbol{Z}} \boldsymbol{B}_{g} \hat{\boldsymbol{e}}\right), \\
& =\left(\boldsymbol{D}^{\prime} \boldsymbol{M}_{\boldsymbol{Z}} \boldsymbol{D}\right)^{-1} \boldsymbol{D}^{\prime}\left(\boldsymbol{M}_{\boldsymbol{Z}} \boldsymbol{B}_{g} \hat{\boldsymbol{e}}\right), \\
& =\left(\boldsymbol{D}^{\prime} \boldsymbol{M}_{\boldsymbol{Z}} \boldsymbol{D}\right)^{-1} \boldsymbol{D}^{\prime}\left(\boldsymbol{M}_{\boldsymbol{Z}} \boldsymbol{B}_{g} \boldsymbol{M}_{\boldsymbol{Z}} \boldsymbol{Y}\right), \text { because } \hat{\boldsymbol{e}}=\boldsymbol{M}_{\boldsymbol{Z}} \boldsymbol{Y} . \tag{16}
\end{align*}
$$

Kennedy (1995) points out that the Freedman and Lane (1983) algorithm is summarized by Equation (16). Notationally, we can use $T_{\boldsymbol{Z}}(\boldsymbol{Y}, g \boldsymbol{D}) ; g \in \mathscr{G}_{X}\left(\right.$ instead of $\left.T(\boldsymbol{Y}, g \boldsymbol{D}) ; g \in \mathscr{G}_{X}\right)$ to represent the distribution of the test statistic associated with the t-statistic of the $\boldsymbol{D}$ covariate in the Freedman and Lane (1983) regression just described. Using this notation, the analysis of the previous sections holds unaltered.

## E Additional Baseline Tables

Table E. 1 presents the statistical description of retention levels by gender and time of survey. Table E. 2 presents the statistical description of selected pre-program variables after 6 years of the program. Table E. 3 presents the statistical description of selected pre-program variables after 12 years of the program.

Table E.1: Retention Rates by Gender

|  | All Males Groups 2 and 4 |  | Control Males Group 2 |  | Treated Males Group 4 |  | Difference Groups 2 and 4 $p$-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Sample | Attrition | Sample | Attrition | Sample | Attrition |  |
| Month 6 | 333 | 0.93 | 232 | 0.93 | 101 | 0.92 | 0.65 |
| Month 12 | 338 | 0.94 | 234 | 0.94 | 104 | 0.95 | 0.83 |
| Year 2 | 339 | 0.94 | 235 | 0.94 | 104 | 0.95 | 0.95 |
| Year 4.5 | 324 | 0.90 | 223 | 0.90 | 101 | 0.92 | 0.51 |
| Year 6 | 323 | 0.90 | 224 | 0.90 | 99 | 0.90 | 0.99 |
| Year 9 | 315 | 0.88 | 218 | 0.88 | 97 | 0.88 | 0.87 |
| Year 12 | 300 | 0.84 | 202 | 0.81 | 98 | 0.89 | 0.06 |
|  | All Females Groups 2 and 4 |  | Control Females Group 2 |  | Treated Females <br> Group 4 |  | Difference <br> Groups 2 and 4 |
|  | Sample | Attrition | Sample | Attrition | Sample | Attrition | $p$-value |
| Month 6 | 338 | 0.94 | 237 | 0.95 | 101 | 0.90 | 0.072 |
| Month 12 | 347 | 0.96 | 239 | 0.96 | 108 | 0.96 | 0.840 |
| Year 2 | 340 | 0.94 | 235 | 0.94 | 105 | 0.94 | 0.814 |
| Year 4.5 | 322 | 0.89 | 220 | 0.88 | 102 | 0.91 | 0.443 |
| Year 6 | 318 | 0.88 | 220 | 0.88 | 98 | 0.88 | 0.817 |
| Year 9 | 312 | 0.86 | 218 | 0.88 | 94 | 0.84 | 0.354 |
| Year 12 | 294 | 0.81 | 205 | 0.82 | 89 | 0.79 | 0.519 |

Notes: The table presents sample attrition over time. The table displays two panes, the top one focuses on data for males, the bottom one describes data for females. The first column in each panel re are The first column of each panel gives the time of survey. Each panel displays four blocks of data description. The first block on the sub-sample consisting of all data for each gender (Groups 2 and 4), the second block assess the control group (Group 2) and the third block focus on the treatment group (Group 4). First column of each of these three blocks provides the sample size and the second block presents the percentages of non-missing data. The last block presents the double-sided $p$-value for testing whether the the difference of non-missing percentage values between treated and control groups is statistically different than zero.
Table E.2: Descriptive Statistic of Baseline Characteristics (Year 6)

$\begin{array}{lllll}0.277 & 0.448 & 0.366 & 0.484 & 0.116 \\ 0.254 & 0.437 & 0.228 & 0.421 & 0.601 \\ 0.237 & 0.426 & 0.188 & 0.393 & 0.317 \\ 0.134 & 0.341 & 0.168 & 0.376 & 0.434 \\ 0.098 & 0.298 & 0.050 & 0.218 & \mathbf{0 . 0 9 9}\end{array}$


Less than $\$ 3000$
$\$ 3000-\$ 6999$
$\$ 7000-\$ 10999$
Greater than $\$ 11000$
Income, No Response
$\begin{array}{lllll}0.304 & 0.461 & 0.277 & 0.450 & 0.628 \\ 0.205 & 0.405 & 0.198 & 0.400 & 0.879 \\ 0.179 & 0.384 & 0.218 & 0.415 & 0.420 \\ 0.313 & 0.465 & 0.307 & 0.464 & 0.920\end{array}$
$989^{\circ} 0$ OtL"0L 2 ZI8.96 LLE*0L LLE*96
 $\begin{array}{ccccc}100.649 & 10.178 & 99.158 & 10.568 & 0.235 \\ 99.307 & 10.212 & 99.008 & 10.419 & 0.810\end{array}$ $\begin{array}{lllll}100.250 & 10.290 & 99.891 & 9.774 & 0.764\end{array}$ $\begin{array}{llllll}100.090 & 10.390 & 99.448 & 10.715 & 0.615\end{array}$ $\begin{array}{lllll}0.286 & 0.453 & 0.303 & 0.462 & 0.755 \\ 0.179 & 0.384 & 0.232 & 0.424 & 0.282\end{array}$
 $\begin{array}{lllll}0.326 & 0.470 & 0.303 & 0.462 & 0.684\end{array}$


$\begin{array}{lllll}0.295 & 0.456 & 0.290 & 0.455 & 0.905 \\ 0.192 & 0.394 & 0.215 & 0.412 & 0.506 \\ 0.194 & 0.396 & 0.190 & 0.393 & 0.900 \\ 0.319 & 0.467 & 0.305 & 0.462 & 0.719\end{array}$

 $\begin{array}{lllll}100.083 & 10.017 & 99.788 & 9.866 & 0.727\end{array}$ $\begin{array}{lllll}100.065 & 10.213 & 99.535 & 9.992 & 0.537\end{array}$ $\begin{array}{ccccc}100.065 & 10.213 & 99.535 & 9.992 & 0.537 \\ 100.060 & 10.045 & 99.533 & 10.649 & 0.554\end{array}$

| uoss!g |
| :--- |
| Ky! $\mathbf{~ r o u ~}$ | Bisson

Cawthon
Hollywood

| Maternal Health Characteristics |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Maternal Height | 164.557 | 7.253 | 164.064 | 6.569 | 0.397 | 164.331 | 7.404 | 164.472 | 6.546 | 0.865 | 164.781 | 7.108 | 163.651 | 6.601 | 0.170 |
| Pre-Pregnancy Weight | 62.097 | 14.866 | 62.339 | 13.588 | 0.839 | 62.828 | 13.775 | 61.394 | 12.375 | 0.355 | 61.362 | 15.885 | 63.264 | 14.683 | 0.294 |
| Gestational Age (Intake) | 16.560 | 5.794 | 16.630 | 5.728 | 0.887 | 16.402 | 5.746 | 16.364 | 5.596 | 0.955 | 16.719 | 5.850 | 16.891 | 5.870 | 0.807 |
| Maternal Social Support |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Grandmother Social Support | 100.197 | 9.474 | 101.517 | 8.566 | 0.081 | 99.357 | 10.486 | 101.434 | 9.100 | 0.073 | 101.034 | 8.285 | 101.599 | 8.054 | 0.563 |
| Husband/Boyfriend Social Support | 100.030 | 9.994 | 100.704 | 9.754 | 0.421 | 99.892 | 10.057 | 99.907 | 9.524 | 0.990 | 100.169 | 9.952 | 101.484 | 9.960 | 0.272 |
| Maternal Risky Behaviors |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Alcohol Consumption (Past 2 wks ) | 0.043 | 0.202 | 0.050 | 0.218 | 0.680 | 0.036 | 0.186 | 0.071 | 0.258 | 0.228 | 0.049 | 0.217 | 0.030 | 0.171 | 0.385 |
| Smoking (Past 3 days) | 0.085 | 0.279 | 0.110 | 0.314 | 0.334 | 0.081 | 0.273 | 0.121 | 0.328 | 0.284 | 0.089 | 0.286 | 0.099 | 0.300 | 0.784 |
| Used Marijuana (Past 2 wks ) | 0.034 | 0.309 | 0.070 | 0.860 | 0.560 | 0.027 | 0.283 | 30.020 | 0.201 | 10.809 | 0.040 | 0.332 | 0.119 | 1.194 | $40.51^{\circ}$ |
| Used Cocaine (Past 2 wks ) | 0.007 | 0.142 | 0.000 | 0.000 | 0.318 | 0.000 | 0.000 | 0.000 | 0.000 | . | 0.013 | 0.200 | 0.000 | 0.000 | 0.318 |
| Sexually Transmitted Diseases | 0.333 | 0.472 | 0.375 | 0.485 | 0.301 | 0.330 | 0.471 | 0.354 | 0.480 | 0.688 | 0.335 | 0.473 | 0.396 | 0.492 | 0.294 |

 whole sample, the female sample and the male sample. Each block has 6 columns: (1) Control mean (C Mean), (2) Control standard deviation (C SD), (3) Treatment mean (T Mean), (4) Treatment standard deviation (T SD), and (5) Asymptotic $p$-value associated with the difference in means. Bold $p$-values indicate that the t-statistic between the control and the treatment means is significant at the $10 \%$ level.
Table E.3: Descriptive Statistic of Baseline Characteristics (Year 12)
 $\begin{array}{lllll}0.282 & 0.451 & 0.384 & 0.489 & \mathbf{0 . 0 8 3} \\ 0.244 & 0.431 & 0.232 & 0.424 & 0.822 \\ 0.230 & 0.422 & 0.162 & 0.370 & 0.151 \\ 0.139 & 0.347 & 0.162 & 0.370 & 0.606 \\ & & & & 0.25 \\ 0.166\end{array}$ $\begin{array}{lllll}0.105 & 0.308 & 0.061 & 0.240 & 0.166\end{array}$
$0.301-0.460-0.273-0.448 \div 0.603$



| 0.291 | 0.455 | 0.283 | 0.452 | 0.825 |
| :--- | :--- | :--- | :--- | :--- |
| 0.194 | 0.396 | 0.225 | 0.419 | 0.391 |
| 0.204 | 0.403 | 0.188 | 0.392 | 0.657 |
| 0.310 | 0.463 | 0.304 | 0.461 | 0.867 | $\begin{array}{llllll}\varsigma_{0} \varepsilon^{\circ} 0 & \varsigma L t^{\prime} 0 & L \varepsilon \varepsilon^{\prime} 0 & 6 \text { t' }^{\prime} 0 & L L Z Z^{0} 0\end{array}$


\section*{| 0.057 | 0.232 | 0.084 | 0.278 | 0.244 |
| :---: | :---: | :---: | :---: | :---: |
| 0.014 | 0.119 | 0.010 | 0.102 | 0.690 |
| 18.052 | 3.215 | 18.047 | 3.268 | 0.986 |
| 10.254 | 1.860 | 10.073 | 2.025 | 0.296 |
| 0.599 | 0.491 | 0.565 | 0.497 | 0.444 |
| 0.556 | 0.497 | 0.495 | 0.501 | 0.163 |
| 34.800 | 21.380 | 35.727 | 20.185 | 0.606 |
| 0.940 | 0.486 | 1.023 | 0.559 | $\mathbf{0 . 0 8 1}$ |}



$\frac{\text { Total Household Income (Past } 6 \text { Months) }}{\text { Less than } \$ 3000}$ Less than $\$ 3000$
$\$ 3000-\$ 6999$ \$7000-\$10999 Greater than $\$ 11000$
Income, No Response

## Region of Residence

Inner City
Bisson
Cawthon
Hollywood

$\begin{array}{lllll}96.057 & 9.997 & 97.455 & 9.585 & 0.240\end{array}$
 $\begin{array}{ccccc}100.451 & 9.968 & 99.216 & 10.821 & 0.339 \\ 98.973 & 10.197 & 99.186 & 10.440 & 0.866\end{array}$ $\begin{array}{lllll}100.165 & 10.193 & 99.781 & 9.718 & 0.751\end{array}$ $\begin{array}{llllll} & 99.765 & 9.923 & 99.782 & 10.984 & 0.990\end{array}$


 $\begin{array}{llllll}100.640 & 99.954 & 10.301 & 99.085 & 10.533 & 0.507\end{array}$ $\begin{array}{llllll}99.947 & 9.401 & 99.537 & 10.896 & 0.754\end{array}$

 $\begin{array}{llllll}100.059 & 10.236 & 99.446 & 10.098 & 0.489\end{array}$ $\begin{array}{ccccc}100.059 & 10.236 & 99.446 & 10.098 & 0.489 \\ 99.857 & 9.652 & 99.664 & 10.914 & 0.834\end{array}$ $\frac{\text { Maternal Mental Health }}{\text { Maternal IQ (Shipley) }}$ Maternal Bavolek Score Maternal Mental Health Maternal Mastery Maternal Mastery
Maternal Psychological Resources
$\left.\begin{array}{llllllllll}164.297 & 7.483 & 164.904 & 6.664 & 0.485 & & & 164.896 & 7.217 & 163.732\end{array}\right) 6.680 \quad 0.170$ $\begin{array}{llllll}61.701 & 16.178 & 63.530 & 15.003 & 0.332 \\ 16.718 & 5.873 & 16.960 & 5.780 & 0.733\end{array}$

|  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- |
| 101.200 | 8.102 | 101.858 | 7.514 | 0.485 |
| 100.511 | 9.971 | 100.731 | 10.275 | 0.859 | $\begin{array}{llllllllll}100.023 & 9.949 & 99.833 & 9.700 & 0.877 & 100.511 & 9.971 & 100.731 & 10.275 & 0.859\end{array}$



## Grandmother Social Support Husband/Boyfriend Social Support

| Maternal Riskey Behaviors |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Alcohol Consumption (Past 2 wks ) | 0.040 | 0.197 | 0.047 | $0.212 \quad 0.710$ | 0.038 | 0.191 | 0.065 | 0.248 | 0.345 | 0.043 | 0.203 | 0.030 | 0.172 | 0.568 |
| Smoking (Past 3 days) | 0.081 | 0.273 | 0.105 | 0.3070 .356 | 0.075 | 0.265 | 0.120 | 0.326 | 0.255 | 0.086 | 0.281 | 0.091 | 0.289 | 0.891 |
| Used Marijuana (Past 2 wks ) | 0.036 | 0.318 | 0.073 | $0.880 \quad 0.566$ | 0.028 | 0.291 | 0.022 | 0.209 | 0.824 | 0.043 | 0.344 | 0.121 | 1.206 | 0.528 |
| Used Cocaine (Past 2 wks ) | 0.007 | 0.146 | 0.000 | $0.000 \quad 0.318$ | 0.000 | 0.000 | 0.000 | 0.000 | . | 0.014 | 0.208 | 0.000 | 0.000 | 0.318 |
| Sexually Transmitted Diseases | 0.347 | 0.477 | 0.372 | 0.4850 .554 | 0.335 | 0.473 | 0.337 | 0.475 | 0.972 | 0.359 | 0.481 | 0.404 | 0.493 | 0.450 |

 Maternal Height
Pre-Pregnancy Weight
Gestational Age (Intake)
Notes: This table presents the statistical description of selected pre-program variables after 6 years of the program. The first column of the table gives the variable
 whole sample, the female sample and the male sample. Each block has 6 columns: (1) Control mean (C Mean), (2) Control standard deviation (C SD), (3) Treatment mean the control and the treatment means is significant at the $10 \%$ level.

## F Additional Inference Results: Unconditional Analysis and Addressing Attrition using Inverse Propensity Weights

Tables F.1-F. 5 present the unconditional analysis of the treatment effects presented in Tables $6-10$ of the main paper.

One aspect of the NFP that may cause concern is attrition. In order to address this issue, we use statistical models that account for missing data by reweighting observations according to the inverse probability of retention, which is usually termed Inverse Probability Weighting (IPW). The probabilities of attrition at each wave are estimated by gender using logit models. To select the covariates in the model, we choose the set of pre-program covariates that minimize the Akaike Information Criterion (AIC). Then, we use the estimated probabilities to reweight the observations and compute the treatment effects. The results do not change much after this correction. Tables F.7-F. 10 show these results. The tables can be read in the same way as Tables 6-10 in the paper.
Table F.1: Child Health Outcomes (Unconditional Effects)

| Outcome DescriptionBirth Outcomes for Child | Female Sample |  |  |  |  |  | Male Sample |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Control Mean | Difference in Means | Effect Size | Asymp. $p$-value | Single $p$-value | Stepdown $p$-value | Control <br> Mean | Difference in Means | Effect Size | Asymp. $p$-value | Single $p$-value | Stepdown $p$-value |
| Placenta Weight | 682.995 | -2.290 | -0.014 | 0.544 | 0.541 | 0.782 | 663.819 | 21.423 | 0.113 | 0.168 | 0.149 | 0.149 |
| Birth Weight | 3055.224 | -121.268 | -0.219 | 0.961 | 0.952 | 0.952 | 2997.486 | 199.869 | 0.275 | 0.006 | 0.006 | 0.025 |
| Head Circumference | 33.262 | 0.023 | 0.013 | 0.454 | 0.452 | 0.771 | 33.511 | 0.352 | 0.151 | 0.088 | 0.080 | 0.142 |
| Length | 49.665 | 0.208 | 0.076 | 0.259 | 0.253 | 0.595 | 49.918 | 0.567 | 0.150 | 0.083 | 0.061 | 0.145 |
| Gestational Age at Delivery | 39.119 | -0.415 | -0.186 | 0.899 | 0.870 | 0.935 | 38.544 | 0.783 | 0.220 | 0.019 | 0.019 | 0.062 |
| Child Health Outcomes (Year 12) |  |  |  |  |  |  |  |  |  |  |  |  |
| Any Injuries Since Last Interview | -0.164 | 0.065 | 0.176 | 0.069 | 0.064 | 0.162 | -0.224 | 0.061 | 0.146 | 0.109 | 0.109 | 0.385 |
| \# Hospitalizations for Injuries Since Last Interview | -0.009 | 0.009 | 0.097 | 0.178 | 0.404 | 0.404 | -0.010 | 0.010 | 0.099 | 0.164 | 0.023 | 0.141 |
| Total \# Injuries Since Last Interview | -0.197 | 0.098 | 0.199 | 0.039 | 0.029 | 0.089 | -0.268 | 0.054 | 0.099 | 0.209 | 0.222 | 0.610 |
| Hospitalized Since Last Interview | -0.042 | 0.042 | 0.210 | 0.023 | 0.027 | 0.111 | -0.039 | -0.033 | -0.171 | 0.892 | 0.871 | 0.995 |
| Have Chronic Condition/Health Problem | -0.197 | 0.012 | 0.031 | 0.401 | 0.404 | 0.639 | -0.361 | -0.072 | -0.150 | 0.886 | 0.886 | 0.986 |
| Standardized Child BMI | -1.121 | 0.276 | 0.308 | 0.007 | 0.007 | 0.034 | -0.797 | -0.163 | -0.179 | 0.921 | 0.917 | 0.917 |

Note: The first column provides the outcome description. Our results are presented in six columns for each gender. The first column (Control Mean) of each result set shows the unconditional mean for the control group. When factor scores were computed, we set the mean in the control group to zero. The second column (Difference in Means)
gives the unconditional difference in means between the treatment group and the control group. As mentioned in Section 2, the control group stands for the 2 of the NFP experiment and the treatment group stands for the original group 4. The third column (Effect Size) presents the unconditional effect size for the respective group. The fourth column (Asymp. $p$-value) provides the asymptotic $p$-value for the one-sided single hypothesis test associated with the $t$-statistic for the unconditional difference in means between treatment and control groups. Th frth column (Unrestricted Permutation-Single $p$-value) presents the one-sided unrestricted permutation $p$-values for the that account for multiple-hypothesis testing based on the Stepdown procedure of Romano and wormilar meaning of outcomes that are tested jointly are separate allude to mothers whose first child is a girl. Likewise, male maternal outcomes alludes to mothers whose first child is a boy.
Table F.2: Family Environment (Unconditional Effects)

| Outcome Description | Female Sample |  |  |  |  |  | Male Sample |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Basic St | istics |  | Unrestricted Permutation <br> Single Stepdown <br> $p$-value $p$-value |  | Control Mean | Basic Statistics |  | Asymp. $p$-value | Unrestricted Permutation <br> Single Stepdown <br> $p$-value $p$-value |  |
|  | Control Mean | Difference in Means | Effect Size | Asymp. $p$-value |  |  | Difference in Means | Effect Size |  |  |  |
| Home Environment, Parenting (Year 1) |  |  |  |  |  |  |  |  |  |  |  |  |
| Family Environment (HOME Score) | 0.000 | 0.354 | 0.354 | 0.002 | 0.002 | 0.004 |  | 0.000 | 0.207 | 0.207 | 0.044 | 0.044 | 0.044 |
| Non-Abusive Parenting Attitudes (Bavolek) | 0.000 | 0.289 | 0.289 | 0.009 | 0.012 | 0.012 | 0.000 | 0.274 | 0.274 | 0.012 | 0.012 | 0.024 |
| Home Environment, Parenting (Year 2) |  |  |  |  |  |  |  |  |  |  |  |  |
| Family Environment (HOME Score) | 0.000 | 0.302 | 0.302 | 0.007 | 0.008 | 0.008 | 0.000 | 0.170 | 0.170 | 0.082 | $0.087$ | 0.087 |
| Non-Abusive Parenting Attitudes (Bavolek) | 0.000 | 0.371 | 0.371 | 0.002 | 0.004 | 0.008 | 0.000 | 0.316 | 0.316 | 0.004 | $0.004$ | 0.009 |
| Maternal Mental Health (Year 2) - Factor Scores |  |  |  |  |  |  |  |  |  |  |  |  |
| Anxiety | 0.000 | 0.247 | 0.247 | 0.024 | 0.030 | 0.072 | 0.000 | 0.038 | 0.038 | 0.376 | 0.382 | 0.561 |
| Depression | 0.000 | 0.129 | 0.129 | 0.139 | 0.137 | 0.221 | 0.000 | 0.062 | 0.062 | 0.302 | 0.306 | 0.516 |
| Positive Well-Being | 0.000 | 0.101 | 0.101 | 0.197 | 0.199 | 0.199 | 0.000 | -0.136 | -0.136 | 0.865 | 0.859 | 0.859 |
| Emotional Stability | 0.000 | 0.207 | 0.207 | 0.047 | 0.052 | 0.106 | 0.000 | 0.050 | 0.050 | 0.340 | 0.341 | 0.526 |
| Overall Mental Health | 0.000 | 0.210 | 0.210 | 0.043 | 0.046 | 0.098 | 0.000 | -0.014 | -0.014 | 0.546 | 0.554 | 0.658 |
| Self-Esteem | 0.000 | 0.313 | 0.313 | 0.005 | 0.006 | 0.024 | 0.000 | 0.073 | 0.073 | 0.282 | 0.295 | 0.563 |
| Mastery | 0.000 | 0.286 | 0.286 | 0.012 | 0.019 | 0.058 | 0.000 | 0.198 | 0.198 | 0.053 | 0.059 | 0.182 |
| Welfare (Child Ages 1-12 Years) |  |  |  |  |  |  |  |  |  |  |  |  |
| AFDC/TANF | -2744.043 | -46.414 | -0.017 | 0.557 | 0.549 | 0.549 | -2743.386 | 439.910 | 0.159 | 0.074 | 0.074 | 0.074 |
| Food Stamp | -2996.965 | 164.090 | 0.089 | 0.218 | 0.208 | 0.303 | -3263.273 | 347.770 | 0.202 | 0.039 | 0.038 | 0.072 |
| Medicaid | -3543.761 | 167.036 | 0.090 | 0.217 | 0.211 | 0.286 | -3823.048 | 317.413 | 0.191 | 0.048 | 0.045 | 0.073 |

Note: The first column provides the outcome description. Our results are presented in six columns for each gender. The first column (Control Mean) of each result set shows the unconditional mean for the control group. When factor scores were computed, we set the mean in the control group to zero. The second column (Difference in Means)
 experiment and the traument group stands for the original group 4. The third column (becs size) presents the unconditional effect size for the respective group. Teans between treatment and control groups. The fifth column (Unrestricted Permutation - Single $p$-value) presents the one-sided unrestricted permutation $p$-values for the single-hypothesis testing based on the $t$-statistic associated with the treatment indicator. Finally, the last column (Unrestricted Permutation - Stepdown) provides $p$-values that account for multiple-hypothesis testing based on the Stepdown procedure of Romano and Wolf (2005). Blocks of outcomes that are tested jointly are separated by lines. The selection of blocks of outcomes is done on the basis of their meaning. Outcomes that share similar meaning are grouped together. Female maternal outcomes allude to
mothers whose first child is a girl. Likewise, male maternal outcomes alludes to mothers whose first child is a boy.
Table F.3: Maternal cumulative subsequent births (Unconditional Analysis)

| Outcome Description | Female Sample |  |  |  |  |  | Male Sample |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Basic Statistics |  |  |  | Unrestricted Permutation <br> Single Stepdown <br> $p$-value $p$-value |  | $\begin{aligned} & \text { Control } \\ & \text { Mean } \\ & \hline \end{aligned}$ | Basic Statistics |  | Asymp | Unrestricted Permutation |  |
|  | Control Mean | Difference in Means | Effect <br> Size | Asymp. $p$-value |  |  | Difference in Means | Effect Size | Asymp. $p$-value | Single $p$-value | $\begin{gathered} \text { Stepdown } \\ p \text {-value } \\ \hline \end{gathered}$ |
| Cumulative Subsequent Births (Years 2-12) |  |  |  |  |  |  |  |  |  |  |  |  |
| Subsequent Children Birth (Years 9-12) | -0.259 | -0.044 | -0.073 | 0.713 | 0.725 | 0.918 |  | -0.397 | 0.014 | 0.020 | 0.437 | 0.446 | 0.659 |
| Subsequent Children Birth (Years 6-9)", | -0.344 | -0.177 | -0.300 | 0.991 | 0.992 | 0.992 | -0.459 | -0.046 | -0.064 | 0.699 | 0.700 | 0.700 |
| Subsequent Children Birth (Years 2-6)", | -0.884 | -0.034 | -0.040 | 0.630 | 0.627 | 0.939 | -1.027 | 0.138 | 0.159 | 0.087 | 0.079 | 0.206 |
| Subsequent Children Birth (Years 0-2)" | -0.298 | 0.050 | 0.110 | 0.171 | 0.173 | 0.532 | -0.315 | 0.105 | 0.226 | 0.023 | 0.023 | 0.090 |

 the unconditional mean for the control group. When factor scores were computed, we set the mean in the control group to zero. The second column (Difference in Means) experiment and the treatment group stands for the original group 4. The third column (Effect Size) presents the unconditional effect size for the respective group. The fourth column (Asymp. p-value) provides the asymptotic $p$-value for the one-sided single hypothesis test associated with the $t$-statistic for the unconditional difference in mean between treatment and control groups. The fifth column (Unrestricted Permutation - Single $p$-value) presents the one-sided unrestricted permutation $p$-values for the bic associan ( that account for multiple-hypothesis testing based on the Stepdown procedure of Romano and Wolf (2005). Blocks of outcomes that are tested jointly are separated by lines. mothers whose first child is a girl. Likewise, male maternal outcomes alludes to mothers whose first child is a boy.
Table F.4: Cognitive Abilities and Achievement Outcomes (Unconditional Effects)

| Outcome Description <br> Kaufman Assessment for Children (Year 6) | Female Sample |  |  |  |  |  | Male Sample |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Control Mean | Basic St Difference in Means | istics Effect Size | Asymp. $p$-value | Unrestricted Single $p$-value | Permutation Stepdown $p$-value | Control Mean | Basic St Difference in Means | Effect Size | Asymp. $p$-value | Unrestrict Single $p$-value | Permutation Stepdown $p$-value |
| Gestalt Closure | 8.981 | 0.172 | 0.055 | 0.330 | 0.332 | 0.690 | 9.775 | -0.456 | -0.154 | 0.888 | 0.885 | 0.885 |
| Hand Movements | 9.282 | 0.351 | 0.158 | 0.106 | 0.110 | 0.454 | 9.287 | 0.043 | 0.019 | 0.438 | 0.447 | 0.763 |
| Matrix Analogies | 8.632 | 0.082 | 0.045 | 0.356 | 0.355 | 0.643 | 8.478 | 0.170 | 0.105 | 0.204 | 0.204 | 0.573 |
| Number Recall | 9.423 | 0.169 | 0.058 | 0.320 | 0.323 | 0.719 | 8.952 | 0.659 | 0.264 | 0.029 | 0.041 | 0.222 |
| Photo Series | 6.967 | 0.343 | 0.156 | 0.108 | 0.112 | 0.423 | 6.774 | -0.063 | -0.028 | 0.589 | 0.586 | 0.787 |
| Spatial Memory | 8.434 | 0.117 | 0.047 | 0.357 | 0.361 | 0.537 | 8.526 | 0.276 | 0.110 | 0.195 | 0.199 | 0.601 |
| Triangles | 8.868 | 0.193 | 0.082 | 0.253 | 0.255 | 0.672 | 9.120 | 0.078 | 0.032 | 0.400 | 0.401 | 0.773 |
| Word Order | 9.693 | -0.193 | -0.069 | 0.713 | 0.709 | 0.709 | 9.191 | 0.542 | 0.203 | 0.059 | 0.063 | 0.290 |
| Kaufman Assessment for Children (Year 6) |  |  |  |  |  |  |  |  |  |  |  |  |
| Nonverbal | 89.203 | 1.409 | 0.148 | 0.127 | 0.144 | 0.253 | 89.244 | 0.778 | 0.079 | 0.269 | 0.275 | 0.333 |
| Sequential Processing | 96.582 | 0.714 | 0.054 | 0.335 | 0.341 | 0.341 | 94.507 | 2.339 | 0.193 | 0.070 | 0.080 | 0.158 |
| Simultaneous Processing | 88.844 | 1.268 | 0.117 | 0.180 | 0.189 | 0.300 | 89.919 | 0.180 | 0.017 | 0.447 | 0.446 | 0.446 |
| WISC-III, PPVT-III for Children (Year 6) |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  |  | 0.312 |  |  |  |  | 0.018 |  | 0.441 | 0.441 |
| Peabody Picture Vocabulary Test (PPVT-III) | $83.299$ | $0.508$ | $0.040$ | $0.373$ | 0.373 | 0.373 | 82.466 | 1.534 | $0.135$ | 0.158 | 0.176 | 0.292 |
| Child Cognition (Year 6) - Factor Scores |  |  |  |  |  |  |  |  |  |  |  |  |
| Cognition + Achievement (KABC, PPVT, WISC) | 0.000 | 0.109 | 0.109 | 0.197 | 0.209 | 0.209 | 0.000 | 0.187 | 0.187 | 0.074 | 0.076 | 0.076 |
| Cognitive skills (Mental Processing KABC) | 0.000 | 0.119 | 0.119 | 0.174 | 0.186 | 0.242 | 0.000 | 0.270 | 0.270 | 0.019 | 0.021 | 0.029 |
| Reading Achievement for the Child (Year 12) |  |  |  |  |  |  |  |  |  |  |  |  |
| Average Reading Grade (Grades 1-5) | 2.703 | -0.028 | -0.036 | 0.606 | 0.607 | 0.840 | 2.327 | 0.106 | 0.133 | 0.159 | 0.153 | 0.329 |
| TCAP \% Language (School Years 1-5, Grd 3+) | 50.854 | -2.399 | -0.099 | 0.760 | 0.756 | 0.879 | 38.063 | 5.143 | 0.224 | 0.055 | 0.053 | 0.161 |
| TCAP \% Reading (School Years 1-5, Grd 3+) | 41.607 | -1.699 | -0.080 | 0.717 | 0.712 | 0.870 | 34.912 | 2.116 | 0.099 | 0.236 | 0.234 | 0.391 |
| PIAT Total Reading (Derived Score) | 90.246 | -0.405 | -0.040 | 0.619 | 0.615 | 0.812 | 89.292 | 1.158 | 0.084 | 0.250 | 0.242 | 0.300 |
| PIAT Reading Comprehension (Derived Score) | 88.307 | -1.091 | -0.114 | 0.811 | 0.805 | 0.805 | 87.585 | 2.108 | 0.172 | 0.091 | 0.092 | 0.230 |
| PIAT Reading Recognition (Derived Score) | 94.221 | 0.870 | 0.069 | 0.306 | 0.310 | 0.583 | 92.456 | 0.752 | 0.050 | 0.344 | 0.339 | 0.339 |
| Math Achievement for the Child (Year 12) |  |  |  |  |  |  |  |  |  |  |  |  |
| Average Math Grade (Grades 1-5) | 2.634 | -0.021 | -0.025 | 0.577 | 0.583 | 0.734 | 2.368 | 0.149 | 0.184 | 0.078 | 0.071 | 0.119 |
| TCAP \% Math (School Years 1-5, Grd 3+) | $46.935$ | -0.115 | $-0.005$ | 0.514 | 0.513 | 0.724 | 40.346 | 3.749 | 0.161 | 0.126 | 0.128 | 0.128 |
| PIAT Mathematics (Derived Score) | 87.188 | -0.790 | -0.080 | 0.728 | 0.724 | 0.724 | 86.316 | 2.102 | 0.198 | 0.062 | 0.065 | 0.150 |

[^2]Table F.5: Socio-emotional Outcomes (Unconditional Effects)

| Outcome Description Child Behavior Checklist (Year 2) - Factor Scores | Female Sample |  |  |  |  |  | Male Sample |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Control Mean | Basic S Difference in Means | Effect Size | Asymp. $p$-value | $\begin{array}{cc}\text { Unrestricted } & \text { Permutation } \\ \text { Single } & \text { Stepdown } \\ p \text {-value } & p \text {-value }\end{array}$ |  | Control Mean | Difference in Means |  |  | Unrestricted Permutation <br> Single Stepdown <br> $p$-value $p$-value |  |
|  |  |  |  |  |  |  | Effect Size |  | Asymp. $p$-value |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |
| Affective Problems Anxiety Problems <br> Pervasion Developmental Problems <br> Attention Deficit Hyperactivity Disorder Oppositional Defiant Problems <br> Child Behavior Checklist (Year 6) - Factor Scores | 0.000 | 0.336 | 0.336 | 0.001 | 0.000 | 0.002 | 0.000 | -0.163 | -0.163 | 0.903 | 0.903 | 0.903 |
|  | 0.000 | 0.191 | 0.191 | 0.048 | 0.040 | 0.040 | 0.000 | 0.029 | 0.029 | 0.407 | 0.421 | 0.758 |
|  | 0.000 | 0.262 | 0.262 | 0.010 | 0.007 | 0.023 | 0.000 | -0.084 | -0.084 | 0.757 | 0.763 | 0.924 |
|  | 0.000 | 0.239 | 0.239 | 0.021 | 0.016 | 0.041 | 0.000 | -0.078 | -0.078 | 0.734 | 0.736 | 0.935 |
|  | 0.000 | 0.224 | 0.224 | 0.029 | 0.026 | 0.047 | 0.000 | -0.113 | -0.113 | 0.832 | 0.847 | 0.947 |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| Affective ProblemsAnxiety ProblemsSomatic ProblemsAttention Deficit Hyperactivity ProblemsOppositional Defiant ProblemsConduct Problems | 0.000 | 0.063 | 0.063 | 0.310 | 0.331 | 0.631 | 0.000 | 0.115 | 0.115 | 0.164 | 0.161 | 0.503 |
|  | 0.000 | 0.085 | 0.085 | 0.230 | 0.220 | 0.515 | 0.000 | -0.106 | -0.106 | 0.797 | 0.797 | 0.797 |
|  | 0.000 | -0.084 | -0.084 | 0.745 | 0.739 | 0.739 | 0.000 | -0.061 | -0.061 | 0.684 | 0.692 | 0.876 |
|  | 0.000 | 0.269 | 0.269 | 0.013 | 0.013 | 0.056 | 0.000 | 0.053 | 0.053 | 0.332 | 0.323 | 0.690 |
|  | 0.000 | 0.031 | 0.031 | 0.398 | 0.400 | 0.617 | 0.000 | 0.103 | 0.103 | 0.208 | 0.220 | 0.578 |
|  | 0.000 | 0.269 | 0.269 | 0.010 | 0.009 | 0.044 | 0.000 | 0.021 | 0.021 | 0.430 | 0.441 | 0.766 |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| Dysregulated AggressionWarmth and EmpathyEmotional IntegrationPerformance AnxietyAggression | 0.000 | 0.040 | 0.040 | 0.366 | 0.354 | 0.798 | 0.000 | 0.179 | 0.179 | 0.096 | 0.126 | 0.465 |
|  | 0.000 | 0.360 | 0.360 | 0.003 | 0.004 | 0.017 | 0.000 | -0.097 | -0.097 | 0.779 | 0.786 | 0.786 |
|  | 0.000 | -0.045 | -0.045 | 0.640 | 0.639 | 0.639 | 0.000 | 0.022 | 0.022 | 0.433 | 0.433 | 0.851 |
|  | 0.000 | 0.037 | 0.037 | 0.375 | 0.359 | 0.649 | 0.000 | -0.051 | -0.051 | 0.644 | 0.622 | 0.903 |
|  | 0.000 | 0.182 | 0.182 | 0.057 | 0.050 | 0.164 | 0.000 | 0.151 | 0.151 | 0.133 | 0.157 | 0.510 |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| Presence of Internalizing Disorders Presence of Externalizing Disorders Average \# of Absences (School Years 1-5) | -0.239 | 0.046 | 0.107 | 0.197 | 0.202 | 0.480 | -0.403 | 0.099 | 0.201 | 0.053 | 0.052 | 0.097 |
|  | -0.182 | 0.023 | 0.060 | 0.317 | 0.328 | 0.545 | -0.187 | -0.069 | -0.177 | 0.908 | 0.907 | 0.907 |
|  | -10.186 | -0.996 | -0.131 | 0.844 | 0.845 | 0.845 | -11.803 | 1.941 | 0.247 | 0.021 | 0.022 | 0.061 |

[^3]Table F.6: Using Logit to Obtain the Inverse Probability Weights

| Sample | $\begin{gathered} \dot{\sim} \\ \sim \\ \sim \\ \dot{\sim} \\ \text { ín } \end{gathered}$ | $\begin{aligned} & 8 \\ & 0 \\ & \text { O } \\ & \text { U } \\ & \hline \end{aligned}$ | $\begin{aligned} & \ddot{0} \\ & \frac{0}{0} \\ & \text { a } \\ & 0 \end{aligned}$ | $$ | $\begin{gathered} \text { U } \\ \text { تِ } \\ \hline \end{gathered}$ | 5 0 0 0 0 B E 0 2 |  | (1) | $\begin{aligned} & \dot{\text { ® }} \\ & \text { I } \\ & \text { In } \\ & \text { In } \end{aligned}$ | $\underset{\substack{0 \\ 60 \\ \hline}}{ }$ | pəuenbs ${ }^{26} \mathrm{~V}$ | $\underset{\sim}{\infty}$ | $\begin{aligned} & \underset{\sim}{\infty} \\ & \underset{\sim}{\infty} \end{aligned}$ |  | $\begin{aligned} & \text { i } \\ & \text { 出 } \\ & \text { 会 } \end{aligned}$ |  |  | $\begin{aligned} & \dot{0} \\ & \mathbb{y y} \\ & \dot{N} \\ & \dot{0} \end{aligned}$ |  |  |  | LR Chi2 | Prob. > Chi2 | AIC | BIC |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Treatment Females |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Year 12 | X |  |  |  | X |  |  |  |  | X | X | X |  | X | X | X | X | X | X | X | X | 25.79 | 0.06 | 45.26 | 90.70 |
| Year 9 | X |  |  |  |  | X |  |  |  | X | X |  |  |  | X | X | X |  | X | X | X | 22.53 | 0.05 | 51.39 | 88.81 |
| Year 6 |  |  |  |  | X | X | X | X |  | X |  |  |  |  |  | X |  |  | X | X | X | 17.62 | 0.02 | 30.361 | 54.417 |
| Year 4.5 |  | X |  |  | X | X |  | X |  | X |  |  |  |  |  | X |  |  | X | X | X | 16.67 | 0.05 | 21.291 | 48.019 |
| Year 2 | X |  |  |  | X | X |  |  |  | X | X | X | X | X | X | X |  |  |  |  |  | 24.18 | 0.01 | -20.92 | 8.79 |
| Control Females |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Year 12 | X | X |  |  |  | X |  | X |  | X |  | X |  |  | X |  |  |  |  | X |  | 20.31 | 0.01 | 205.54 | 237.16 |
| Year 9 |  | X |  |  |  | X |  | X |  | X | X | X |  |  |  |  |  |  | X | X |  | 19.29 | 0.01 | 122.33 | 153.81 |
| Year 6 |  |  |  |  |  |  | X | X |  | X | X | X |  |  |  | X |  |  | X |  |  | 17.84 | 0.01 | 103.72 | 131.73 |
| Year 4.5 |  |  |  |  | X |  |  |  |  | X | X |  | x |  | X |  |  | X |  |  | X | 14.68 | 0.02 | 114.54 | 139.16 |
| Year 2 |  |  |  |  |  |  |  |  |  | X | X |  |  | X | X |  | X |  |  | X | X | 10.00 | 0.12 | -19.62 | 5.01 |
| Treatment Males |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Year 12 |  |  |  |  | X |  |  | X |  |  |  | X | X | X |  | X | X | X | X |  | X | 22.97 | 0.04 | 106.90 | 144.96 |
| Year 9 |  |  |  |  |  |  |  | X |  | X | X |  |  | X |  |  |  |  |  |  |  | 9.38 | 0.05 | 82.55 | 96.14 |
| Year 6 |  |  |  |  | X |  |  | X |  | X | X |  | X | X |  | X |  |  | X |  |  | 17.66 | 0.02 | 62.91 | 87.37 |
| Year 4.5 |  | X |  |  | X |  |  | X |  | X | X | X |  | X | X |  |  |  | X | X | X | 14.03 | 0.23 | 46.04 | 78.66 |
| Year 2 |  |  |  |  |  | X |  |  |  |  |  | X |  | X |  | X |  |  | X | X |  | 15.73 | 0.02 | -3.54 | 15.49 |
| Control Males |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Year 12 | X |  |  |  | X | X |  | X | X | X |  |  |  |  |  |  | X |  | X | X |  | 32.07 | 0.00 | 155.12 | 200.37 |
| Year 9 |  |  |  |  | X |  |  | X |  |  | X |  |  | X |  |  | X |  | X | X |  | 37.02 | 0.00 | 95.72 | 134.01 |
| Year 6 |  |  |  |  | X |  |  | X |  |  | X |  |  |  | X |  | X |  | X | X | X | 34.24 | 0.00 | 68.20 | 109.91 |
| Year 4.5 |  |  | X | X |  | x |  | X | X |  |  | X |  | X | X | X |  |  | X |  |  | 22.03 | 0.02 | 121.86 | 160.14 |
| Year 2 |  |  |  |  |  |  |  |  |  | X | X |  |  | X | X |  |  |  |  |  |  | 12.38 | 0.01 | -24.20 | -6.63 |

 control females, treatment males, and control males. Additionally, there is a corresponding time period for each row. The next 23 columns represents the set of pre-program
 information criterion and Bayesian information criterion respectively.
Table F.7: Child Health Outcomes (Correcting for Attrition)

| Outcome Description | Females |  |  |  |  |  | Males |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Basic Statistics |  |  |  | Block Perm. FL |  | Basic Statistics |  |  |  | Block Perm. FL |  |
|  | Cntr. Mean | Cd. Diff. Mn. | Cd. Eff. Size | Asy P-val | Single P-val | Stepdown | Cntr. Mean | Cd. Diff. Mn. | Cd. Eff. Size | Asy P-val | Single P-val | Stepdown |
| Birth Outcomes for Child |  |  |  |  |  |  |  |  |  |  |  |  |
| Placenta Weight | 683.488 | -11.638 | -0.073 | 0.707 | 0.467 | 0.717 | 662.401 | 27.965 | 0.157 | 0.112 | 0.014 | 0.036 |
| Birth Weight | 3050.565 | -128.456 | -0.235 | 0.966 | 0.903 | 0.903 | 2993.726 | 204.977 | 0.292 | 0.006 | 0.000 | 0.001 |
| Head Circumference | 33.257 | 0.038 | 0.023 | 0.425 | 0.203 | 0.459 | 33.506 | 0.327 | 0.146 | 0.107 | 0.060 | 0.060 |
| Length | 49.652 | 0.234 | 0.087 | 0.236 | 0.202 | 0.513 | 49.908 | 0.711 | 0.196 | 0.042 | 0.018 | 0.033 |
| Gestational Age at Delivery | 39.092 | -0.545 | -0.242 | 0.940 | 0.854 | 0.919 | 38.526 | 0.745 | 0.214 | 0.028 | 0.001 | 0.005 |
| Child Health Outcomes (Year 12) |  |  |  |  |  |  |  |  |  |  |  |  |
| Any Injuries Since Last Interview | 0.175 | -0.043 | -0.122 | 0.171 | 0.216 | 0.386 | 0.232 | -0.059 | -0.149 | 0.132 | 0.120 | 0.474 |
| \# Hospitalizations for Injuries Since Last Interview | 0.009 | -0.011 | -0.116 | 0.138 | 0.185 | 0.451 | 0.011 | -0.013 | -0.134 | 0.132 | 0.170 | 0.582 |
| Total \# Injuries Since Last Interview | 0.200 | -0.068 | -0.156 | 0.102 | 0.074 | 0.224 | 0.278 | -0.057 | -0.110 | 0.212 | 0.268 | 0.685 |
| Hospitalized Since Last Interview | 0.059 | -0.044 | -0.226 | 0.033 | 0.035 | 0.140 | 0.040 | 0.054 | 0.299 | 0.975 | 0.890 | 0.890 |
| Have Chronic Condition/Health Problem | 0.203 | -0.003 | -0.009 | 0.473 | 0.639 | 0.639 | 0.360 | 0.077 | 0.163 | 0.885 | 0.849 | 0.965 |
| Standardized Child BMI | 1.090 | -0.240 | -0.277 | 0.019 | 0.012 | 0.060 | 0.778 | 0.224 | 0.257 | 0.968 | 0.833 | 0.988 |

 the unconditional mean for the control group. When factor scores were computed, we set the mean in the control group to zero. The second column (Difference in Means) gives the conditional difference in means between the treatment group and the control group. As mentioned in Section 2 , the control group stands for the original treatment group. The fourth column (Asymp. $p$-value) provides the asymptotic $p$-value for the one-sided single hypothesis test associated with the $t$-statistic for the conditional difference in means between treatment and control groups. The fifth column (Block Permutation - Single $p$-value) presents the one-sided restricted permutation $p$-values for the single-hypothesis testing based on the $t$-statistic associated with the treatment indicator in the Freedman and Lane (1983) regression as described in Section 3 . By restricted permutation we mean that permutations are done within strata defined by the baseline variables used in the randomization protocol: maternal age and race, gestational age at enrollment, employment status of the head of the household, and geographic region. The covariates used in the Freedman and Lane (1983) regression are: maternal height, household income, grandmother support, maternal parenting attitudes and mother currently in school. Finally, the last column (Block Permutation Stepdown) provides $p$-values that account for multiple-hypothesis testing based on the Stepdown procedure of Romano and Wolf (2005). Blocks of outcomes that are tested jointly are separated by lines. The selection of blocks of outcomes is done on the basis of their meaning. Outcomes that share similar meaning are grouped together. Female
 that is described at the beginning of this section.
Table F.8: Family Environment (Correcting for Attrition)

| Outcome Description | Females |  |  |  |  |  | Males |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Basic Statistics |  |  |  | Block Perm. FL |  | Basic Statistics |  |  |  | Block Perm. FL |  |
|  | Cntr. Mean | Cd. Diff. Mn. | Cd. Eff. Size | Asy P-val | Single P-val | Stepdown | Cntr. Mean | Cd. Diff. Mn. | Cd. Eff. Size | Asy P-val | Single P-val | Stepdown |
| Home Environment, Parenting (Year 1) - Factor Scores |  |  |  |  |  |  |  |  |  |  |  |  |
| Home Observation Measurement of the Environment (HOME) | 0.000 | 0.338 | 0.338 | 0.004 | 0.004 | 0.004 | -0.005 | 0.154 | 0.154 | 0.114 | 0.079 | 0.079 |
| Non-Abussive Parenting Attitudes (Bavolek) | 0.007 | 0.294 | 0.293 | 0.010 | 0.003 | 0.006 | -0.002 | 0.364 | 0.364 | 0.002 | 0.001 | 0.002 |
| Home Environment, Parenting (Year 2)- Factor Scores |  |  |  |  |  |  |  |  |  |  |  |  |
| Home Observation Measurement of the Environment (HOME) | 0.001 | 0.298 | 0.297 | 0.010 | 0.004 | 0.007 | -0.008 | 0.116 | 0.116 | 0.186 | 0.111 | 0.111 |
| Non-Abussive Parenting Attitudes (Bavolek) | 0.012 | 0.374 | 0.372 | 0.003 | 0.005 | 0.005 | -0.005 | 0.481 | 0.481 | 0.000 | 0.001 | 0.001 |
| Maternal Mental Health (Year 2) |  |  |  |  |  |  |  |  |  |  |  |  |
| Anxiety | -0.001 | -0.226 | -0.226 | 0.042 | 0.038 | 0.086 | 0.012 | -0.052 | -0.052 | 0.340 | 0.348 | 0.633 |
| Depression | 0.000 | -0.115 | -0.115 | 0.180 | 0.102 | 0.169 | 0.010 | -0.011 | -0.011 | 0.465 | 0.524 | 0.692 |
| Positive Well-Being | -0.002 | 0.096 | 0.096 | 0.222 | 0.413 | 0.413 | -0.006 | -0.213 | -0.214 | 0.950 | 0.947 | 0.947 |
| Emotional Stability | 0.001 | 0.185 | 0.185 | 0.076 | 0.056 | 0.113 | -0.012 | 0.042 | 0.042 | 0.367 | 0.427 | 0.689 |
| Overall Mental Health | 0.000 | 0.193 | 0.193 | 0.066 | 0.066 | 0.122 | -0.011 | -0.047 | -0.047 | 0.644 | 0.666 | 0.772 |
| Self-Esteem | 0.011 | 0.283 | 0.283 | 0.014 | 0.003 | 0.014 | -0.011 | 0.045 | 0.045 | 0.367 | 0.467 | 0.707 |
| Mastery | 0.009 | 0.251 | 0.250 | 0.030 | 0.018 | 0.057 | -0.010 | 0.253 | 0.252 | 0.026 | 0.040 | 0.137 |
| Total Cost of Govt. Programs (Child Ages 1-12 Years) |  |  |  |  |  |  |  |  |  |  |  |  |
| AFDC/TANF | 2585.286 | -177.226 | -0.070 | 0.280 | 0.627 | 0.627 | 2657.084 | -426.434 | -0.165 | 0.073 | 0.087 | 0.156 |
| Food Stamp | 2900.613 | -374.602 | -0.229 | 0.026 | 0.241 | 0.362 | 3191.672 | -288.782 | -0.187 | 0.061 | 0.118 | 0.155 |
| Medicaid | 3462.064 | -367.166 | -0.221 | 0.035 | 0.275 | 0.377 | 3747.045 | -271.420 | -0.183 | 0.068 | 0.153 | 0.153 |

 gives the conditional difference in means between the treatment group and the control group. As mentioned in Section 2 , the control group stands for the original treatment group 2 of the NFP experiment and the treatment group stands for the original group 4. The third column (Effect Size) presents the conditional effect size for the respective group. The fourth column (Asymp. p-value) provides the asymptotic $p$-value for the one-sided single hypothesis test associated with the $t$-statistic for the conditional , the single-hypothesis testing based on the $t$-statistic associated with the treatment indicator in the Freedman and Lane (1983) regression as described in Section 3. By restricted permutation we mean that permutations are done within strata defined by the baseline variables used in the randomization protocol: maternal age and race, gestational age at enrollment, employment status of the head of the household, and geographic region. The covariates used in the Freedman and Lane (1983) regression are: ( jointly are separated by lines. The selection of blocks of outcomes is done on the basis of their meaning. Outcomes that share similar meaning are grouped together. Female

 that is described at the beginning of this section.
Table F.9: Cognitive Abilities and Achievement Outcomes (Correcting for Attrition)

 the unconditional mean for the control group. When factor scores were computed, we set the mean in the control group to zero. The second column (Difference in Means) gives the conditional difference in means between the treatment group and the control group. As mentioned in Section 2 , the control group stands for the original treatment
group 2 of the NFP experiment and the treatment group stands for the original group 4. The third column (Effect Size) presents the conditional effect size for the respective group. The fourth column (Asymp. p-value) provides the asymptotic $p$-value for the one-sided single hypothesis test associated with the $t$-statistic for the conditional difference in means between treatment and control groups. The fifth column (Block Permutation - Single $p$-value) presents the one-sided restricted permutation $p$-values for the single-hypothesis testing based on the $t$-statistic associated with the treatment indicator in the Freedman and Lane (1983) regression as described in Section 3 . By restricted permutation we mean that permutations are done within strata defined by the baseline variables used in the randomization protocol: maternal age and race, gestational age at enrollment, employment status of the head of the household, and geographic region. The covariates used in the Freedman and Lane (1983) regression are: n procedure of Romano and Wolf (2005). Blocks of outcomes that are tested jointly are separated by lines. The selection of blocks of outcomes is done on the basis of their meaning. Outcomes that share similar meaning are grouped together. Female

 that is described at the beginning of this section.
Table F.10: Socio-Emotional Abilities (Correcting for Attrition)

| Outcome Description | Females |  |  |  |  |  | Males |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Basic Statistics |  |  |  | Block Perm. FL |  | Basic Statistics |  |  |  | Block Perm. FL |  |
|  | Cntr. Mean | Cd. Diff. Mn. | Cd. Eff. Size | Asy P-val | Single P-val | Stepdown | Cntr. Mean | Cd. Diff. Mn. | Cd. Eff. Size | Asy P-val | Single P-val | Stepdown |
| Cbild Behavior Checklist (Year 2) - Factor Scores |  |  |  |  |  |  |  |  |  |  |  |  |
| Affective Problems | -0.001 | -0.337 | -0.337 | 0.002 | 0.003 | 0.015 | 0.004 | 0.287 | 0.287 | 0.985 | 0.955 | 0.955 |
| Anxiety Problems | -0.002 | -0.181 | -0.181 | 0.066 | 0.249 | 0.249 | 0.007 | 0.016 | 0.016 | 0.550 | 0.636 | 0.907 |
| Pervasion Developmental Problems | -0.005 | -0.261 | -0.261 | 0.013 | 0.060 | 0.100 | 0.005 | 0.185 | 0.185 | 0.925 | 0.817 | 0.950 |
| Attention Deficit Hyperactivity Disorder | -0.001 | -0.243 | -0.242 | 0.025 | 0.019 | 0.060 | 0.003 | 0.056 | 0.056 | 0.670 | 0.706 | 0.923 |
| Oppositional Defiant Problems | -0.001 | -0.217 | -0.217 | 0.040 | 0.053 | 0.120 | 0.005 | 0.126 | 0.126 | 0.853 | 0.880 | 0.962 |
| Cbild Behavior Checklist (Year 6) - Factor Scores |  |  |  |  |  |  |  |  |  |  |  |  |
| Affective Problems | -0.010 | -0.007 | -0.007 | 0.479 | 0.612 | 0.796 | -0.004 | -0.103 | -0.103 | 0.203 | 0.151 | 0.481 |
| Anxiety Problems | -0.008 | -0.061 | -0.061 | 0.306 | 0.492 | 0.759 | 0.009 | 0.083 | 0.082 | 0.729 | 0.813 | 0.813 |
| Somatic Problems | -0.003 | 0.130 | 0.130 | 0.832 | 0.884 | 0.884 | 0.007 | 0.063 | 0.063 | 0.678 | 0.442 | 0.757 |
| Attention Deficit Hyperactivity Problems | -0.012 | -0.230 | -0.230 | 0.035 | 0.096 | 0.307 | -0.006 | -0.040 | -0.040 | 0.379 | 0.310 | 0.713 |
| Oppositional Defiant Problems | 0.000 | -0.027 | -0.027 | 0.415 | 0.286 | 0.608 | -0.013 | -0.083 | -0.083 | 0.270 | 0.317 | 0.672 |
| Conduct Problems | -0.002 | -0.267 | -0.266 | 0.013 | 0.003 | 0.015 | -0.009 | -0.011 | -0.011 | 0.467 | 0.485 | 0.665 |
| MacArthur Story Stem Battery (MSSB) (Year 6) - Factor Scores |  |  |  |  |  |  |  |  |  |  |  |  |
| Dysregulated Aggression | -0.006 | -0.027 | -0.027 | 0.413 | 0.135 | 0.269 | -0.009 | -0.130 | -0.130 | 0.189 | 0.137 | 0.496 |
| Warmth and Empathy | -0.011 | 0.388 | 0.388 | 0.002 | 0.005 | 0.019 | -0.011 | -0.099 | -0.099 | 0.770 | 0.535 | 0.832 |
| Emotional Integration | -0.005 | -0.028 | -0.028 | 0.585 | 0.765 | 0.765 | -0.015 | 0.055 | 0.055 | 0.349 | 0.429 | 0.849 |
| Performance Anxiety | 0.010 | -0.038 | -0.038 | 0.373 | 0.093 | 0.259 | -0.009 | 0.077 | 0.077 | 0.701 | 0.843 | 0.843 |
| Aggression | -0.005 | -0.164 | -0.164 | 0.084 | 0.003 | 0.012 | -0.010 | -0.095 | -0.095 | 0.260 | 0.177 | 0.570 |
| Internalizing, Externalizing, Absences (Year 12) |  |  |  |  |  |  |  |  |  |  |  |  |
| Internalizing Disorders | 0.240 | -0.028 | -0.066 | 0.309 | 0.453 | 0.813 | 0.397 | -0.087 | -0.183 | 0.090 | 0.082 | 0.154 |
| Externalizing Disorders | 0.183 | -0.013 | -0.032 | 0.402 | 0.641 | 0.866 | 0.183 | 0.089 | 0.239 | 0.951 | 0.859 | 0.859 |
| Average \# of Absences (School Years 1-5) | 10.144 | 0.263 | 0.035 | 0.605 | 0.666 | 0.666 | 11.548 | -1.838 | -0.246 | 0.029 | 0.027 | 0.077 |

 the unconditional mean for the control group. When factor scores were computed, we set the mean in the control group to zero. The second column (Difference in Means) group 2 of the NFP experiment and the treatment group stands for the original group 4. The third column (Effect Size) presents the conditional effect size for the respective group. The fourth column (Asymp. p-value) provides the asymptotic $p$-value for the one-sided single hypothesis test associated with the $t$-statistic for the conditional difference in means between treatment and control groups. The fifth column (Block Permutation - Single p-value) presents the one-sided restricted permutation $p$-values for restricted permutation we mean that permutations are done within strata defined by the baseline variables used in the randomization protocol: maternal age and race, gestational age at enrollment, employment status of the head of the household, and geographic region. The covariates used in the Freedman and Lane (1983) regression are: maternal height, household income, grandmother support, maternal parenting attitudes and mother currently in school. Finally, the last column (Block Permutation Stepdown) provides $p$-values that account for multiple-hypothesis testing based on the Stepdown procedure of Romano and Wolf ( 2005 ). Blocks of outcomes that are tested jointly are separated by lines. The selection of blocks of outcomes is done on the basis of their meaning. Outcomes that share similar meaning are grouped together. Female
 that is described at the beginning of this section.

## G A Framework for Mediation Analysis

This section develops a theoretical framework to conduct our mediation analysis. Our model is motivated by the literature on the technology of skill formation (Cunha and Heckman, 2007). In it, subsequent skills build on earlier skills to generate human capital. Notationally, let $\boldsymbol{\theta}_{i, t}$ be the vector of skills during childhood for individual $i$ at period $t$ and $t \in\{0,1, \ldots, T\}$, where T is the number of periods of childhood. Let $\boldsymbol{I}_{i, t}$ represent investments at the same period. We use $\boldsymbol{X}_{i}$ for family background characteristics and $v_{i, t}$ for an exogenous error term independent of $\boldsymbol{\theta}_{i, t}, \boldsymbol{I}_{i, t}$ and $\boldsymbol{X}_{i}$. The structural equations that govern the evolution of skills are given by:

$$
\begin{equation*}
\boldsymbol{\theta}_{i, t+1}=\boldsymbol{q}_{t+1}\left(\boldsymbol{\theta}_{i, t}, \boldsymbol{I}_{i, t+1}, \boldsymbol{X}_{i}, v_{i, t+1}\right) ; t \in\{0,1, \ldots, T-1\} . \tag{17}
\end{equation*}
$$

By the term "structural equations," we mean autonomous functions in the language of Frisch (1938), i.e. deterministic functions whose functional forms do not change as their arguments vary. We also allow for skills to affect investments, that is:

$$
\begin{equation*}
\boldsymbol{I}_{i, t+\boldsymbol{1}}=\boldsymbol{h}_{t+1}\left(\boldsymbol{\theta}_{i, t}, \boldsymbol{X}_{i}, \varepsilon_{i, t+1}\right) ; t \in\{0,1, \ldots, T-1\} \tag{18}
\end{equation*}
$$

where $\varepsilon_{i, t+1}$ is an exogenous error term independent of $\boldsymbol{\theta}_{i, t}$ and $\boldsymbol{X}_{i}$. Our model is completed by the following structural outcome equation at period $T$ :

$$
\begin{equation*}
\boldsymbol{Y}_{\boldsymbol{i}}=\boldsymbol{g}_{T}\left(\boldsymbol{\theta}_{i, T}, \boldsymbol{X}_{i}, \xi_{i, T}\right) \tag{19}
\end{equation*}
$$

where $\xi_{i, T}$ is an exogenous error term independent of $\boldsymbol{\theta}_{i, T}$ and $\boldsymbol{X}_{i}$.
We can use a recursive substitution of investments and skills of Equations (17)-(18) into (19) to generate the following equation:

$$
\begin{equation*}
\boldsymbol{Y}_{\boldsymbol{i}}=\boldsymbol{f}_{t^{\prime}}\left(\boldsymbol{\theta}_{i, t^{\prime}}, \boldsymbol{X}_{i},\left\{v_{i, \tilde{t}}\right\}_{\tilde{t}=t^{\prime}}^{T},\left\{\varepsilon_{i, \tilde{t}}\right\}_{\tilde{t}=t^{\prime}}^{T}, \xi_{i, T}\right), \tag{20}
\end{equation*}
$$

where $\left\{v_{i, \tilde{t}}\right\}_{\tilde{t}=t^{\prime}}^{T}=\left\{v_{i, t^{\prime}}, v_{i, t^{\prime}+1}, \ldots, v_{i, T}\right\}$ and $\left\{\varepsilon_{i, \tilde{t}}\right\}_{\tilde{t}=t^{\prime}}^{T}=\left\{\varepsilon_{i, t^{\prime}}, \varepsilon_{i, t^{\prime}+1}, \ldots, \varepsilon_{i, T}\right\}$.
Suppose that an intervention occurs at period $t^{\prime}$ where $t^{\prime} \in\{1, \ldots, T\}$. Let $D_{i} \in\{0,1\}$ be the treatment indicator of this intervention which takes value 1 if participant $i$ is treated and 0 otherwise. The intervention enters our technology of skill formation model as a form of skill investment. Thus we append the investment Equation (18) at period $t^{\prime}$ by:

$$
\begin{equation*}
\boldsymbol{I}_{i, t^{\prime}}=\boldsymbol{h}_{t^{\prime}}\left(\boldsymbol{\theta}_{i, t^{\prime}-1}, D_{i}, \boldsymbol{X}_{i}, \varepsilon_{i, t^{\prime}}\right) ; \text { for some } t^{\prime} \in\{0,1, \ldots, T-1\} \tag{21}
\end{equation*}
$$

The counterfactual values investment $\boldsymbol{I}_{i, t^{\prime}}$ are defined by the value $\boldsymbol{I}_{i, t^{\prime}}$ takes when the intervention $D_{i}$ is fixed at a level $d \in\{0,1\}$. By fixing, I mean the causal operation defined in Haavelmo (1944) where $D_{i}$ is set to $d \in\{0,1\}$ as argument in the structural equation (21). That is:

$$
\begin{equation*}
\boldsymbol{I}_{i, t^{\prime}, \boldsymbol{d}}=\boldsymbol{h}_{t^{\prime}}\left(\boldsymbol{\theta}_{i, t^{\prime}-1}, d, \boldsymbol{X}_{i}, \varepsilon_{i, t^{\prime}} ; d \in\{0,1\} \text { for some } t^{\prime} \in\{0,1, \ldots, T-1\} .\right. \tag{22}
\end{equation*}
$$

Let the counterfactual skills be defined in a symmetric fashion by:

$$
\boldsymbol{\theta}_{i, t^{\prime}, d}=\boldsymbol{q}_{t^{\prime}}\left(\boldsymbol{\theta}_{i, t^{\prime}-1}, \boldsymbol{I}_{i, t^{\prime}, d}, \boldsymbol{X}_{i}, v_{i, t^{\prime}}\right)
$$

We also define the counterfactual skills and investments for periods $t>t^{\prime}$ by:

$$
\begin{aligned}
\boldsymbol{I}_{i, t+1, d} & =\boldsymbol{h}_{t+1}\left(\boldsymbol{\theta}_{i, t, d}, \boldsymbol{X}_{i}, \varepsilon_{i, t+1}\right), \text { and } \\
\boldsymbol{\theta}_{i, t+1, d} & =\boldsymbol{q}_{t+1}\left(\boldsymbol{\theta}_{i, t, d}, \boldsymbol{I}_{i, t+1, d}, \boldsymbol{X}_{i}, v_{i, t+1}\right) ; t>t^{\prime}
\end{aligned}
$$

We can also define the counterfactual outcomes by:

$$
\begin{equation*}
\boldsymbol{Y}_{i, d}=\boldsymbol{f}_{t^{\prime}}\left(\boldsymbol{\theta}_{i, t^{\prime}, d}, \boldsymbol{X}_{i},\left\{\epsilon_{i, \tilde{t}}\right\}_{\tilde{t}=t^{\prime}}^{T},\left\{\varepsilon_{i, \tilde{t}}\right\}_{\tilde{t}=t^{\prime}}^{T}, \xi_{i, T}\right) \tag{23}
\end{equation*}
$$

If the intervention assignment uses the method of randomization, then we have that:

$$
\left(\boldsymbol{Y}_{i, \boldsymbol{d}}, \boldsymbol{\theta}_{i, t^{\prime}, d}\right) \Perp D_{i} \mid \boldsymbol{X}_{i} ; d \in\{0,1\} .
$$

We can also write the realized values of skills and outcomes as:

$$
\begin{aligned}
\boldsymbol{Y}_{\boldsymbol{i}} & =\boldsymbol{Y}_{i, 1} D_{i}+\boldsymbol{Y}_{i, 0}\left(1-D_{i}\right), \text { and } \\
\boldsymbol{\theta}_{i, t} & =\boldsymbol{\theta}_{i, t, 1} D_{i}+\boldsymbol{\theta}_{i, t, 0}\left(1-D_{i}\right) ; t \geq t^{\prime}
\end{aligned}
$$

We use Equation (23) to generate a tractable equation to examine mediation effects. Note that Equation (23) holds not only for $t^{\prime}$ but for any $t \geq t^{\prime}$.

$$
\begin{equation*}
\boldsymbol{Y}_{i, d}=\boldsymbol{f}_{t}\left(\boldsymbol{\theta}_{i, t, d}, \boldsymbol{X}_{i},\left\{v_{i, \tilde{t}}\right\}_{\tilde{t}=t}^{T},\left\{\varepsilon_{i, \tilde{t}}\right\}_{\tilde{t}=t}^{T}, \xi_{i, T}\right), \text { for any } t \in\left\{t^{\prime}, t^{\prime}+1, \ldots, T\right\} . \tag{24}
\end{equation*}
$$

Error terms $\left(\left\{v_{i, \tilde{t}}\right\}_{\tilde{t}=t}^{T},\left\{\varepsilon_{i, t}\right\}_{\tilde{t}=t}^{T}, \xi_{i, T}\right)$ are independent of $\boldsymbol{\theta}_{i, t, d}$ and $\boldsymbol{X}_{i}$. For sake of notational simplicity, we can substitute those error terms by $\zeta_{t}$ without loss of generality. Equation (24) then becomes:

$$
\begin{equation*}
\boldsymbol{Y}_{i, d}=f_{t}\left(\boldsymbol{\theta}_{i, t, d}, \boldsymbol{X}_{i}, \zeta_{i, t}\right) \tag{25}
\end{equation*}
$$

We achieve a linear form of Equation (25) by approximating it through a Maclaurin expansion. This generates the following equation:

$$
\begin{equation*}
\boldsymbol{Y}_{i, d, t}=\boldsymbol{\kappa}_{\boldsymbol{t}}+\boldsymbol{\alpha}_{t, d} \boldsymbol{\theta}_{i, t, d}+\boldsymbol{\beta}_{t, d} \boldsymbol{X}_{i}+\boldsymbol{\epsilon}_{i, t, \boldsymbol{d}}, \quad d \in\{0,1\} . \tag{26}
\end{equation*}
$$

where $\boldsymbol{\epsilon}_{\boldsymbol{t}, \boldsymbol{d}}$ accounts for the approximation error. Equations (25)-(26) are used in our mediation analysis in Section 5.

## H Mediation Methodology

## H. 1 Three Step Procedure

This part of the appendix explains in detail the three step procedure that we use in order to decompose the NFP treatment effects. As noted in the paper, we perform two sets of analyses. First, we study whether the treatment effects on child skills at age 6 were mediated by program enhancement of birth weight, parenting attitudes and investments, and maternal socio-emotional skills at age 2 . Second, we study whether the program impact on outcomes at age 12 was mediated by the NFP enhancement of skills at age 6 . The results from these analysis shed light on the complementarity of investments and skills in explaining the NFP treatment effects.

Step One The idea is to develop a measurement system that links the observed items and the latent skills. In order to do that, we assume that our measurements are dedicated. This means that each observed measurement is linked to a unique skill. Specifically, let $\mathcal{M}^{j}$ be the index set of measures associated with trait $j$, where $j \in \mathcal{J}=\{P, C, S E\} . P, C, S E$ denote, respectively, parenting skills, child cognitive skills, and child socio-emotional abilities. ${ }^{3}$ Thus, our linear measurement system is as follows: ${ }^{4}$

$$
\begin{equation*}
M_{m^{j}, d}^{j}=\nu_{m^{j}}^{j}+\varphi_{m^{j}}^{j} \boldsymbol{\theta}_{d}^{j}+\eta_{m^{j}, d}^{j}, \tag{27}
\end{equation*}
$$

where $\nu_{m^{j}}^{j}$ is the intercept term and $\varphi_{m^{j}}^{j}$ represents the loading factor of trait $j$. We cannot reject the null hypothesis that the intercepts and loading factors depend on treatment status. $\eta_{m^{j}, d}$ is a mean zero idiosyncratic error term which, by assumption, is independent of $\boldsymbol{\theta}_{d}^{j}$ $\forall j \in \mathcal{J}$. We normalize the loading factor associated with the first measure of each factor

[^4]to 1 in order to set a scale, otherwise the scale is arbitrary. ${ }^{5}$ Finally, we allow for factor correlation.

The parameters that identify the measurement system are the factor means, the factor covariances, the intercepts, the factor loadings, and the variances of the error terms: $\left.E\left[\boldsymbol{\theta}^{j}(\boldsymbol{d})\right]=\mu_{d}^{j}, \operatorname{Var}\left[\boldsymbol{\theta}_{\boldsymbol{d}}\right]=\Sigma_{\boldsymbol{\theta}_{\boldsymbol{d}}}, \boldsymbol{\nu}_{\boldsymbol{m}^{j}}^{j}, \boldsymbol{\varphi}_{\boldsymbol{m}^{j}}^{j}, \operatorname{Var}\left[\eta_{m^{j}}^{j}\right]\right)$. Heckman et al. (2013) show that the existence of at least three measures for each latent skill guarantees identification. ${ }^{6}$ Broadly, means, variances, and covariances across the measures identify the parameters of the system.

We estimate the parameters of the measurement system that links skills with measures both at ages 2 and 6 . Variables become potential mediators if we estimate an effect of the NFP on it, so that they are potential meaningful channels. For age 2, non-abusive parenting attitudes are approximated by the Adult-Adolescent Parenting Inventory (Bavolek), which comprises 32 items, and home investments are measured by the Bradley and Caldwell Home Observation for measurement of the Environment (HOME) inventory, which is composed of 45 items.

The maternal skills selected correspond to anxiety, assessed by the Rand Mental Health Inventory, self-esteem, measured by the Rosenberg scale, and mastery, approximated by the Pearlin scale. Similarly, for age 6, we select children's skills influenced by the NFP as plausible mediators . Child cognition is measured by 8 subtests from the K-ABC mental processing composite. For children's socio-emotional skills, we identify as potential mediators the treatment reduction in conduct, attention and aggression problems, as well as the enhancement of children's pro-social skills. Attention and conduct problems are approximated by items from the Child Behavior Checklist. Pro-social skills (warmth or empathy) and aggression problems are approximated by items from the MacArthur Story Stem Battery. Section C of the Appendix explains in more detail these tests, as well as the instruments they use.

We estimate the parameters of the measurement system by maximum likelihood. In

[^5]order to do this, we assume that the latent skills and the error terms, $\theta^{j}$ and $\eta_{m^{j}}^{j}$, are normal and i.i.d. We use full-information maximum likelihood to deal with the missing values in the measures for some individuals. FIML yields unbiased estimates that are more efficient than ad hoc methods like list-wise and pair-wise deletion, which work under the implicit assumption of random missing data. By missing at random we mean that the probability of data associated with a variable $x$ can depend on other observed variables but not on the values of $x$ itself.

For the case of the measurement system at age 2, we have 146 items. Although it is ideal to estimate the complete set of items (skills) jointly, it is not feasible. Thus, we estimate them in two blocks: one for parenting and home investments and other for maternal characteristics. This allows us to account for the correlation between the skills that are in the same block. For the case of the measurement system at age 6 , the set of items is smaller and we do a joint estimation.

Step Two In the second step we use the parameter estimates from the first step to construct factor scores for each children. The objective of this is to construct approximations for the latent skills. The two most common linear scoring methods are the regression method and the Bartlett method, which resembles GLS (Thomson, 1934). We use the Bartlett (1937) method because it estimates unbiased approximations of the unobserved skills. Actually, this guarantees that the difference in means between the factor scores for children in the treatment and the control groups equals the difference in means in the true scores. The derivation of the Bartlett estimator begins with the measurement system summarized as:

$$
\underbrace{\boldsymbol{M}_{i}}_{|\mathcal{M}| \times 1}=\underbrace{\varphi}_{|\mathcal{M}| \times|\mathcal{J}| \mathcal{J} \mid \times 1} \underbrace{\boldsymbol{\theta}_{i}}_{|\mathcal{M}| \times 1}+\underbrace{}_{\boldsymbol{\eta}_{i}}
$$

where the dimension of each term is below the braces (recall that $\mathcal{J}$ and $\mathcal{M}$ are the indexing sets for skills and measures respectively). Assume that the $\left(\boldsymbol{\theta}_{i}, \boldsymbol{\eta}_{i}\right), i \in\{1, \ldots, I\}$, are
independent across $i$. For simplicity, we assume that they are i.i.d. ${ }^{7}$ Let $\operatorname{Cov}\left(\boldsymbol{M}_{i}, \boldsymbol{M}_{i}\right)=\boldsymbol{\Sigma}$, $\operatorname{Cov}\left(\boldsymbol{\theta}_{i}, \boldsymbol{\theta}_{i}\right)=\boldsymbol{\Phi}$ and $\operatorname{Cov}\left(\boldsymbol{\eta}_{i}, \boldsymbol{\eta}_{i}\right)=\boldsymbol{\Omega}$. The linear relation between the factor scores and the measures is the following:

$$
\begin{equation*}
\boldsymbol{\theta}_{S, i}=\boldsymbol{L}^{\prime} \boldsymbol{M}_{i} \tag{28}
\end{equation*}
$$

In order to obtain unbiased estimates, Barlett imposes the restriction that $\boldsymbol{L}^{\prime} \boldsymbol{\varphi}=\boldsymbol{I}_{|\mathcal{J}|}$. The Bartlett estimator for the vector of approximated skills $\left(\boldsymbol{\theta}_{\boldsymbol{i}}\right)$ is:

$$
\begin{equation*}
\boldsymbol{\theta}_{S, i}=\left(\hat{\varphi}^{\prime} \hat{\Omega}^{-\mathbf{1}} \hat{\boldsymbol{\varphi}}\right)^{-1} \hat{\boldsymbol{\varphi}}^{\prime} \hat{\Omega}^{-\mathbf{1}} \boldsymbol{M}_{\boldsymbol{i}} \tag{29}
\end{equation*}
$$

where the matrix of loading factors, $\hat{\boldsymbol{\varphi}}$, and $\hat{\boldsymbol{\Omega}}=\operatorname{Cov}\left(\boldsymbol{\eta}_{i}, \boldsymbol{\eta}_{i}\right)$ are both estimated in the first step. Bartlett's estimator is a Generalized Least Squares, $G L S$, procedure where measures are used as dependent variables and loading factors are treated as regressors. By the GaussMarkov theorem, the Bartlett GLS estimator is optimal and hence leads to the best linear unbiased predictor (BLUE).

There are individuals that have missing data in some of the items that compose the measurement system. In order to take advantage of the information that they have (instead of list-wise delete them), we predict factor scores for them. We use the covariance between the measures and the factors from the sample with complete measurement system to predict scores for these people. Additionally, for the cases were individuals are missing a factor score because they did not have any item in that measurement system, we impute factor scores with the regression method. ${ }^{8}$ This procedure recovers around $10 \%$ of the randomized sample.

Step 3 In this step, we use factor scores as approximations of the true skills to estimate the models that link the later outcomes with the intermediate skills. The factor scores are

[^6]measured with error, which produces downward-biased estimates of the parameters of the outcome equations. This bias corresponds to the traditional attenuation that results from classical measurement error. In factor scored regressions, Bolck et al. (2008) prove this. We adopt the bias correction strategy proposed by Croon (2002). In summary, this approach takes advantage of the fact that we have estimates of all the components of the bias. This strategy, also used by Heckman et al. (2013), can be summarized as follows:

Consider the model following model. To simplify notation, we use $W$ to denote preprogram variables $\boldsymbol{X}$, treatment indicator and the intercept of equation 4:

$$
\begin{equation*}
\boldsymbol{Y}_{\boldsymbol{i}}=\boldsymbol{\alpha} \boldsymbol{\theta}_{i}+\gamma \boldsymbol{W}_{i}+\boldsymbol{\epsilon}_{i}, \quad i=1, \ldots, N . \tag{30}
\end{equation*}
$$

The covariance matrix of $\left(\boldsymbol{\theta}_{i}, \boldsymbol{W}_{i}\right)$ is

$$
\left(\begin{array}{cc}
\operatorname{Cov}(\boldsymbol{\theta}, \boldsymbol{\theta}) & \operatorname{Cov}(\boldsymbol{\theta}, \boldsymbol{W}) \\
\operatorname{Cov}(\boldsymbol{W}, \boldsymbol{\theta}) & \operatorname{Cov}(\boldsymbol{W}, \boldsymbol{W})
\end{array}\right)
$$

We measure $\boldsymbol{\theta}_{i}$ with error. Thus,

$$
\begin{gathered}
\boldsymbol{\theta}_{S, i}=\boldsymbol{\theta}_{i}+\boldsymbol{V}_{i}, \quad i=1, \ldots, N \\
\left(\boldsymbol{W}_{i}, \boldsymbol{\theta}_{i}\right) \Perp \boldsymbol{V}_{i}, \quad E\left(\boldsymbol{V}_{i}\right)=0, \quad \operatorname{Cov}(\boldsymbol{V}, \boldsymbol{V})=\boldsymbol{\Sigma}_{\boldsymbol{V} \boldsymbol{V}}
\end{gathered}
$$

Denote $\operatorname{Cov}\left(\boldsymbol{\theta}_{S, i}, \boldsymbol{\theta}_{S, i}\right)=\boldsymbol{\Sigma}_{\boldsymbol{\theta}_{S}, \boldsymbol{\theta}_{S}}$. We assume that the $\left(\boldsymbol{\theta}_{i}, \boldsymbol{W}_{i}, \boldsymbol{\epsilon}_{i}\right)$ are i.i,d, but much weaker conditions suffice. Note that we do not assume that $\boldsymbol{\theta}_{i} \Perp \boldsymbol{W}_{i}$ as in traditional factor analysis. We do assume that $\left(\boldsymbol{\theta}_{i}, \boldsymbol{W}_{i}\right) \Perp \boldsymbol{\epsilon}_{i}$ and $E\left(\boldsymbol{\epsilon}_{i}\right)=0$.

If we use $\boldsymbol{\theta}_{S, i}$ in place of $Y_{i}$, it follows that:

$$
\begin{equation*}
\boldsymbol{Y}_{\boldsymbol{i}}=\boldsymbol{\alpha} \boldsymbol{\theta}_{S, i}+\gamma \boldsymbol{W}_{i}+\boldsymbol{\epsilon}_{i}-\alpha \boldsymbol{V}_{i} . \tag{31}
\end{equation*}
$$

The estimation of Equation 31 using OLS produces estimates that are biased:
$\operatorname{plim}\binom{\hat{\boldsymbol{\alpha}}}{\hat{\gamma}}=\left(\begin{array}{cc}\operatorname{Cov}\left(\boldsymbol{\theta}_{S}, \boldsymbol{\theta}_{S}\right) & \operatorname{Cov}\left(\boldsymbol{\theta}_{S}, \boldsymbol{W}\right) \\ \operatorname{Cov}\left(\boldsymbol{W}, \boldsymbol{\theta}_{S}\right) & \operatorname{Cov}(\boldsymbol{W}, \boldsymbol{W})\end{array}\right)^{-1}\left(\begin{array}{cc}\operatorname{Cov}(\boldsymbol{\theta}, \boldsymbol{\theta}) & \operatorname{Cov}(\boldsymbol{\theta}, \boldsymbol{W}) \\ \operatorname{Cov}(\boldsymbol{W}, \boldsymbol{\theta}) & \operatorname{Cov}(\boldsymbol{W}, \boldsymbol{W})\end{array}\right)\binom{\boldsymbol{\alpha}}{\gamma}$.
Let $\boldsymbol{\Sigma}_{\mathbf{B}, \mathbf{C}}$ be $\operatorname{Cov}(\mathbf{B}, \mathbf{C})$. Observe that $\boldsymbol{\Sigma}_{\boldsymbol{\theta}, \boldsymbol{W}}=\boldsymbol{\Sigma}_{\boldsymbol{\theta}_{S}, \boldsymbol{W}}$ as a consequence of our assumptions. In this notation

$$
\operatorname{plim}\binom{\hat{\alpha}}{\hat{\gamma}}=\underbrace{\left(\begin{array}{cc}
\Sigma_{\theta, \theta}+\Sigma_{V, V} & \Sigma_{\theta, W}  \tag{32}\\
\Sigma_{W, \theta} & \Sigma_{W, W}
\end{array}\right)^{-1}\left(\begin{array}{cc}
\Sigma_{\theta, \theta} & \Sigma_{\theta, W} \\
\Sigma_{W, \theta} & \Sigma_{W, W}
\end{array}\right)}_{\boldsymbol{A}}\binom{\alpha}{\gamma}
$$

which is the usual attenuation formula.
From the estimation of the measurement system, we can identify $\boldsymbol{\Sigma}_{\boldsymbol{\theta}, \boldsymbol{\theta}}, \boldsymbol{\Sigma}_{\boldsymbol{\theta}, \boldsymbol{W}}, \boldsymbol{\Sigma}_{\boldsymbol{V}, \boldsymbol{V}}$, and we have all the components of $\boldsymbol{A}$. Hence if we pre-multiply the least squares estimator by $\boldsymbol{A}^{-1}$, we obtain:

$$
\operatorname{plim} \boldsymbol{A}^{-1}\binom{\hat{\boldsymbol{\alpha}}}{\hat{\gamma}}=\binom{\boldsymbol{\alpha}}{\gamma} .
$$

This is called "Croon's method" in psychometrics (Croon, 2002). In our application, there are two groups corresponding to $D=0$ and $D=1$ (control and treatment, respectively). We allow $\boldsymbol{\theta}_{i}$ to vary by treatment status. Indeed, our method assumes that treatment only operates through shifting the distribution of $\boldsymbol{\theta}$. We do not normalize the means of $\boldsymbol{\theta}$ (or $\boldsymbol{W}$ ) to be zero.

In the third step of our estimation procedure we compute bootstrapped $p$-values for each decomposition channel of the treatment effects. We take 100,000 resamples with replacement. The bootstrapped p-value for the null hypothesis $H_{0}: \alpha_{j}=0$ is calculated as follows:

$$
\begin{equation*}
p \text {-value }=\frac{1}{B} \sum_{b=1}^{B} 1\left(t_{b}^{j, *}>t^{j}\right) \text { with } t^{j}=\frac{\hat{\alpha}^{j}}{\hat{\sigma}\left(\hat{\alpha}^{j}\right)} \text { and } t_{b}^{j, *}=\frac{\left(\hat{\alpha}_{b}^{j}-\hat{\alpha}^{j}\right)}{\hat{\sigma}\left(\hat{\alpha}_{b}^{j}\right)} \tag{33}
\end{equation*}
$$

where $\hat{\alpha_{b}^{j}}$ is bootstrapped estimated in the $b^{t h}$ resample and $\hat{\alpha_{j}}$ is estimated from the original data. Given the estimates of the outcome equation and of the factor scores, we construct the bootstrapped p-value for the contribution of skill $k$ under the null hypothesis $H_{0}$ : $\hat{\alpha}^{j} E\left(\theta_{1}^{j}-\theta_{0}^{j}\right)=0$ as follows:

$$
\begin{equation*}
p \text {-value }=\frac{1}{B} \sum_{b=1}^{B} 1\left(T_{b}^{j, *}>T^{j}\right) \text { with } T^{j}=\frac{\hat{\alpha}^{j} * E\left(\theta^{j}\left(\widehat{1)-\theta^{j}}(0)\right)\right.}{\hat{\sigma}\left(\hat{\alpha}^{j} * E\left(\theta^{j}\left(\widehat{1)-\theta^{j}}(0)\right)\right)\right.} \tag{34}
\end{equation*}
$$

where $T_{b}^{j, *}$ is the statistic $T^{j}$ computed with the parameters obtained in the $b^{t h}$ resample. Notice that the p-value combines the variation in two population parameters: 1) the coefficient of the outcome equation; 2) the experimentally induced difference in means in the skills. It could be the case that each of these parameters are, separately, statistically significant. However, the p-value may increase due to a loss in power when they are combined.

Tables H.1-H. 4 shows the parameters of the outcome equations as wells as the decompositions components.
Table H.1: Female Decomposition (Year 6)

|  | Treatment |  | Birth Weight |  | Home y2 |  | Parenting y 2 |  | Anx |  | Self-E | m y 2 | Mastery y 2 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coefficient $p$-value |  | Coefficient $p$-value |  | Coefficient $p$-value |  | Coefficient $p$-value |  | Coefficient $p$-value |  | Coefficient $p$-value |  | Coefficient $p$-value |  | Sample Size |
| Outcome Coefficients |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Cognitive | 0.04 | 0.391 | 0.08 | 0.089 | 0.23 | 0.035 | 0.06 | 0.139 | 0.21 | 0.087 | 0.16 | 0.288 | -0.08 | 0.363 | 304 |
| Attention Problems | -0.15 | 0.083 | -0.11 | 0.013 | -0.11 | 0.169 | -0.07 | 0.042 | -0.14 | 0.139 | 0.24 | 0.206 | -0.19 | 0.181 | 304 |
| Conduct Problems | -0.15 | 0.036 | -0.09 | 0.018 | -0.07 | 0.249 | -0.03 | 0.192 | -0.18 | 0.032 | -0.20 | 0.190 | 0.11 | 0.255 | 304 |
| Warmth/Empathy | 0.18 | 0.060 | 0.05 | 0.192 | 0.29 | 0.003 | 0.09 | 0.014 | -0.01 | 0.481 | -0.41 | 0.101 | 0.26 | 0.125 | 304 |
| Aggression | -0.13 | 0.103 | -0.04 | 0.218 | -0.15 | 0.107 | -0.01 | 0.416 | 0.13 | 0.177 | -0.13 | 0.299 | -0.06 | 0.386 | 304 |
| Treatment Effect |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Cognitive | 0.04 | 0.391 | -0.01 | 0.110 | 0.04 | 0.032 | 0.02 | 0.079 | 0.03 | 0.081 | 0.03 | 0.246 | -0.02 | 0.321 | 304 |
| Attention Problems | -0.15 | 0.083 | 0.01 | 0.099 | -0.02 | 0.144 | -0.02 | 0.046 | -0.02 | 0.115 | 0.04 | 0.168 | -0.05 | 0.134 | 304 |
| Conduct Problems | -0.15 | 0.036 | 0.01 | 0.112 | -0.01 | 0.223 | -0.01 | 0.170 | -0.03 | 0.065 | -0.04 | 0.153 | 0.03 | 0.208 | 304 |
| Warmth/Empathy | 0.18 | 0.060 | -0.01 | 0.169 | 0.05 | 0.007 | 0.03 | 0.018 | -0.00 | 0.434 | -0.07 | 0.073 | 0.08 | 0.093 | 304 |
| Aggression | -0.13 | 0.103 | 0.00 | 0.173 | -0.03 | 0.090 | -0.00 | 0.401 | 0.02 | 0.127 | -0.02 | 0.263 | -0.02 | 0.351 | 304 |
| Treatment Effect Fraction |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Cognitive | 0.29 | 0.391 | -0.09 | 0.110 | 0.35 | 0.032 | 0.14 | 0.079 | 0.25 | 0.081 | 0.24 | 0.246 | -0.19 | 0.321 | 304 |
| Attention Problems | 0.73 | 0.083 | -0.07 | 0.099 | 0.10 | 0.144 | 0.09 | 0.046 | 0.10 | 0.115 | -0.20 | 0.168 | 0.26 | 0.134 | 304 |
| Conduct Problems | 0.79 | 0.036 | -0.06 | 0.112 | 0.06 | 0.223 | 0.05 | 0.170 | 0.14 | 0.065 | 0.19 | 0.153 | -0.16 | 0.208 | 304 |
| Warmth/Empathy | 0.71 | 0.060 | -0.03 | 0.169 | 0.21 | 0.007 | 0.11 | 0.018 | -0.01 | 0.434 | -0.29 | 0.073 | 0.29 | 0.093 | 304 |
| Aggression | 0.74 | 0.103 | -0.03 | 0.173 | 0.16 | 0.090 | 0.02 | 0.401 | -0.12 | 0.127 | 0.13 | 0.263 | 0.10 | 0.351 | 304 |

Notes: The first column provides the outcome description and the top row provides information on the mediators. For Year 6, the mediators are treatment, birth weight, divided into 3 groups: Outcome Coefficients, Treatment Effect and Treatment Effect Fraction. The last of these groups is also shown visually in Figure 3. Each mediator has two subcolumns of information: the coefficient and the $p$-value. Bold $p$-values are significant at the $10 \%$ level. We used the following controls: maternal race, maternal age, maternal height, gestational age, household density, region, employment status of household head, grandmother support, randomization wave, income category, mother currently in school, and maternal parenting attitudes.
Table H.2: Male Decomposition (Year 6)
 home environment, parenting, anxiety, self-esteem and mastery. The last column provides the sample size for the corresponding outcome in the first column. The rows are divided into 3 groups: Outcome Coefficients, Treatment Effect and Treatment Effect Fraction. The last of these groups is also shown visually in Figure 4 . Each mediator has two subcolumns of information: the coefficient and the $p$-value. Bold $p$-values are significant at the $10 \%$ level. We used the following controls: maternal race, maternal age,
maternal height, gestational age, household density, region, employment status of household head, grandmother support, randomization wave, income category, mother currently in school, and maternal parenting attitudes.
Table H.3: Female Decomposition (Year 12)

|  | Treatment |  | Cognition |  | Attention problems |  | Conduct Problems |  | Warmth/Empathy |  | Aggression |  | Sample Size |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coefficient | $p$-value | Coefficient | $p$-value | Coefficient | $p$-value | Coefficient | $p$-value | Coefficient | $p$-value | Coefficient | $p$-value |  |
| Outtome Copfitients |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Total \# Days Child Used Marijuana | ${ }^{-0.15}$ | 0.060 | -0.12 | ${ }^{0.086}$ | ${ }^{0.03}$ | ${ }^{0.373}$ | ${ }^{0.00}$ | ${ }^{0.515}$ | ${ }^{0.12}$ | ${ }^{0.158}$ | ${ }^{0.04}$ | ${ }^{0.192}$ | ${ }^{271}$ |
| Child Used Alcohol, Marijuana, or Tobacco in Last 30 Days | -0.02 | 0.198 | -0.02 | 0.048 | 0.00 | 0.467 | 0.02 | 0.338 | -0.00 | 0.477 | 0.00 | 0.457 | 268 |
| Standardized Child BMI (Year 12) | -0.30 | 0.016 | -0.11 | 0.100 | -0.22 | 0.115 | 0.37 | 0.030 | 0.08 | 0.206 | -0.11 | 0.197 | 272 |
| $\underline{T r a t m e n t ~ E f f e r t ~}$ |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Child Ever Used Marijuana | -0.15 | 0.060 | -0.01 | 0.218 | -0.00 | 0.318 | -0.00 | 0.484 | 0.04 | 0.111 | -0.01 | 0.155 | 271 |
| Child Used Alcohol, Marijuana, or Tobacco in Last 30 Days | -0.02 | 0.198 | -0.00 | 0.217 | -0.00 | 0.431 | -0.00 | 0.273 | -0.00 | 0.469 | -0.00 | 0.424 | 268 |
| Standardized Child BMI (Year 12) | -0.30 | 0.016 | -0.01 | 0.209 | 0.03 | 0.109 | -0.05 | 0.060 | 0.02 | 0.162 | 0.02 | 0.145 | 272 |
| Tratment Effect Fraction |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Child Ever Used Marijuana | 1.12 | 0.060 | 0.05 | 0.218 | 0.03 | 0.318 | 0.00 | 0.484 | -0.26 | 0.111 | 0.06 | 0.155 | 271 |
| Child Used Alcohol, Marijuana, or Tobacco in Last 30 Days | 0.83 | 0.198 | 0.05 | 0.217 | 0.02 | 0.431 | 0.07 | 0.273 | 0.01 | 0.469 | 0.01 | 0.424 | 268 |
| Standardized Child BMI (Year 12) | 1.06 | 0.016 | 0.02 | 0.209 | -0.11 | 0.109 | 0.19 | 0.060 | -0.08 | 0.162 | -0.08 | 0.145 | 272 |

Notes: The first column provides the outcome description and the top row provides information on the mediators. For Year 12, the mediators are treatment, cognition, rows are divided into 3 groups: Outcome Coefficients, Treatment Effect and Treatment Effect Fraction. The last of these groups is also shown visually in Figure 7. Each
 We used the following controls: maternal race, maternal age, maternal height, gestational age, household density, region, employment status of household head, grandmother support, randomization wave, income category, mother currently in school, and maternal parenting attitudes.
Table H.4: Male Decomposition (Year 12)

|  | Treatment |  | Cognition |  | Attention problems |  | Conduct Problems |  | Warmth/Empathy |  | Aggression |  | Sample Size |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coefficient | $p$-value | Coefficient | $p$-value | Coefficient | $p$-value | Coefficient | $p$-value | Coefficient | $p$-value | Coefficient | $p$-value |  |
| Outcome Coefficients |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Average TCAP percentile, y1-5: language composite | 2.88 | 0.188 | 11.93 | 0.000 | 2.21 | 0.379 | -0.94 | 0.442 | 2.40 | 0.178 | -3.88 | 0.170 | 222 |
| PIAT reading comprehension derived score | 1.69 | 0.134 | 7.44 | 0.000 | -1.41 | 0.326 | 0.45 | 0.480 | 0.31 | 0.391 | 0.57 | 0.371 | 272 |
| Average math grade grades 1-5 | 0.03 | 0.368 | 0.50 | 0.000 | -0.21 | 0.145 | 0.22 | 0.137 | 0.05 | 0.288 | -0.13 | 0.118 | 243 |
| Average math grade. Years 1-5 after KG | -0.00 | 0.538 | 0.50 | 0.000 | -0.15 | 0.205 | 0.16 | 0.198 | 0.03 | 0.324 | -0.14 | 0.096 | 246 |
| average tcap percentile y1-5: math | 0.81 | 0.348 | 15.29 | 0.000 | 5.32 | 0.231 | -2.86 | 0.336 | -0.98 | 0.380 | -3.47 | 0.204 | 223 |
| PIAT math derived score | 1.64 | 0.106 | 7.86 | 0.000 | -1.36 | 0.324 | 0.20 | 0.515 | 0.71 | 0.232 | 1.53 | 0.146 | 270 |
| SC ever tried smoking: $1=$ yes | -0.05 | 0.079 | 0.02 | 0.254 | 0.04 | 0.279 | 0.02 | 0.341 | -0.02 | 0.192 | 0.01 | 0.404 | 274 |
| SC use alc, mar, tob last 30 days | -0.05 | 0.04 | -0.00 | 0.40 | -0.01 | 0.41 | 0.04 | 0.25 | -0.02 | 0.07 | 0.03 | 0.20 | 272 |
| Internalizing disorders - Youth report | -0.05 | 0.213 | -0.07 | 0.047 | 0.02 | 0.421 | 0.03 | 0.406 | 0.03 | 0.295 | 0.10 | 0.072 | 274 |
| Anxious/depressed - clinical or borderline disorder, youth report | -0.05 | 0.082 | -0.05 | 0.016 | -0.05 | 0.167 | 0.06 | 0.101 | 0.01 | 0.332 | 0.07 | 0.083 | 273 |
| Average number of absences, school years 1-5 | -1.05 | 0.146 | -2.25 | 0.001 | 4.36 | 0.010 | -3.87 | 0.009 | 0.22 | 0.372 | -1.10 | 0.156 | 267 |
| Treatment Effect |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Average TCAP percentile, y1-5: language composite | 2.88 | 0.188 | 2.09 | 0.064 | -0.02 | 0.432 | 0.03 | 0.361 | -0.39 | 0.135 | 0.45 | 0.161 | 222 |
| PIAT reading comprehension derived score | 1.69 | 0.134 | 1.44 | 0.041 | 0.11 | 0.255 | -0.03 | 0.364 | -0.04 | 0.313 | -0.07 | 0.270 | 272 |
| Average math grade grades 1-5 | 0.03 | 0.368 | 0.08 | 0.080 | -0.00 | 0.366 | 0.00 | 0.449 | -0.01 | 0.200 | 0.02 | 0.107 | 243 |
| Average math grade. Years 1-5 after KG | -0.00 | 0.538 | 0.08 | 0.061 | 0.00 | 0.451 | -0.00 | 0.335 | -0.00 | 0.230 | 0.02 | 0.121 | 246 |
| average tcap percentile y1-5: math | 0.81 | 0.348 | 2.68 | 0.070 | -0.09 | 0.367 | 0.09 | 0.313 | 0.16 | 0.303 | 0.41 | 0.203 | 223 |
| PIAT math derived score | 1.64 | 0.106 | 1.55 | 0.042 | 0.11 | 0.242 | -0.01 | 0.420 | -0.08 | 0.169 | -0.18 | 0.102 | 270 |
| SC ever tried smoking: $1=$ yes | -0.05 | 0.079 | 0.00 | 0.198 | -0.00 | 0.234 | -0.00 | 0.334 | 0.00 | 0.194 | -0.00 | 0.324 | 274 |
| SC use alc, mar, tob last 30 days | -0.05 | 0.043 | -0.00 | 0.342 | 0.00 | 0.328 | -0.00 | 0.262 | 0.00 | 0.144 | -0.00 | 0.170 | 272 |
| Internalizing disorders - Youth report | -0.05 | 0.213 | -0.01 | 0.058 | -0.00 | 0.342 | -0.00 | 0.303 | -0.00 | 0.245 | -0.01 | 0.116 | 274 |
| Anxious/depressed - clinical or borderline disorder, youth report | -0.05 | 0.082 | -0.01 | 0.028 | 0.00 | 0.213 | -0.00 | 0.219 | -0.00 | 0.251 | -0.01 | 0.085 | 273 |
| Average number of absences, school years 1-5 | -1.05 | 0.146 | -0.36 | 0.063 | -0.27 | 0.258 | 0.07 | 0.424 | -0.03 | 0.272 | 0.14 | 0.127 | 267 |
| Treatment Effect Fraction |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Average TCAP percentile, y1-5: language composite | 0.57 | 0.188 | 0.41 | 0.064 | -0.00 | 0.432 | 0.01 | 0.361 | -0.08 | 0.135 | 0.09 | 0.161 | 222 |
| PIAT reading comprehension derived score | 0.54 | 0.134 | 0.46 | 0.041 | 0.03 | 0.255 | -0.01 | 0.364 | -0.01 | 0.313 | -0.02 | 0.270 | 272 |
| Average math grade. Years 1-5 after KG | 0.23 | 0.368 | 0.68 | 0.080 | -0.03 | 0.366 | 0.01 | 0.449 | -0.05 | 0.200 | 0.16 | 0.107 | 243 |
| average tcap percentile y1-5: math | 0.20 | 0.348 | 0.66 | 0.070 | -0.02 | 0.367 | 0.02 | 0.313 | 0.04 | 0.303 | 0.10 | 0.203 | 223 |
| PIAT math derived score | 0.54 | 0.106 | 0.51 | 0.042 | 0.04 | 0.242 | -0.00 | 0.420 | -0.03 | 0.169 | -0.06 | 0.102 | 270 |
| SC use alc, mar, tob last 30 days | 0.92 | 0.043 | 0.02 | 0.342 | -0.02 | 0.328 | 0.04 | 0.262 | -0.04 | 0.144 | 0.08 | 0.170 | 272 |
| Internalizing disorders - Youth report | 0.63 | 0.213 | 0.17 | 0.058 | 0.01 | 0.342 | 0.02 | 0.303 | 0.03 | 0.245 | 0.14 | 0.116 | 274 |
| Anxious/depressed - clinical or borderline disorder, youth report | 0.71 | 0.082 | 0.14 | 0.028 | -0.05 | 0.213 | 0.05 | 0.219 | 0.02 | 0.251 | 0.12 | 0.085 | 273 |
| Average number of absences, school years 1-5 | 0.70 | 0.146 | 0.24 | 0.063 | 0.18 | 0.258 | -0.05 | 0.424 | 0.02 | 0.272 | -0.09 | 0.127 | 267 |

Notes: The first column provides the outcome description and the top row provides information on the mediators. For Year 12 , the mediators are treatment, cognition,
 mediator has two subcolumns of information: the coefficient and the $p$-value. Bold $p$-values are significant at the $10 \%$ level. We used the following controls: maternal race, maternal age, maternal height, gestational age, household density, region, employment status of household head, grandmother support, randomization wave, income category, mother currently in school, and maternal parenting attitudes.

## I Mediation Specification Tests

In this section we specify how do we empirically test the effect that the mediators have on the final outcomes. We use $\mathcal{J}$ for an indexing set of skills. We use $\mathcal{J}_{p} \subseteq \mathcal{J}$ for the subset of measured skills. Our model for the outcome equation is:

$$
Y_{d}=\kappa_{d}+\sum_{j \in \mathcal{J}} \alpha_{d}^{j} \theta_{d}^{j}+\boldsymbol{\beta}_{d} \boldsymbol{X}+\tilde{\epsilon}_{d}, \quad d \in\{0,1\}
$$

where $\kappa_{d}$ is an intercept, $\left(\alpha_{d}^{j} ; j \in \mathcal{J}\right)$ are loading factors and $\boldsymbol{\beta}_{d}$ are $|\boldsymbol{X}|$-dimensional vectors of parameters. The error term $\tilde{\epsilon}_{d}$ is a zero-mean i.i.d. random variable assumed to be independent of regressors $\left(\theta_{d}^{j} ; j \in \mathcal{J}\right)$ and $\boldsymbol{X}$.

The NFP analysts collected a rich array of measures of cognitive and personality skills. However, it is likely that there are skills that they did not measure. As noted before, we use $\mathcal{J}_{p} \subseteq \mathcal{J}$ be the index set of measured skills. Namely, skills for which we have enough psychological instruments for estimation. We rewrite the equation for scalar potential outcome $Y_{d}$ as:

$$
\begin{align*}
Y_{d} & =\kappa_{d}+\sum_{j \in \mathcal{J}} \alpha_{d}^{j} \theta_{d}^{j}+\boldsymbol{\beta}_{d} \boldsymbol{X}+\tilde{\epsilon}_{d} \\
& =\kappa_{d}+\underbrace{\sum_{j \in \mathcal{J}_{p}} \alpha_{d}^{j} \theta_{d}^{j}}_{\substack{\text { skills that } \\
\text { we measure }}}+\underbrace{\sum_{j \in \mathcal{J} \backslash \mathcal{J}_{p}} \alpha^{j} \theta_{d}^{j}}_{\substack{\text { skills that we } \\
\text { do not measure }}}+\boldsymbol{\beta}_{d} \boldsymbol{X}+\tilde{\epsilon}_{d} \\
& =\underbrace{\kappa_{d}+\sum_{j \in \mathcal{J} \backslash \mathcal{J}_{p}} \alpha_{d}^{j} \mathrm{E}\left(\theta_{d}^{j}\right)}_{\text {new intercept }}+\underbrace{\sum_{j \in \mathcal{J}_{p}} \alpha_{d}^{j} \theta_{d}^{j}}_{\substack{\text { skills that } \\
\text { we measure }}}+\underbrace{\sum_{j \in \mathcal{J} \backslash \mathcal{J}_{p}} \alpha_{d}^{j}\left(\theta_{d}^{j}-\mathrm{E}\left(\theta_{d}^{j}\right)\right)}_{\begin{array}{c}
\text { skills that we } \\
\text { do not measure }
\end{array}}+\boldsymbol{\beta}_{d} \boldsymbol{X}+\tilde{\epsilon}_{d}, \\
& =\underbrace{\tau_{d}}_{\text {new intercept }}+\underbrace{\sum_{j \in \mathcal{J}_{p}} \alpha_{d}^{j} \theta_{d}^{j}}_{\substack{\text { skills that } \\
\text { we measure }}}+\boldsymbol{\beta}_{d} \boldsymbol{X}+\underbrace{\sum_{j \in \mathcal{J} \backslash \mathcal{J}_{p}} \alpha_{d}^{j}\left(\theta_{d}^{j}-\mathrm{E}\left(\theta_{d}^{j}\right)\right)+\tilde{\epsilon}_{d}}_{\text {new error term }} \tag{35}
\end{align*}
$$

where $d \in\{0,1\}, \tau_{d}=\kappa_{d}+\sum_{j \in \mathcal{J} \backslash \mathcal{J}_{p}} \alpha_{d}^{j} \mathrm{E}\left(\theta_{d}^{j}\right)$.

Any differences in the error terms between treatment and control groups can be attributed to differences in unmeasured skills. Thus, we assume, without loss of generality, that $\tilde{\epsilon}_{1} \stackrel{d}{=} \tilde{\epsilon}_{0}$, where $\stackrel{d}{=}$ means equality in distribution.

The goal of this section is to examine the statistical assumptions needed to estimate unbiased parameters $\left(\alpha_{d}^{j}: j \in \mathcal{J}_{p}, d \in\{0,1\}\right)$. These parameters are used to perform the decomposition of outcome treatment effects into parts associated with skills enhancement $\left(\theta_{1}^{j}-\theta_{0}^{j}: j \in \mathcal{J}_{p}\right)$. Parameters $\alpha$ may suffer from confounding effects if measured and unmeasured skills are not independent. We can solve this confounding problem by assuming that unmeasured skills are independent of measured skills. Namely,

$$
\left(\theta_{d}^{j} ; j \in \mathcal{J} \backslash \mathcal{J}_{p}\right) \Perp\left(\theta_{d}^{j} ; j \in \mathcal{J}_{p}\right) \mid \boldsymbol{X} ; d \in\{0,1\}
$$

then the regression:

$$
\begin{equation*}
Y_{d}=\tau_{d}+\sum_{j \in \mathcal{J}_{p}} \alpha_{d}^{j} \theta_{d}^{j}+\boldsymbol{\beta}_{d} \boldsymbol{X}+\epsilon_{d} \tag{36}
\end{equation*}
$$

produces unbiased estimates of parameter $\left(\alpha_{d}^{j} ; j \in \mathcal{J}_{p}\right) ; d \in\{0,1\}$. Indeed error terms $\epsilon_{d}$ in equation (36) are given by

$$
\epsilon_{d}=\tilde{\epsilon}_{d}+\sum_{j \in \mathcal{J} \backslash \mathcal{J}_{p}} \alpha_{d}^{j}\left(\theta_{d}^{j}-\mathrm{E}\left(\theta_{d}^{j}\right)\right)
$$

which are independent of $\left(\theta_{d}^{j} ; j \in \mathcal{J}_{p}\right)$ conditional on $\boldsymbol{X}$ under the assumption that skills are independent.

Now suppose that instead of the skills independence assumption for both groups, we focus only on the control group, thus,

$$
\left(\theta_{0}^{j} ; j \in \mathcal{J} \backslash \mathcal{J}_{p}\right) \Perp\left(\theta_{0}^{j} ; j \in \mathcal{J}_{p}\right) \mid \boldsymbol{X} .
$$

Moreover, suppose we also assume that $\alpha_{1}^{j}=\alpha_{0}^{j} ; j \in \mathcal{J}$. Equivalently, the outcome
loading factors for both treatment and control groups are the same. In this new setup, the regression

$$
\begin{equation*}
Y_{0}=\tau_{0}+\sum_{j \in \mathcal{J}_{p}} \alpha^{j} \theta_{0}^{j}+\boldsymbol{\beta}_{0} \boldsymbol{X}+\epsilon_{0}, \tag{37}
\end{equation*}
$$

also produces unbiased estimates of $\left(\alpha^{j} ; j \in \mathcal{J}_{p}\right)$. Now consider the regression

$$
Y_{1}=\tau_{1}+\sum_{j \in \mathcal{J}_{p}} \alpha^{j} \theta_{1}^{j}+\boldsymbol{\beta}_{1} \boldsymbol{X}+\epsilon_{1} .
$$

According to our rationale, this regression only produces unbiased estimates of ( $\alpha^{j} ; j \in$ $\left.\mathcal{J}_{p}\right)$ if:

$$
\begin{equation*}
\left(\theta_{1}^{j} ; j \in \mathcal{J} \backslash \mathcal{J}_{p}\right) \Perp\left(\theta_{1}^{j} ; j \in \mathcal{J}_{p}\right) \mid \boldsymbol{X}, \tag{38}
\end{equation*}
$$

or, alternatively,

$$
\begin{equation*}
\left(\theta_{1}^{j}-\theta_{0}^{j} ; j \in \mathcal{J} \backslash \mathcal{J}_{p}\right) \Perp\left(\theta_{1}^{j}-\theta_{0}^{j} ; j \in \mathcal{J}_{p}\right) \mid \boldsymbol{X} . \tag{39}
\end{equation*}
$$

Thus, under this new set of assumptions, testing $H_{0}: \boldsymbol{\alpha}_{1}=\boldsymbol{\alpha}_{0}$ is translated into testing the independence relations of equations (38)-(39).

While the skill independence assumption in equation (38) may appear strong, the rich settlement of information on NFP surveys makes this assumption more plausible. NFP data has a huge selection of psychological questionnaires that aims to measure both cognitive and non-cognitive skills though childhood. We examine all the available data and only a subset of these measures turns out to be statistically relevant for mediation analysis. We use these measures to estimate factors that are able to explain the majority of the treatment effects. Thus, it seems unlikely that some unobserved skills overlooked by psychologists could have a major impact on mediating treatment effects.

## I. 1 Skills and the Measurement System

The assumption that the loading factors in the measurement system (Equation 27) are the same for treatment and control is not necessary to identify the model. It is useful for clarity
in the interpretation because the treatment operates by the shift in latent skills and not by the map between measures and skills.

Ultimately, we need the decomposition of the treatment effects, (6), to be invariant to the choice of the measurement system we used. Thus, for each skill's contribution to treatment effect on each outcome, we want to test the null hypothesis that:

$$
\begin{equation*}
H_{0}: \boldsymbol{\alpha}_{0}\left(\mathrm{E}\left(\boldsymbol{\theta}_{1}-\boldsymbol{\theta}_{0}\right)\right)=\boldsymbol{\alpha}_{1}\left(\mathrm{E}\left(\boldsymbol{\theta}_{1}-\boldsymbol{\theta}_{0}\right)\right) \tag{40}
\end{equation*}
$$

where $\boldsymbol{\alpha}_{d}=\left(\alpha_{d}^{j}: j \in \mathcal{J}_{p}\right)$ and $\boldsymbol{\theta}_{d}=\left(\theta_{d}^{j}: j \in \mathcal{J}_{p}\right)$ such that $d \in\{0,1\}$ denotes treatment status.

Let $\hat{\boldsymbol{\theta}}_{i}$ be the estimated factor score for individual $i$, assigned to treatment status $D_{i} \in$ $\{0,1\}$, using the estimated loading factors from the subsample of individuals with the same treatment status, i.e. for each individual factor score:

$$
\hat{\boldsymbol{\theta}}_{i}=\left(\boldsymbol{\varphi}_{D_{i}}{ }^{\prime}\left(\boldsymbol{\Omega}_{D_{i}}\right)^{-1} \boldsymbol{\varphi}_{D_{i}}\right)^{-1} \boldsymbol{\varphi}_{D_{i}}{ }^{\prime}\left(\boldsymbol{\Omega}_{D_{i}}\right)^{-1} \boldsymbol{M}_{i}
$$

We would like to test whether the contribution to the treatment effects is independent if we use the parameters from a different measurement system (i.e if we estimate a different set of loading factors for the treatment and control group).

Hence, an appropriate single hypothesis test statistic for each skill $j \in \mathcal{J}_{p}$ becomes:

$$
\hat{\alpha}_{0}^{j}\left(\hat{\theta}_{1}^{j}-\hat{\theta}_{0}^{j}\right)-\hat{\alpha}_{1}^{j}\left(\hat{\theta}_{1}^{j}-\hat{\theta}_{0}^{j}\right)
$$

where we use a hat superscript to denote estimated parameters. $\hat{\alpha}$ are Croon corrected estimates of $\alpha$. We can use a summary statistic to test the joint hypothesis stated in (40).

Independence between $\hat{\boldsymbol{\alpha}}_{d}$ and $\hat{\boldsymbol{\theta}}_{d}-\hat{\boldsymbol{\theta}}_{d}$ yields:

$$
\operatorname{Var}\left(\hat{\boldsymbol{\alpha}}_{d}\left(\hat{\boldsymbol{\theta}}_{1}-\hat{\boldsymbol{\theta}}_{0}\right)\right)=\left(\hat{\boldsymbol{\alpha}}_{d}\right)^{2} \operatorname{Var}\left(\hat{\boldsymbol{\theta}}_{1}-\hat{\boldsymbol{\theta}}_{0}\right)+\operatorname{Var}\left(\hat{\boldsymbol{\alpha}}_{d}\right)\left(\hat{\boldsymbol{\theta}}_{1}-\hat{\boldsymbol{\theta}}_{0}\right)^{2}+\operatorname{Var}(\hat{\boldsymbol{\alpha}}) \operatorname{Var}\left(\hat{\boldsymbol{\theta}}_{1}-\hat{\boldsymbol{\theta}}_{0}\right)
$$

Independence between the quantities estimated for each of the d's yields:

$$
\operatorname{Var}\left(\hat{\boldsymbol{\alpha}}_{0}\left(\hat{\boldsymbol{\theta}}_{1}-\hat{\boldsymbol{\theta}}_{0}\right)-\hat{\boldsymbol{\alpha}}^{1}\left(\hat{\boldsymbol{\theta}}_{1}-\hat{\boldsymbol{\theta}}_{0}\right)\right)=\operatorname{Var}\left(\hat{\boldsymbol{\alpha}}_{0}\left(\hat{\boldsymbol{\theta}}_{1}-\hat{\boldsymbol{\theta}}_{0}\right)\right)+\operatorname{Var}\left(\hat{\boldsymbol{\alpha}}_{1}\left(\overline{\hat{\boldsymbol{\theta}}}_{1}-\overline{\hat{\boldsymbol{\theta}}}_{0}\right)\right)
$$

This variance helps us to get the $z$-statistic:

$$
z=\frac{\hat{\boldsymbol{\alpha}}_{0}\left(\hat{\boldsymbol{\theta}}_{1}-\hat{\boldsymbol{\theta}}_{0}\right)-\hat{\boldsymbol{\alpha}}_{1}\left(\hat{\boldsymbol{\theta}}_{1}-\hat{\boldsymbol{\theta}}_{0}\right)}{\sqrt{\operatorname{Var}\left(\hat{\boldsymbol{\alpha}}_{0}\left(\hat{\boldsymbol{\theta}}_{1}-\hat{\boldsymbol{\theta}}_{0}\right)-\hat{\boldsymbol{\alpha}}^{1}\left(\hat{\boldsymbol{\theta}}_{1}-\hat{\boldsymbol{\theta}}_{0}\right)\right)}}
$$

A two-sided $z$-test gives a $p$-value associated with the skill and outcome null hypothesis of invariance to the choice of the measurement system.

These paired (outcome, skill) $p$-values are shown in Tables I. 1 and I.2. We find that we can not reject the null hypothesis for any skill-outcome pair, which suggests that our decompositions of the NFP treatment effects are not driven by the choice of the measurement system.

## I.1.1 Additional Specification Tests for the Outcome Equations

In order to clearly interpret the channels through which the NFP affects later outcomes, (1) assumes that the parameters that map skills and pre-program variables with the outcomes are not affected by the programs. Put another way, the mediated channels operate exclusively through the program effect on skills. This assumption is not necessary to identify the model.

For each outcome decomposed, we test the hypothesis that $\alpha_{1}^{j}=\alpha_{0}^{j}, \forall j \in \mathcal{J}$ and $\boldsymbol{\beta}_{\mathbf{1}}=\boldsymbol{\beta}_{\mathbf{0}}$ with a Wald test. Tables I. 3 and I. 4 show the results of this test. We cannot reject the null hypothesis of equality of the coefficients for the treatment and control groups. This evidence strengthens the validity of our interpretation of the decomposition of the NFP treatment effect.

## J Oaxaca-Blinder Decomposition Results

Oaxaca-Blinder decompositions are often used to examine sources of treatment effects. This method decomposes the difference in means between two groups (treatment and control) into the part that is due to the group differences in the channels and into the part that is due to group differences in the parameters that capture the relationship between the channels and the outcomes. In our context, the Oaxaca-Blinder decomposition is summarized as follows: ${ }^{9}$

$$
\begin{equation*}
\underbrace{E(\boldsymbol{Y} \mid D=1)-\mathrm{E}(\boldsymbol{Y} \mid D=0)}_{\text {Treatment Effects }}=\underbrace{\left(\boldsymbol{\alpha}_{1}-\boldsymbol{\alpha}_{0}\right) \boldsymbol{\theta}_{0}}_{\text {Differences unexplained by the skills }}+\underbrace{\left(\boldsymbol{\theta}_{1}-\boldsymbol{\theta}_{0}\right) \boldsymbol{\alpha}}_{\text {Differences explained by the skills }} . \tag{41}
\end{equation*}
$$

The decomposition that we propose summarizes the unexplained part in the above equation through the difference in the intercepts between the treatment and the control groups. In order to assess whether our decomposition is a plausible specification, we estimate an Oaxaca-Blinder decomposition. The results in Tables J.1- J. 5 present evidence that the unexplained component accounting for differences in the mapping of the skills on outcomes is not statistically significant for any outcome. Therefore, the results from the decomposition of the NFP treatment effects presented in the paper seem to be correctly specified.

[^7]Table I.1: Specification Test - Invariance of the Contribution of Skills to the Choice of the Measurement System (Females)
Factor Testing Results - Females
Maternal Skills Age 2

| Age 6 outcomes | Home | Parenting | anxiety | esteem | mastery |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Cognition | 0.263 | 0.907 | 0.859 | 0.698 | 0.672 |
| Attention problems | 0.363 | 0.709 | 0.702 | 0.667 | 0.748 |
| Conduct problems | 0.421 | 0.694 | 0.922 | 0.721 | 0.677 |
| Warmth-empathy (pro-social | 0.267 | 0.907 | 0.644 | 0.833 | 0.973 |
| skills) | 0.692 | 0.819 | 0.862 | 0.821 | 0.786 |
| Agression Problems |  |  |  |  |  |

[^8]Notes: The table shows p-values for the Wald test: $z=\frac{\hat{\alpha}^{0}\left(\overline{\hat{\theta}}_{1}^{0}-\overline{\hat{\theta}}_{0}^{0}\right)-\hat{\alpha}^{1}\left(\overline{\hat{\theta}}_{1}^{1}-\overline{\hat{\theta}}_{0}^{1}\right)}{\sqrt{\operatorname{Var}\left(\hat{\alpha}^{0}\left(\hat{\theta}_{1}^{0} \tilde{\hat{\theta}}_{0}^{0}\right)-\hat{\alpha}^{1}\left(\overline{\hat{\theta}}_{1}^{1}\right.\right.} \overline{\left.\hat{\theta}_{0}^{1}\right)}}$
Table I.2: Specification Test - Invariance of the Contribution of Skills to the Choice of the Measurement System (Males)

Maternal Skills Age 2

| Age 6 outcomes | Home | Parenting | anxiety | esteem | mastery |
| :--- | :---: | :---: | ---: | ---: | ---: |
| Cognition | 0.349 | 0.394 | 0.971 | 0.927 | 0.946 |
| Agression Problems | 0.928 | 0.959 | 0.950 | 0.843 | 0.537 |

[^9][^10]| Outcome |  | Test Stat | $P$-Val |
| :---: | :---: | :---: | :---: |
| 6 Years |  |  |  |
|  | Cognition | 0.982 | 0.490 |
|  | Attention prob. | 1.753 | 0.018 |
|  | Conduct Prob. | 0.846 | 0.675 |
|  | Pro-social | 1.264 | 0.189 |
|  | Aggression | 0.558 | 0.955 |


| SC use alc, mar, tob last 30 days |  | 1.266 | 0.189 |
| ---: | :--- | :--- | :--- |
| SC \# days use of alc, mar, tob last 30 days | 1.271 | 0.186 |  |
| Standardized Child BMI (Year 12) | 1.172 | 0.270 |  |

Notes: The table shows p-values for Wald tests for the equality of slopes between treatment and control group in the outcome equation

| Outcome | Test Stat | $P$-Val |
| :---: | :---: | :---: |
| 6 Years |  |  |
| Cognition | 0.609 | 0.926 |
| Aggression | 0.881 | 0.628 |
| 12 Years |  |  |
| average tcap percentile, y1-5: language composite | 1.162 | 0.283 |
| PIAT reading comprehension derived score | 1.286 | 0.175 |
| Average math grade. Years 1-5 after KG | 1.493 | 0.073 |
| average tcap percentile y1-5: math | 1.242 | 0.213 |
| PIAT math derived score | 1.102 | 0.343 |
| SC ever tried smoking: 1=yes | 0.838 | 0.686 |
| Internalizing disorders - Youth report | 0.993 | 0.477 |
| Anxious/depressed - clinical or borderline disorder | 0.682 | 0.867 |
| Average number of absences, school years 1-5 | 0.798 | 0.738 |

Notes: The table shows p-values for Wald tests for the equality of slopes between treatment and control group in the outcome equation
Table J.1: Oaxaca-Blinder Decomposition, outcomes at age 6 (Females)

| Aggression |  |  |  |
| :---: | :---: | :---: | :---: |
| Effect | SE | $P$-Val | Fraction |
|  |  |  |  |
| -0.187 | 0.094 | $\mathbf{0 . 0 4 6}$ | - |
| -0.024 | 0.030 | 0.421 | 0.260 |
| -0.163 | 0.090 | $\mathbf{0 . 0 7 1}$ | 0.740 |
|  |  |  |  |
|  |  |  |  |
| -0.015 | 0.015 | 0.344 | 0.160 |
| 0.002 | 0.019 | 0.935 | 0.015 |
| 0.009 | 0.014 | 0.536 | -0.116 |
| -0.022 | 0.021 | 0.283 | 0.131 |
| -0.003 | 0.023 | 0.893 | 0.096 |
| 0.005 | 0.007 | 0.483 | -0.027 |
|  |  |  |  |
|  |  |  |  |
| -0.009 | 0.019 | 0.624 | 0.051 |
| 0.010 | 0.019 | 0.615 | -0.051 |
| 0.005 | 0.016 | 0.757 | -0.026 |
| 0.006 | 0.024 | 0.802 | -0.033 |
| 0.001 | 0.029 | 0.974 | -0.005 |
| 0.002 | 0.005 | 0.749 | -0.009 |
| -0.176 | 0.091 | $\mathbf{0 . 0 5 2}$ | 0.943 |

Table J.2: Oaxaca-Blinder Decomposition, outcomes at age 6 (Males)

|  | Cognition |  |  |  | Attention Problems |  |  |  | Conduct Problems |  |  |  | Warmth/Empathy |  |  |  | Aggression |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Effect | SE | $P$-Val | Fraction | Effect | SE | $P$-Val | Fraction | Effect | SE | $P$-Val | Fraction | Effect | SE | $P$-Val | Fraction | Effect | SE | $P$-Val | Fraction |
| Overall |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Total Diff. in Means | 0.173 | 0.105 | 0.100 | - | -0.025 | 0.095 | 0.797 | - | -0.012 | 0.088 | 0.891 | - | -0.080 | 0.104 | 0.442 | - | -0.099 | 0.086 | 0.250 | - |
| Explained | 0.092 | 0.037 | 0.012 | 0.496 | -0.064 | 0.033 | 0.051 | -5.856 | -0.056 | 0.035 | 0.111 | -1.130 | 0.022 | 0.025 | 0.387 | -0.254 | -0.025 | 0.025 | 0.322 | 0.158 |
| Unexplained | 0.081 | 0.101 | 0.422 | 0.504 | 0.039 | 0.092 | 0.670 | 6.856 | 0.044 | 0.090 | 0.628 | 2.130 | -0.102 | 0.105 | 0.332 | 1.254 | -0.074 | 0.085 | 0.383 | 0.842 |
| Explained Portion |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Home Index | 0.021 | 0.022 | 0.344 | 0.220 | -0.003 | 0.006 | 0.627 | -0.303 | -0.009 | 0.010 | 0.395 | -0.395 | 0.006 | 0.009 | 0.513 | -0.161 | 0.002 | 0.005 | 0.728 | -0.074 |
| Parenting Index | 0.044 | 0.022 | 0.042 | 0.112 | -0.021 | 0.019 | 0.258 | -0.479 | -0.011 | 0.018 | 0.540 | -0.041 | 0.008 | 0.017 | 0.640 | -0.007 | -0.019 | 0.013 | 0.150 | 0.085 |
| Maternal Anxiety Index | 0.001 | 0.004 | 0.847 | 0.019 | -0.007 | 0.012 | 0.561 | -0.544 | -0.008 | 0.014 | 0.546 | -0.297 | 0.003 | 0.007 | 0.629 | -0.046 | -0.011 | 0.018 | 0.538 | 0.125 |
| Maternal Self-Esteem Index | -0.005 | 0.009 | 0.607 | -0.016 | 0.005 | 0.010 | 0.592 | 1.536 | -0.011 | 0.015 | 0.490 | -0.514 | -0.003 | 0.009 | 0.690 | 0.034 | 0.012 | 0.014 | 0.399 | -0.230 |
| Maternal Mastery Index | 0.009 | 0.018 | 0.598 | 0.021 | -0.015 | 0.018 | 0.385 | -3.978 | -0.001 | 0.014 | 0.964 | 0.553 | 0.004 | 0.018 | 0.831 | -0.020 | -0.003 | 0.013 | 0.833 | 0.197 |
| Birthweight | 0.022 | 0.017 | 0.192 | 0.140 | -0.023 | 0.018 | 0.214 | -2.089 | -0.016 | 0.015 | 0.283 | -0.435 | 0.004 | 0.014 | 0.747 | -0.054 | -0.007 | 0.010 | 0.534 | 0.055 |
| Unexplained Portion |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Home Index | 0.002 | 0.011 | 0.834 | 0.013 | -0.003 | 0.009 | 0.763 | 0.111 | 0.002 | 0.009 | 0.846 | -0.152 | 0.008 | 0.012 | 0.529 | -0.096 | -0.002 | 0.009 | 0.836 | 0.019 |
| Parenting Index | -0.007 | 0.024 | 0.760 | -0.043 | 0.012 | 0.024 | 0.627 | -0.471 | 0.016 | 0.022 | 0.467 | -1.314 | 0.010 | 0.027 | 0.698 | -0.129 | 0.034 | 0.021 | 0.110 | -0.340 |
| Maternal Anxiety Index | 0.006 | 0.012 | 0.622 | 0.035 | 0.000 | 0.006 | 0.945 | 0.016 | -0.005 | 0.011 | 0.625 | 0.451 | 0.003 | 0.009 | 0.708 | -0.042 | -0.004 | 0.008 | 0.627 | 0.040 |
| Maternal Self-Esteem Index | 0.004 | 0.010 | 0.665 | 0.024 | -0.002 | 0.008 | 0.771 | 0.091 | 0.004 | 0.010 | 0.688 | -0.321 | 0.005 | 0.010 | 0.642 | -0.059 | 0.000 | 0.006 | 0.970 | -0.002 |
| Maternal Mastery Index | -0.033 | 0.028 | 0.241 | -0.192 | -0.007 | 0.026 | 0.780 | 0.294 | 0.020 | 0.023 | 0.396 | -1.627 | -0.023 | 0.029 | 0.421 | 0.288 | 0.029 | 0.026 | 0.268 | -0.289 |
| Birthweight | 0.020 | 0.024 | 0.418 | 0.113 | -0.060 | 0.028 | 0.034 | 2.422 | -0.065 | 0.028 | 0.022 | 5.335 | -0.015 | 0.024 | 0.525 | 0.191 | -0.003 | 0.018 | 0.877 | 0.028 |
| Residual | 0.090 | 0.107 | 0.400 | 0.519 | 0.100 | 0.093 | 0.284 | -4.061 | 0.072 | 0.098 | 0.462 | -5.970 | -0.090 | 0.116 | 0.436 | 1.121 | -0.128 | 0.076 | 0.091 | 1.293 |

Table J.3: Oaxaca-Blinder Decomposition, outcomes at age 12 (Females)

|  | SC \# days ever used marijuana |  |  |  | SC use alc, mar, tob last 30 days |  |  |  | Standardized Child BMI (12Y) |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Effect | SE | $P$-Val | Fraction | Effect | SE | $P$-Val | Fraction | Effect | SE | $P$-Val | Fraction |
| Total Diff. in Means | -0.141 | 0.085 | 0.098 | - | -0.021 | 0.020 | 0.281 | - | -0.197 | 0.110 | 0.072 | - |
| Explained | 0.026 | 0.038 | 0.488 | -0.184 | -0.005 | 0.007 | 0.460 | 0.242 | 0.020 | 0.038 | 0.601 | -0.102 |
| Unexplained | -0.167 | 0.111 | 0.133 | 1.184 | -0.016 | 0.021 | 0.429 | 0.758 | -0.217 | 0.114 | 0.056 | 1.102 |
| Explained |  |  |  |  |  |  |  |  |  |  |  |  |
| Cognitive | -0.006 | 0.018 | 0.723 | 0.046 | -0.001 | 0.003 | 0.708 | 0.058 | -0.004 | 0.013 | 0.777 | 0.019 |
| Attention Problems | 0.000 | 0.013 | 0.991 | 0.001 | -0.001 | 0.005 | 0.827 | 0.051 | 0.016 | 0.017 | 0.354 | -0.081 |
| Conduct Problems | -0.003 | 0.010 | 0.775 | 0.019 | -0.003 | 0.004 | 0.468 | 0.149 | -0.027 | 0.022 | 0.215 | 0.138 |
| Warmth/Empathy | 0.038 | 0.038 | 0.318 | -0.268 | 0.000 | 0.004 | 0.934 | -0.015 | 0.023 | 0.023 | 0.331 | -0.115 |
| Aggression | -0.003 | 0.005 | 0.644 | 0.018 | 0.000 | 0.002 | 0.988 | -0.001 | 0.013 | 0.021 | 0.554 | -0.064 |
| Unexplained |  |  |  |  |  |  |  |  |  |  |  |  |
| Cognitive | 0.012 | 0.018 | 0.523 | -0.084 | 0.003 | 0.004 | 0.496 | -0.135 | 0.004 | 0.015 | 0.773 | -0.022 |
| Attention Problems | 0.002 | 0.015 | 0.889 | -0.015 | 0.008 | 0.007 | 0.288 | -0.364 | -0.036 | 0.034 | 0.293 | 0.181 |
| Conduct Problems | 0.013 | 0.018 | 0.482 | -0.091 | -0.008 | 0.012 | 0.522 | 0.364 | 0.095 | 0.047 | 0.043 | -0.481 |
| Warmth/Empathy | -0.029 | 0.032 | 0.374 | 0.202 | -0.005 | 0.004 | 0.254 | 0.228 | -0.026 | 0.023 | 0.260 | 0.132 |
| Aggression | -0.002 | 0.009 | 0.803 | 0.016 | -0.001 | 0.003 | 0.697 | 0.062 | 0.013 | 0.037 | 0.724 | -0.066 |
| Residual | -0.163 | 0.096 | 0.089 | 1.155 | -0.013 | 0.027 | 0.629 | 0.603 | -0.268 | 0.122 | 0.028 | 1.358 |

Notes: The indices are means of the non-missing items. The fractions are proportions of the total conditional difference in means.
Table J.4: Oaxaca-Blinder outcomes at age 12, Decomposition Part 1 (Males)

|  | language composite |  |  |  | derived score |  |  |  | Average math grade grades 1-5 |  |  |  | after KG |  |  |  | average tcap percentile y1-5: math |  |  |  | PIAT math derived score |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Effect | SE | $P$-Val | Fraction | Effect | SE | $P$-Val | Fraction | Effect | SE | $P$-Val | Fraction | Effect | SE | $P$-Val | Fraction | Effect | SE | $P$-Val | Fraction | Effect | SE | $P$-Val | Fraction |
| Total Diff. in Means | 4.403 | 3.289 | 0.181 | - | 2.022 | 1.584 | 0.202 | - | 0.117 | 0.107 | 0.275 | - | 0.083 | 0.103 | 0.422 | - | 2.818 | 3.429 | 0.411 | - | 2.447 | 1.324 | 0.065 | - |
| Explained | 1.564 | 1.520 | 0.304 | 0.355 | 0.941 | 0.808 | 0.244 | 0.466 | 0.077 | 0.063 | 0.219 | 0.660 | 0.067 | 0.058 | 0.242 | 0.814 | 2.226 | 1.820 | 0.221 | 0.790 | 0.946 | 0.817 | 0.247 | 0.387 |
| Unexplained | 2.839 | 3.105 | 0.361 | 0.645 | 1.080 | 1.454 | 0.457 | 0.534 | 0.040 | 0.090 | 0.658 | 0.340 | 0.015 | 0.087 | 0.860 | 0.186 | 0.592 | 3.111 | 0.849 | 0.210 | 1.501 | 1.103 | 0.174 | 0.613 |
| Explained |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Cognitive | 1.831 | 1.337 | 0.171 | 0.416 | 0.910 | 0.705 | 0.197 | 0.450 | 0.057 | 0.054 | 0.285 | 0.492 | 0.057 | 0.051 | 0.266 | 0.682 | 2.146 | 1.664 | 0.197 | 0.762 | 0.939 | 0.750 | 0.211 | 0.384 |
| Attention Problems | 0.075 | 0.275 | 0.787 | 0.017 | 0.060 | 0.131 | 0.647 | 0.030 | 0.009 | 0.016 | 0.570 | 0.080 | 0.007 | 0.012 | 0.548 | 0.087 | -0.057 | 0.312 | 0.855 | -0.020 | 0.067 | 0.123 | 0.589 | 0.027 |
| Conduct Problems | -0.003 | 0.273 | 0.992 | -0.001 | 0.037 | 0.122 | 0.765 | 0.018 | 0.000 | 0.010 | 0.964 | -0.004 | -0.001 | 0.008 | 0.933 | -0.008 | 0.020 | 0.281 | 0.945 | 0.007 | 0.035 | 0.110 | 0.751 | 0.014 |
| Warmth/Empathy | -0.535 | 0.468 | 0.253 | -0.122 | -0.079 | 0.156 | 0.611 | -0.039 | -0.009 | 0.013 | 0.475 | -0.079 | -0.012 | 0.014 | 0.362 | -0.149 | -0.063 | 0.385 | 0.870 | -0.022 | -0.094 | 0.122 | 0.441 | -0.038 |
| Aggression | 0.197 | 0.419 | 0.639 | 0.045 | 0.014 | 0.161 | 0.930 | 0.007 | 0.020 | 0.018 | 0.267 | 0.171 | 0.017 | 0.017 | 0.328 | 0.203 | 0.181 | 0.439 | 0.681 | 0.064 | -0.001 | 0.142 | 0.995 | 0.000 |
| Unexplained |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Cognitive | 0.540 | 0.980 | 0.582 | 0.123 | -0.015 | 0.288 | 0.958 | -0.007 | 0.005 | 0.019 | 0.787 | 0.045 | 0.009 | 0.019 | 0.652 | 0.105 | 0.563 | 0.780 | 0.470 | 0.200 | -0.016 | 0.228 | 0.943 | -0.007 |
| Attention Problems | 0.699 | 0.769 | 0.363 | 0.159 | 0.449 | 0.342 | 0.189 | 0.222 | 0.009 | 0.020 | 0.659 | 0.077 | 0.006 | 0.020 | 0.752 | 0.075 | 0.518 | 0.663 | 0.435 | 0.184 | -0.162 | 0.292 | 0.579 | -0.066 |
| Conduct Problems | 0.003 | 0.427 | 0.995 | 0.001 | -0.092 | 0.242 | 0.705 | -0.045 | -0.003 | 0.016 | 0.858 | -0.024 | -0.003 | 0.015 | 0.864 | -0.031 | 0.002 | 0.402 | 0.995 | 0.001 | 0.004 | 0.183 | 0.983 | 0.002 |
| Warmth/Empathy | 0.373 | 0.691 | 0.590 | 0.085 | 0.134 | 0.271 | 0.622 | 0.066 | 0.011 | 0.020 | 0.588 | 0.092 | 0.012 | 0.021 | 0.559 | 0.150 | -0.076 | 0.650 | 0.907 | -0.027 | 0.070 | 0.222 | 0.752 | 0.029 |
| Aggression | -0.078 | 0.495 | 0.875 | -0.018 | -0.071 | 0.270 | 0.793 | -0.035 | -0.006 | 0.017 | 0.740 | -0.049 | -0.001 | 0.013 | 0.933 | -0.014 | -0.266 | 0.614 | 0.664 | -0.094 | -0.168 | 0.231 | 0.468 | -0.069 |
| Residual | 1.302 | 3.447 | 0.706 | 0.296 | 0.675 | 1.617 | 0.676 | 0.334 | 0.023 | 0.090 | 0.796 | 0.200 | -0.008 | 0.089 | 0.927 | -0.099 | -0.150 | 3.289 | 0.964 | -0.053 | 1.773 | 1.191 | 0.137 | 0.724 |

[^11]Table J.5: Oaxaca-Blinder outcomes at age 12, Decomposition Part 2 (Males)

| Average number of absences |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Effect | SE | $P$-Val | Fraction |  |
| -1.017 | 0.931 | 0.275 | - |  |
| -0.524 | 0.346 | 0.130 | 0.515 |  |
| -0.493 | 0.918 | 0.591 | 0.485 |  |

$\begin{array}{llll}-0.237 & 0.228 & 0.298 & 0.233\end{array}$

|  |  |
| :---: | :---: |
|  |  |
|  |  |
|  |  |




$\begin{array}{llll}-0.006 & 0.009 & 0.477 & 0.090 \\ -0.004 & 0.008 & 0.588 & 0.062 \\ -0.002 & 0.007 & 0.811 & 0.024 \\ -0.002 & 0.007 & 0.762 & 0.030 \\ -0.016 & 0.015 & 0.267 & 0.239\end{array}$



| Total Diff. in Means |
| ---: |
| Explained |
| Unexplained |




Cognitive
Attention Problems
 Warmth/Empathy

[^12]
## K Summary of Previous Analyses of NFP

In this section, we summarize the findings from previous studies that examine the treatment effects of the NFP by each of the three trials. Tables K.1-K. 8 present the studies for Elmira; Tables K.9-K. 13 for Memphis and Tables K.14-K. 15 for Denver.

## Table K.1: Summary of Olds et al. (1986), Elmira Trial

## A. Paper Title

Improving the Delivery of Prenatal Care and Outcomes of Pregnancy: A Randomized Trial of Nurse Home Visitation
B. Period of Investigation

Time of registration in the program, at the 32nd week of pregnancy and medical records at labor delivery
C. Sample Size

500 women invited, 400 enrolled. Comparison: 165 (group 1 and 2). Treatment: 189 (Group 3 and 4).
From the initial 400 women enrolled, 46 non-white women were removed because of the small sample sizes (when conditioned on other pre-program variables of interest).

## D. Main Goal

Evaluation of the effectiveness of the comprehensive prenatal program as means of improving antepartum social support, health habits and obstetrician health status on on length of gestation and birth weight

## E. Outcomes

Use of services and support systems, obstetrician complication after enrollment, obstetrician conditions and health habits, number of cigarettes, birth weight and length of gestation
F. Methods

Differences in means. OLS for continuous outcomes and logistic linear model for dichotomous outcomes

## G. Main Results

Nurse home visited group improved in the use of community services, informal social support, and health habits. No overall effect on either birth weight or length of gestation. But, positive effects were present for the children of young adolescents $(<17)$ and smokers

# Table K.2: Summary of Olds et al. (1986), Elmira Trial 

## A. Paper Title

Preventing Child Abuse and Neglect: A Randomized Trial of Nurse Home Visitation

## B. Period of Investigation

Time of registration in the program, at $6,12,24$ months of the child's life.
C. Sample Size

Comparison: 165 (group 1 and 2) Treatment: 189 (Group 3 and 4). From the initial 400 women enrolled 46 non-white women were removed in this analysis because of the small number to crossclassify race with other variables important for the statistical analysis. In the 2 years of child's life attrition between $12 \%$ and $21 \%$

## D. Main Goal

Effect of prenatal program on childhood health and developmental problems in the 2 years of child's life, including abuse and neglect

## E. Outcomes

Child abuse and neglect.
Reports of infant temperament, behavior problems, and maternal reaction to behavioral problems.
Restriction and punishment and provision of play material.
Infant mental development (Bayley and Cattell)
Emergency room visits.
F. Methods

Differences in means (Adjusted for baseline characteristics). Simultaneous statistical inference. For continuous outcomes, OLS and for dichotomous outcomes logistic linear model (Assuming a binominal distribution). And low incidence outcomes, in the form of counts (number of emergency room visits) in the log-linear model (assuming a poisson distribution)
G. Main Results

Positive results concentrated among women at greater risk (younger mothers, poor, unmarried) of caregiving dysfunction. These group had fewer records of child abuse and neglect during the first two years of child lives; they punished their children less; they provided with more playing material. Children had less emergency room visits. For the infants of all the nurse-visited women: they visited the emergency room less, they were seen by physicians less frequently for accidents and poisoning in the second years of life

# Table K.3: Summary of Olds et al. (1988), Elmira Trial 

## A. Paper Title

Improving the Life-Course Development of Socially Disadvantaged Mothers: A Randomized Trial of Nurse Home Visitation

## B. Period of Investigation

Time of registration in the program, at $6,10,22,46$ months of children life. SSA records of number of days that women and their children received public assistance from the index child's birth to fourth birthday

## C. Sample Size

Comparison: 165 (Groups 1 and 2).
Treatment: 189 (Group 3 and 4).
During the first 4 years of child's life attrition was between $15 \%$ and $21 \%$

## D. Main Goal

Effect on improving maternal life-course development

## E. Outcomes

Mother's educational achievement (enrollment, graduation, years of schooling)
Employment
Child Care
Public assistance.
Subsequent-pregnancy

## F. Methods

Differences in means (Adjusted for baseline characteristics, classification factors and interaction). For continuous outcomes, OLS and for dichotomous outcomes logistic linear model (Assuming a binominal distribution). And low incidence outcomes, in the form of counts (number of emergency room visits) in the log-linear model (assuming a poisson distribution)

## G. Main Results

Up through the four year old index child's life, the nurse visited women who had not graduated from high school returned to school more rapidly than the comparison group. Treated poor, unmarried women showed an $82 \%$ increase in the number of months employed, had $43 \%$ fewer subsequent pregnancies, and postponed the birth of the second child on average by 12 months. During the first two years after delivery, nurse-visited, poor unmarried older women received $40 \%$ less of public assistance than comparison group

# Table K.4: Summary of Olds et al. (1994), Elmira Trial 

## A. Paper Title

Intellectual Impairment in Children of Women who Smoke Cigarettes During Pregnancy

## B. Period of Investigation

Time of registration in the program, 34th week of gestation, measures at $6,10,22,36,48$ months of the child's life
C. Sample Size

Comparison: 165 (Groups 1 and 2).
Treatment: 189 (Groups 3 and 4).
During the first 4 years of child's life attrition was between $15 \%$ and $21 \%$.
Analysis limited to whites.
From the initial 400 women enrolled 46 non-white women were removed in this analysis because of the small number to cross-classify race with other variables important for the statistical analysis. The estimation of the effect of smoking focused on the comparison sample because the nurse visited group altered the relationship prenatal smoking and Children IQ

## D. Main Goal

Study the effect of maternal cigarette smoking during pregnancy on children's intellectual functioning during the first 4 years of life, adjusting for the primary confounding influences

## E. Outcomes

Intellectual functioning scores: Bayley mental development index (12 months), Cattell (24 months), Stanford-Binet (36 months and 48 months)
F. Methods

General linear model methods, including mixed models to analyse repeated measures with missing data. Newton Raphson and EM algorithms. Adjustment for baseline characteristics, classification factors (marital status, SES), covariates and their interactions. To analyze the effect of smoking, the comparison is made between women in the comparison group who smoke 10 or more cigarettes per day during pregnancy and comparison women who smoke 0
G. Main Results

Children in the comparison group whose mothers smoke 10 or more cigarettes per day during pregnancy had Stanford-Binet scores at 3 and 4 years that were 4.35 points lower (after controlling for several variables) than their counterparts who did not smoke prenatally

## Table K.5: Summary of Olds et al. (1994), Elmira Trial

## A. Paper Title

Does Prenatal and Infancy Nurse Home Visitation Have Enduring Effects on Qualities of Parental Caregiving and Child Health at 25 to 50 Months of Life?

## B. Period of Investigation

Time of registration in the program, and at $34,36,46$, and 48 months of the child's life.

## C. Sample Size

Comparison: 165 (group 1 and 2)
Treatment: 189 Group 3 and 4.
During the first 4 years of child's life attrition was between $15 \%$ and $21 \%$.
Analysis limited to whites.
From the initial 400 women enrolled 46 non-white women were removed in this analysis because of the small number to cross-classify race with other variables important for the statistical analysis

## D. Main Goal

Examine the effect of a randomized trial of a nurse home visitation program on the health, development, rates of child maltreatment, and living conditions of children from 3 to 4 years of age

## E. Outcomes

Cases of abuse and neglect
Intellectual functioning: Stanford-Binet
Home hazards
Health care encounters
Home inventory
F. Methods

Differences in means (Adjusted for baseline characteristics, classification factors, covariates and interactions). For continuous outcomes, OLS and for dichotomous outcomes logistic linear model (Assuming a binominal distribution). And low incidence outcomes, in the form of counts (number of emergency room visits) in the log-linear model (assuming a poisson distribution)

## G. Main Results

No treatment differences in the rates of child abuse and neglect children's intellectual function from 25 to 48 moths of age. However, nurse-visited children lived in homes with fewer hazards, they had $40 \%$ fewer injuries and $45 \%$ fewer behavioral and parental coping problems. They made $35 \%$ fewer visits to the emergency room. treatment mothers were more involved with and punished their children to a greater extend than comparison mothers. The functional meaning of punishments is different between the treatment group and the comparison group

# Table K.6: Summary of Olds et al. (1997), Elmira Trial 

## A. Paper Title

Long-term Effects of Home Visitation on Maternal Life Course and Child Abuse and Neglect: Fifteen-Year Follow-up of a Randomized Trial

## B. Period of Investigation

Time of registration in the program, and at 15 years of the child's life
C. Sample Size

Of the 400 pregnant women who enrolled,
324 participated in the fifteen year follow up.
Comparison group (Groups 1 and 2)
Treatment group (Group 4)
D. Main Goal

Evaluate the long-term effects of the program on women's life course and child abuse and neglect
E. Outcomes
*Rates of subsequent births (self-report) *Use of welfare (AFDC, food stamps, medicaid, self report) *Maternal substance abuse, arrests, convictions, and child abuse and neglect reports from birth up to 15 years of child life (New York State records)

## F. Methods

Intent to treat approach. Differences in means (Adjusted for baseline characteristics, classification factors, covariates and interactions). For continuous outcomes, OLS was used and for low frequency count data (number of reports pf child maltreatment) the log-linear model was used (assuming a poisson distribution). The analysis reported here was not limited to one race

## G. Main Results

Women visited by nurses during pregnancy and infancy were involved in fewer child abuse and neglect episodes than comparison group women. Among unmarried and low SES women at initial of enrollment, treated women had 1.3 vs. 1.6 subsequent births (in contrast to comparison group), 65 vs 37 months between the birth of the first child and the second, 60 vs 90 months receiving AFDC, 0.41 vs 0.73 behavioral impairments due to alcohol and drugs, 0.18 vs 0.58 arrest by self report, and 0.16 vs 0.90 arrest according to the state. All differences are significant at the $95 \%$ confidence level

# Table K.7: Summary of Olds et al. (1998), Elmira Trial 

## A. Paper Title

Long-term Effects of Nurse Home Visitation on Children's Criminal and Antisocial Behavior: 15-Year Follow-up of a Randomized Controlled Trial
B. Period of Investigation

Time of registration in the program, and at 15 years of the child's life
C. Sample Size

400 pregnant women enrolled. A total of 315 adolescent offspring participated in the 15 years follow up study. Comparison group ( 1 and 2) and treatment group (group 3 and 4 separately)
D. Main Goal

Evaluate the long-term effects of the program on children's criminal and antisocial behavior

## E. Outcomes

Children's self-reports of running away, arrests, convictions, initiation of sexual intercourse, number of sex partners, pregnancy, and use of illegal substances.
School records of suspensions.
Teachers'reports of children's disruptive behavior in school.
Parents'reports of the children arrests and behavioral problems

## F. Methods

Intent to treat approach. Differences in means (Adjusted for baseline characteristics, classification factors, covariates and some interactions). For continuous outcomes, OLS was used and for low frequency count data (eg, number of arrests) the log-linear model was used (assuming a poisson distribution). Low incidence count outcomes with values higher than 20 were analyzed in a loglinear model, correcting for over-dispersion. For outcomes reported by more than one respondent (eg, child, teacher), they used repeated measured analysis (adding fixed factors for respondent and random factor for individual). For children self-reports of antisocial and delinquent acts, they used factor analysis and created two factors for multiple hypothesis testing: major delinquency and minor antisocial acts. The analysis reported here was not limited to one race

## G. Main Results

Adolescents born to women who received the program during pregnancy and infancy and who were unmarried and from low SES at registration, in contrast to the comparison group, reported lower incidence of running away ( 0.24 vs 0.60 ), fewer arrests ( 0.20 vs 0.45 ), fewer convictions ( 0.09 vs 0.47 ), fewer lifetime sexual partners ( 0.92 vs 2.48 ), fewer cigarettes per day ( 1.50 vs 2.50 ), and fewer days of alcohol consumption (1.09 vs 2.49). Parents in the treatment group (4) reported that their children had fewer problems related to alcohol and drugs use ( 0.15 vs $0.35)$. Differences statistically significant. No effect on teachers' reports, short-term or long term suspensions, adolescent initiation of sexual life, and the two factors: major delinquency and minor antisocial acts

Table K.8: Summary of Eckenrode et al. (2010), Elmira Trial

## A. Paper Title

Long-term Effects of Prenatal and Infancy Nurse Home Visitation on the Life Course of Youths: 19 year follow up

## B. Period of Investigation

Time of registration in the program, and at 19 years of the child's life
C. Sample Size

400 pregnant women enrolled. A total of 310 adolescent offspring participated in the 19 years follow up study. Comparison group (1 and 2) and treatment group (group 3 and 4 separately)

## D. Main Goal

Evaluate the impact of the prenatal and infancy nurse visits on youths' life course development

## E. Outcomes

Youth self reports of educational achievement, reproductive behaviors, welfare use, criminal involvement, and drug use
F. Methods

Differences in means (Adjusted for baseline characteristics, classification factors, covariates). For continuous outcomes, OLS was used. For dichotomous outcomes, generalized linear model with log link and binomial error distributions was used. For count data, log link and negative binomial was assumed. To estimate the hazard of first arrest, the Cox proportional hazards method was used. Growth curves for arrest episodes over time were estimated in a generalized mixed model with cubic age, with log link and negative binomial error. The analysis reported here was not limited to one race

## G. Main Results

In contrast to the comparison group, girls born to women in the pregnancy and infancy nursevisited group were less likely to be arrested ( $10 \%$ vs $30 \%$ ), and convicted ( $4 \%$ vs $20 \%$ ) and had fewer lifetime arrests ( 0.10 vs 0.54 ) and convictions ( 0.04 vs 0.37 ).
Nurse-visited girls born to unmarried and low SES mothers had fewer children and were less likely to use Medicaid use than their comparison group counterparts

# Table K.9: Summary of Kitzman et al. (1997), Memphis Trial 

## A. Paper Title

Effect of Prenatal and Infancy Home Visitation by Nurses on Pregnancy Outcomes, Childhood Injuries, and Repeated Childbearing: A Randomized Controlled Trial

## B. Period of Investigation

Time of registration in the program, at 28th and 36th week of pregnancy,and at $6,12,24$ months of the child's life. Medical and social service records were abstracted

## C. Sample Size

1290 women invited, 1139 enrolled. Comparison group 1 (166), Comparison group 2 (515), treatment group 3 (230), treatment group 4 (228)

## D. Main Goal

To examine the impact of pregnancy and infancy home visits by nurses on pregnancy-induced hypertension, pre-term delivery, and low birth weight; on children's injuries, immunizations, mental development, and behavioral problems; and on maternal life course

## E. Outcomes

Medical records: Pregnancy-induced hypertension (PIH), preterm delivery, low birth weight, children's injuries, ingestions and immunizations.
Mothers' reports of children's behavioral problems;
Children mental development (Bayley scales and Achenbach Child Behavior Checklist)
Mothers'reports of subsequent pregnancy, educational achievement, and labor force participation
Use of welfare: AFDC, from state records

## F. Methods

Differences in means (Adjusted for baseline characteristics, classification factors, covariates and interactions). For continuous outcomes, OLS and for dichotomous outcomes (eg, PIH) logistic linear model (Assuming a binominal distribution).
And low incidence outcomes, in the form of counts (number of health care encounters) in the log-linear model (assuming a Poisson distribution). Pregnancy models contrast comparison group (group 1 and 2) vs treatment group 3, and vs treatment group 4.
Postnatal models contrast comparison group (1 and 2) with treatment group 4 (the one that received both prenatal and infancy nurse visits)

## G. Main Results

Women visited by nurses during pregnancy had had less PIH ( $13 \%$ vs $20 \%$ ) compared to the comparison group. During the first two years after delivery, women in the treatment group had fewer health care encounters for children in which injuries were detected ( 0.43 vs 0.55 ), fewer days of children's hospitalization ( 0.03 vs 0.16 ), and fewer second pregnancies ( $36 \%$ vs $47 \%$ ).
No program effects on pre-term delivery, or low birth weight; children's immunization rates, mental development or behavioral problems; or mothers' education and employment

# Table K.10: Summary of Kitzman et al. (2000), Memphis Trial 

## A. Paper Title

Enduring Effects of Nurse Home Visitation on Maternal Life Course: A 3-Year Follow-up of a Randomized Trial

## B. Period of Investigation

Data from assessments at time of registration in the program and 54th months of the child's life

## C. Sample Size

Of those cases randomized with no fetal or child death, follow up interviews were completed on $91 \%$ of the cases ( 443 in comparison group 2 and 203 in treatment group 4)

## D. Main Goal

Effectiveness of the NFP prenatal and infancy home visitation program on the maternal life course 3 years after the program ended

## E. Outcomes

Mothers: Rate of subsequent pregnancy, mean of interval between first and second child, educational achievement, number of months in the labor force, and number of months enrolled in AFDC, food stamps (FS), Medicaid, WIC.
Administrative data from the Tennessee Dept. of Social Service were obtained for AFDC and Food Stamp.

## F. Methods

Intent to treat approach. Differences in means (classification factors, covariates and interactions). For continuous outcomes, OLS was used, for dichotomous outcomes (eg, cohabitation) the logistic linear model (assuming binominal distribution) and for low frequency count data (subsequent pregnancies) the log-linear model was used (assuming a poisson distribution) Models focused on contrasting comparison group 2 with treatment group 4

## G. Main Results

Contrasted with women in the control group, women who received the NFP treatment had fewer pregnancies ( 1.15 vs 1.34 ), fewer closely spaced subsequent pregnancies ( 0.22 vs 0.32 ), longer intervals between the birth of the first and second child ( 30.25 vs 26.60 ), and fewer months using AFDC ( 32.55 vs 36.29 ) and FS ( 41.57 vs 45.04 ). Compared with the effect of the program while the visits were being conducted, the effect after it ended was essentially equal for AFDC, greater for FS, and greater for rates of closely spaced pregnancies.

# Table K.11: Summary of Olds et al. (2004), Memphis Trial 

## A. Paper Title

Effects of Nurse Home-Visiting on Maternal Life Course and Child Development: Age 6 Follow-Up Results of a Randomized Trial

## B. Period of Investigation

Data from assessments at time of registration in the program and 6 years of the child's life

## C. Sample Size

Of those cases randomized with no fetal or child death, 6 yr follow up interviews were completed on $91 \%$ of the mothers ( 444 in comparison group 2 and 197 in treatment group 4) and $88 \%$ of the children ( 425 in comparison group 2 and 190 in treatment group 4)
D. Main Goal

Effectiveness of the NFP on mothers' fertility and economic self-sufficiency and the academic and behavioral adjustment of their children as they finish kindergarden near their sixth birthday

## E. Outcomes

Mother: Number and timing of subsequent pregnancies; months of employment; use of welfare; food stamps; Medicaid; rates of marriage, cohabitation, and duration of relationships; Child Educational Achievement; Behavioral problems resulting from illegal substances;
Children's behavioral problems (Achenbach Child Behavior Check list), responses to story stems, Intellectual functioning (Kaufman Assessment Battery and Peabody Picture Vocabulary),
Receptive language and academic achievement; Teachers completed the High-tower Teacher-Child Rating Scales

## F. Methods

Differences in means (Adjusted for baseline characteristics, classification factors, covariates and some interactions). For continuous outcomes that didn't violate the normality assumption, OLS was used and for dichotomous correlated outcomes they used generalized estimating equations with logit link function. The timing of the first subsequent birth was explored using proportionalhazards analysis. For teaches' reports of children's classroom behavior, children's representation of aggressive behavior and parental warmth/empathy factors were obtained using principal components analysis

## G. Main Results

Women visited by nurses had fewer subsequent pregnancies (1.16 vs 1.38 ) and births (1.08 vs 1.28 ), longer intervals between births ( 34.28 vs 30.23 ), longer relationships with current partners ( 54.36 vs 45 months), and since the previous follow up, fewer months of using AFDC ( 7.21 vs 8.96 ) and FS ( 9.67 vs 11.50 ) than control group mothers. Nurse visited children were more likely to have been enrolled in formal out of home care between 2 and 4.5 years ( $82 \%$ vs $74.9 \%$ ). Children visited by nurses demonstrated higher intellectual functioning (scores 92.34 vs 90.24 ) and receptive vocabulary scores ( 84.32 vs 82.13 ) and fewer behavioral problems in the borderline or clinical range ( $1.8 \%$ vs $5.4 \%$ ). For the cases of mother with low levels of psychological resources, children had higher achievement test scores, and expressed lefs aggression and incoherence in response to story stems. No statistically significant effect on women's education, duration of employment, rates of marriage, being in a partnered relationship, behavioral problems related to alcohol or drug abuse

# Table K.12: Summary of Olds et al. (2007), Memphis Trial 

## A. Paper Title

Effects of Nurse Home-Visiting on Maternal and child functioning: Age-9 Follow-Up of a Randomized Trial

## B. Period of Investigation

Data from assessments at time of registration in the program and 9 years of the child's life. However, whenever possible they use data from earlier phases

## C. Sample Size

From the initially 743 primary black women randomize to comparison group 2 and treatment group 4 (Core of posnatal evaluations), follow up assessments at child age 9 were completed by $91 \%$ of the mothers, school records were obtained for $88 \%$ of the children and achievement test scores for $83 \%$ of the children.

## D. Main Goal

To examine the impact of pregnancy and infancy home visits by nurses mothers' fertility and children development

## E. Outcomes

Mothers: interval between births, number of children born per year, mothers' stability of relationships, use of welfare, FS and Medicaid, mother's use of substances, mothers' arrest and incarcerations.
Child: academic achievement (GPAs, Tennessee Comprehensive Assessment Test), school conduct, and mental disorders.
Secondary outcomes: women's employment, experience of domestic violence and children's mortality

## F. Methods

Differences in means (Adjusted for baseline characteristics, classification factors, covariates and some interactions). For outcomes on which there are multiple assessments for each mother or child, mixed models were used. This is the first time in this trial where they examined the full longitudinal effects of some of the maternal outcomes. Quantitative dependent variables were analyzed using OLS; and, dichotomous outcomes using logit model. Low-frequency outcomes were analyzed in generalized linear models with negative binominal error and log link assumptions. Factor analysis was used to summarize the information from children Social Competence Scale, the Social Health Profile and the Teachers Observation of Child Adjustment revisited. 3 indices were produced: antisocial behavior, academically focused behavior, and peer affiliation.

## G. Main Results

Nurse-visited women had longer intervals between births, few cumulative subsequent births per year, and longer relationships with current partners. From birth through child age 9, treated mothers used AFDC and FS for fewer months. Nurse-visited children whose mothers have low psychological resources, had better GPA and achievement test scores in math and reading in grades 1 through 3.

## Table K.13: Summary of Kitzman et al. (2010), Memphis Trial

## A. Paper Title

Enduring Effects of Prenatal and Infancy Home Visiting by Nurses on Children: Follow-up of a Randomized Trial Among children at Age 12

## B. Period of Investigation

Data from assessments at time of registration in the program and 12 years of the child's life

## C. Sample Size

From the initially 743 primary black women randomized to comparison group 2 and treatment group 4 (Core of posnatal evaluations), follow up assessments at child age 12 were abstracted for 613 children
D. Main Goal

To evaluate the impact of a nurse visiting program on 12-year-old first born children's use of substances, behavioral adjustment, and academic achievement

## E. Outcomes

Use of cigarettes, alcohol, and marijuana;
Internalizing, externalizing and total behavior problems from parents', teachers' and children's reports.
Academic achievement.
Reading and math achievement using the Peabody Individual Achievement Tests (PIATS).
Reading and Math GPA from grade 1 to 6. Reading and Math from the Tennessee Comprehensive Assessment Program (grade 1 to 6 ).
Arrests reported by mother and child

## F. Methods

Differences in means (Adjusted for baseline characteristics, classification factors, covariates and some interactions). For outcomes on which there are multiple assessments for each child (eg, GPAs), mixed models were used. Quantitative dependent variables were analyzed using OLS; and, dichotomous outcomes using logit model. Low-frequency outcomes were analyzed in generalized linear models with negative binominal error

## G. Main Results

Nurse-visited children reported fewer days of using cigarettes, alcohol, and marijuana and were less likely to report the presence of internalizing disorders that met the borderline or clinical threshold compared to the control group kids. Treated children born to mothers with low psychological resources compared to the control group, had higher scores on PIATs, and on group-administered standardized tests of math and reading. No statistically significant program effects were found on children's externalizing or total behavior problems

# Table K.14: Summary of Olds et al. (2002), Denver Trial 

## A. Paper Title

Home Visiting by Paraprofessionals and by Nurses: a randomized controlled trial

## B. Period of Investigation

Time of registration in the program, and 36 th week of pregnancy, and at $6,12,15,21,24$ months of the child's life
C. Sample Size

1178 women invited, 735 enrolled and randomized. Control group (255), Paraprofessional group (245), and Nurse group (235)

## D. Main Goal

To evaluate the effectiveness of home visiting by paraprofessionals and by nurses as separate means of improving maternal and child health when both types of visitors are trained with the same program model

## E. Outcomes

Mothers: Urine cotinine over the course of pregnancy; women's use of auxiliary services during pregnancy, subsequent pregnancies and births, educational achievement, labor market participation, and use of welfare.
Mother-infant reponsive interactions; family home environments.
Infants emotional vulnerability in response to a fear stimuli and low emotional vitality in response to joy and anger stimuli; children's language and mental development index (MDI), temperament and behavioral problems

## F. Methods

Differences in means (Adjusted for baseline characteristics, classification factors, covariates and some interactions). Comparisons between nurse vs control and paraprofessional vs control. Quantitative dependent variables were analyzed using OLS; and, dichotomous outcomes using logit model. For outcomes on which there are more than one observation in time (eg, maternal-child interaction and home environment), repeated measures were used, adding a fixed factor for time and random factor for individuals. The timing of subsequent pregnancy was analyzed with proportional hazards analysis. Factor analysis of measures for maternal and infancy interaction identified a single internally consisted factor: responsive interaction

## G. Main Results

Mother-child pairs in the paraprofessional group in which the mother had low psychological resources interacted with one another more responsively than the control group. There are no statistically significant paraprofessional effects. Nurse-visited smokers, compared with the control group women, had larger reductions in nicotine; by the index child second birthday, women visited by nurses had fewer subsequent pregnancies; they delayed subsequent pregnancies for longer time, and they worked more during the second year of index child life. Mother-child pairs in the nurse group interacted with one another more responsively than the control group pairs. Nurse-visited children exhibit less emotional vulnerability. Nopre-visited children born to women with low psychological resources were less likely to exhibit low emotional vitality and language delays, and had higher MDI scores. No statistically significant effects on mothers' use of prenatal services, educational achievement, use of welfare or their children behavior problems

## Table K.15: Summary of Olds et al. (2004), Denver Trial

## A. Paper Title

Effects of Home Visits by Paraprofessionals and by Nurses: Age 4 Follow-Up Results of a Randomized Trial

## B. Period of Investigation

Data from assessments at time of registration in the program and 48th months of the child's life

## C. Sample Size

From the initial 735 mothers randomized, 635 completed 4 -y interviews, and 605 completed $4-\mathrm{y}$ child assessments

## D. Main Goal

To evaluate the effects of prenatal and infancy home visiting by paraprofessionals and nurses from child age 2 through age 4

## E. Outcomes

Mothers: Subsequent pregnancies, participation in education and work, use of welfare, marriage, cohabitation, domestic violence, mental health, substance abuse, and sense of mastery.
Mother-child interaction and home environment.
Children: tests of language and executive functioning, mother's report of child externalizing behavior

## F. Methods

Differences in means (Adjusted for baseline characteristics, classification factors, covariates and some interactions). Comparisons between nurse vs control and paraprofessional vs control. Quantitative dependent variables were analyzed using OLS; and, dichotomous outcomes using logit model. The timing of subsequent pregnancy was analyzed with proportional hazards analysis. Principal component analysis was used to create factors from mothers and children externalizing behavior reports (one factor), cognitive tasks on children's ability to sustain attention and inhibitory control (single factor); examination of children's ability to regulate their behavior and emotion (two factors).

## G. Main Results

In general, there are greater effects for paraprofessional-visited mothers than nurse-visited mothers, while greater effects for children in nurse-visited families than in paraprofessional ones.
Paraprofessional: Women were less likely to be married (compared to control), work more and reported better mental health and mastery, had fewer subsequent miscarriages and low birth weight babies. Mother and Children in this group showed greater responsiveness and sensitivity; and in cases of low levels of psychological resources they had home environments more supportive of children learning.
Nurses: Women reported greater intervals between births, less domestic violence, and enrolled the children less frequently in preschool. Children in this group and whose mothers had low levels of psychological resources had home environments that were supportive for learning, more advance language, superior executive functioning, and bgtter behavioral adaptation.

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[^0]:    ${ }^{1}$ The regions are: Inner City, Bisson, Cawthon and Hollywood.

[^1]:    ${ }^{2}$ Nurses completed on average 7 visits during pregnancy (range: $0-18$ ) and 26 visits during the first two years of the child's life (range: 0-71) (Olds and Korfmacher, 1998)

[^2]:     the unconditional mean for the control group. When factor scores were computed, we set the mean in the control group to zero. The second column (Difference in Means)
     experiment and the treatment group stands for the original group 4. The third column (Effect Size) presents the unconditional effect size for the respective group. The four column (Asymp. $p$-value) provides the asymptotic $p$-value for the one-sided single hypothesis test associated with the $t$-statistic for the unconditional difference in means lathe based the $t$-statist (2005). Blocks of outcomes that are tested jointly are separated by lines. The selection of blocks of outcomes is done on the basis of their meaning. Outcomes that share similar meaning are grouped together. Female maternal outcomes allude to mothers whose first child is a girl. Likewise, male maternal outcomes alludes to mothers whose first child is a boy.

[^3]:     the unconditional mean for the control group. When factor scores were computed, we set the mean in the control group to zero. The second column (Difference in Means) gives the unconditional difference in means between the treatment group and the control group. As mentioned in Section 2 , the control group stands for the 2 of the NFP . (Asympor $p$-value) poriated with the $t$-statistic for the unconditional difference in means otwnn (Asymp. $p$-value) provides the asymptotic $p$-value (Unrestricted Permutation - Single $p$-value) presents the one-sided unrestricted permutation $p$-values for the single-hypothesis testing based on the $t$-statistic associated with the treatment indicator. Finally, the last column (Unrestricted Permutation - Stepdown) provides $p$-values that account for multiple-hypothesis testing based on the Stepdown procedure of Romano and Wolf (2005). Blocks of outcomes that are tested jointly are separated by lines. The selection of blocks of outcomes is done on the basis of their meaning. Outcomes that share similar meaning are grouped together. Female maternal outcomes allude to mothers whose first child is a girl. Likewise, male maternal outcomes alludes to mothers whose first child is a boy.

[^4]:    ${ }^{3}$ This follows the same notation as Heckman et al. (2013)
    ${ }^{4}$ We control for pre-program variables $X$ but we keep it implicit to shorten notation.

[^5]:    ${ }^{5}$ Given that the first measure sets the scale, we choose it to be the most correlated with the skill. The results are robust to alterations of this.
    ${ }^{6}$ Carneiro et al. (2003) and Cunha et al. (2010) also discuss identification of factor models.

[^6]:    ${ }^{7}$ This is not strictly required but simplifies the notation.
    ${ }^{8}$ We impute factor scores for individuals that have at least two other factor scores.

[^7]:    ${ }^{9}$ We implicitly control for pre-program variables.

[^8]:    

[^9]:    Children's Skills Age 6 | Age 12 outcomes | Cognition | $\begin{array}{c}\text { Attention } \\ \text { probs }\end{array}$ | $\begin{array}{c}\text { conduct } \\ \text { probs }\end{array}$ | Empathy | Agression |
    | :--- | :---: | :---: | :---: | :---: | :---: |
    | $\begin{array}{l}\text { Average TCAP percetile. Years 1- } \\ \text { 5 after KG: Language }\end{array}$ | 0.529 | 0.975 | 0.993 | 0.636 | 0.882 |
    | $\begin{array}{l}\text { PIAT reading comprehension } \\ \text { derived score }\end{array}$ | 0.420 | 0.794 | 0.941 | 0.867 | 0.812 |
    | $\begin{array}{l}\text { Average math grades. Years 1-5 } \\ \text { after KG } \\ \text { Average TCAP percetile. Years 1- }\end{array}$ | 0.425 | 0.953 | 0.830 | 0.871 | 0.792 |
    | $\begin{array}{l}\text { 5 after KG: Math }\end{array}$ | 0.940 | 0.951 | 0.940 | 0.875 |  |
    | $\begin{array}{l}\text { PIAT mathematics derived score } \\ \text { SC use of alc, mar, tob. Lat 30 }\end{array}$ | 0.433 | 0.503 | 0.817 | 0.976 | 0.652 |
    | $\begin{array}{l}\text { days }\end{array}$ | 0.845 | 0.970 | 0.751 | 0.791 | 0.538 |
    | $\begin{array}{l}\text { Internalizing disorders - youth } \\ \text { report } \\ \text { Clinical or borderline } \\ \text { anxious/depressed disorder } \\ \text { Average number of absences, }\end{array}$ | 0.582 | 0.934 | 0.911 | 0.905 | 0.509 |
    | $\begin{array}{l}\text { school years 1-5 after KG }\end{array}$ | 0.537 | 0.936 | 0.771 | 0.916 | 0.687 |

[^10]:    Notes: The table shows p-values for the Wald test: $z=\frac{\hat{\boldsymbol{\alpha}}^{0}\left(\overline{\hat{\boldsymbol{\theta}}}_{1}^{0}-\overline{\hat{\boldsymbol{\theta}}}_{0}^{0}\right)-\hat{\boldsymbol{\alpha}}^{1}\left(\overline{\hat{\boldsymbol{\theta}}}_{1}^{1}-\overline{\hat{\boldsymbol{\theta}}}_{0}^{1}\right)}{\sqrt{\operatorname{Var}\left(\hat{\boldsymbol{\alpha}}^{0}\left(\overline{\hat{\boldsymbol{\theta}}}_{1}^{0}-\overline{\boldsymbol{\theta}}_{0}^{0}\right)-\hat{\boldsymbol{\alpha}}^{1}\left(\overline{\boldsymbol{\theta}}_{1}^{1}-\overline{\hat{\boldsymbol{\theta}}}_{0}^{1}\right)\right)}}$

[^11]:    Notes: The indices are means of the non-missing items. The fractions are proportions of the total conditional difference in means.

[^12]:    Notes: The indices are means of the non-missing items. The fractions are proportions of the total conditional difference in means.

